

# TRUST, TRUTH, STATUS AND IDENTITY

an experimental inquiry \*

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**Abstract:** In an experiment involving a standard trust game and a costless signalling game, it is demonstrated that economically relevant norm-based behaviors (trust, reciprocity and truth-telling) vary with *social* identity.

The experimental procedure induced two trivial social identities. In one version, a status difference was induced. The results permitted a succinct description of identity effects: subjects held own-group members to a higher standard; and high status subjects held everyone, including themselves, to a higher standard. To illustrate the “high status/high standards” phenomenon, subjects’ “standards” were estimated from a simple identity model for a subset of the data.

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# 1 Introduction

An ever-growing body of experimental research in economics demonstrates undeniable and systematic deviations from pure self-interest. Many of the situations in which these deviations have been documented share a particularly simple form: what a purely self-interested person would do is diametrically opposed to what a “decent” person *should* do. For lack of a better term, call what one *should* do a norm. And call the aforementioned class of results norm-based deviations from pure self-interest.

The simplest example of this type of behavior involves a sharing norm: we’ve all been taught since grade school that we *should* share. Consequently, there is considerable experimental evidence that subjects share money earnings in the lab even when there is no plausible purely self-interested reason for doing so. Beyond simply sharing, however, the class of norm-based deviations from pure self-interest encompasses phenomena such as trust, reciprocity and altruism which are thought to be central to everything from the functioning of labor markets (Bewley 1999; Akerlof 1982) to aggregate economic growth (Knack and Keefer 1997). For an overview of many economics experiments dealing with these phenomena, see Camerer (2003).

In terms of economic theory, there are two main approaches to explaining norm-based deviations from pure self-interest. In the “social identity” approach—represented in the economics literature by Akerlof and Kranton (2000, 2005)—the building blocks are categorizations. Individuals place themselves and others into social categories. Each category is a social identity. Norms in this view—how we *should* and *should not* behave—are tied inextricably to social identities. To complete the setup, individuals have preferences over their own and others’ norm-compliance.

In another, more established, approach, norms are modeled as stable, *individual* traits. In these models—generally referred to as “social preferences” models—norm-concern can be boiled down to individual-specific parameters, with norm-based variation in behavior explained by individual heterogeneity in these parameters, sometimes complemented by

informational phenomena such as signaling (Benabou and Tirole 2004; Levine 1998; and Charness and Rabin 2002). Some of the most widely-cited social preferences models make the further assumption that norms are the same for everyone—i.e., assuming that one norm is *the* norm (see, e.g., Fehr and Schmidt 1999; Bolton and Ockenfels 2000). Behavioral heterogeneity in these models stems from differences in how much individuals care about *the* norm vis-a-vis standard economic incentives. Again, this tradeoff is modeled as a stable, individual-specific trait: some people are purely self-interested, some people care only about, e.g., “fairness,” and most people lie somewhere in between.

The key difference between the social identity and social preferences approaches is where norms reside, and this difference leads to a testable hypothesis. Specifically, since the social identity approach asserts that norms are tied to social categories, manipulating categories will change observed behavior in situations where norms are important. On the other hand, if we are careful to not change anything systematically about the *individuals* involved, the social *preferences* approach predicts that merely changing categorizations will not produce significant changes in behavior.<sup>2</sup>

Herein I report the results of an experiment in which I tested this hypothesis by varying

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<sup>2</sup>A handful of recent papers in experimental economics—of which I was unaware when designing and implementing the current experiment—also investigate identity in the lab. In contrast to the current experiment, these other papers treat identity as more about small-group/team interaction than pure categorization and thus induce identity in ways that substantively change subjects’ strategic situations—making the attribution of observed patterns in behavior to identity less straightforward.

Specifically, in Chen and Li (2007), McLeish and Oxoby (2007) and Eckel and Grossman (2005) the authors used intra-group communication about, and intra-group cooperation on, shared group projects to induce identities—thus making identity literally about small-group interaction. Such pre-game communication and cooperation can change the set of equilibria for many games, and possibly confounds identity with trust-building/familiarity effects and/or repeated game effects. In contrast to these experiments, the current experiment did not allow pre-game cooperation or communication.

In Charness, Rigotti and Rustichini (2007), because individual earnings were made to depend on own-group earnings, the mere expectation that *others* will exhibit in-group bias may be sufficient to induce actual in-group bias for some forms of social *preferences* (e.g. Charness and Rabin 2002) in the particular normal-form games examined. Heuristically, if in-group members are expected to be more likely to cooperate, then defecting in a Prisoner’s Dilemma destroys more expected surplus when one’s partner is an in-group member. Additionally, the introduction of an audience of own-group or other-group members may implicate preferences for conformity (e.g., Asch 1956) since the composition of the audience determined which player’s preferences were aligned with the majority of those present. In fact, such numerical majority/minority effects have a long history in the social identity literature beginning with Sachdev and Bourhis (1984). By way of contrast, the current experiment did not incorporate a group component to individuals’ earnings from game-play in order to keep incentives as transparent as possible; and majority/minority effects were ameliorated by ensuring that both social identities were of equal size and that each social identity had the same number of subjects present in each session.

subjects' social identities without changing anything typically considered decision-relevant about the individuals involved: subjects were publicly randomly divided into two groups; subjects' group affiliations served as trivial social identities. Subjects then played standard experimental games: the Trust Game (Berg, Dickhaut and McCabe 1995) and a two-player costless signalling game I call the Truth Game, where the central strategic consideration was whether to lie for potential monetary gain.<sup>3</sup>

The results support the social identity approach, in that even these simple, information-less, categorizations significantly affected subjects' behavior. In particular, the strategic situations considered were rich enough to go beyond mere "in-group bias"—treating members of your own group more favorably—and reveal a more nuanced phenomenon: subjects held members of their own social identity to a higher standard. Furthermore, when a status difference between groups was introduced, members of the higher status group held *everyone*, including themselves, to a higher standard.<sup>4</sup>

Before describing the results more precisely, it will prove helpful to give an example of what I mean by "social identity." The concept of social identity is vividly illustrated by a classic experiment: Robbers Cave (Sherif, *et al.* 1954.) In the Robbers Cave experiment, two groups of otherwise-similar schoolboys were sequestered in separate camps at Robbers Cave State Park in Oklahoma. During this initial week-long phase, each group engaged in cohesion-building activities: running, hiking, swimming—standard summer camp fare. The groups became cohesive enough to spontaneously name themselves: one group called itself the Rattlers, while the other group deemed itself the Eagles.

In the second phase, the two groups engaged in competitive intergroup activities—baseball tournaments, tug-of-war and the like. To put it dryly, in the second phase each group demonstrated out-group aversion. In fact, the groups' aversion to each other was sufficiently strong to induce minor acts of arson and multiple attempted assaults.<sup>5</sup>

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<sup>3</sup>The game was similar to the cheap talk games considered in Sanchez-Pages and Vorsatz (2007) and Gneezy (2005).

<sup>4</sup>There is existing experimental and empirical evidence to support this "high status/high standards" phenomenon (see the related literature section, below). However, to my knowledge, the individual pieces of evidence have not previously been connected, and the phenomenon has not previously been articulated.

<sup>5</sup>For example, one attempted assault stemmed from the Rattlers' desire to settle the score for a midnight

In a social identity framework, we can think of the experiment as creating two social identities—Eagles and Rattlers—and revealing the existence of a particular (social) identity-based norm—out-group aversion, more commonly referred to as in-group bias. While the strength of the manipulation raises many concerns about whether this was a purely identity-based phenomenon, most of the obvious alternative explanations have been ruled out by four decades of experimental work in social psychology. Starting with Tajfel, *et al* (1971), and using simple disinterested money-splitting decisions, in-group bias has been shown to be a significant phenomenon even when the social identities are the result of simply randomizing subjects into two different groups.<sup>6</sup>

Given this background, the current experiment proceeded in two phases: an identity-inducement phase and a game-playing phase. Following the social psychology literature, the identity-inducement phase created two social identities by randomly dividing subjects into two groups. In one version of the experiment, the two groups were of equal status. In a second version of the experiment—following, in the economics literature, Eckel, Ball, Grossman and Zame (2001)—I randomly chose one of the groups to have “high status,” and reinforced subjects’ sense of status by varying tasks and comfort levels.<sup>7</sup> I provide evidence later in the paper suggesting that the status manipulation did not simply induce a mood effect or a wealth effect.

The game-playing phase of the experiment was standard experimental economics fare. Subjects played two widely-researched economic games. In the Trust Game (Berg, Dickhaut and McCabe 1995), one subject—the sender—is given a fixed sum of money, of which he can send some, all or none of to his anonymous co-player—the receiver. The amount sent is tripled, at which point the receiver can return any of this tripled amount to the sender. The sender’s action involves trust, as his co-player has the option of returning nothing; and

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raid on their cabin by the Eagles, and featured a mess hall raid where Rattlers armed themselves with sticks and bats. A later attempt involved the Rattlers lying in wait for the Eagles with rock-filled socks as weapons.

<sup>6</sup>Most importantly, these experiments have demonstrated that the effects of social identity are *not* about familiarity created through in-group interaction, nor about any true similarity among group members.

<sup>7</sup>High status subjects were seated three per row while low status subjects sat in rows of five; and high status subjects enjoyed refreshments while low status subjects performed a boring, tedious task: re-alphabetizing a list of names by hand

the receiver’s action involves reciprocity, as pure self-interest dictates keeping the entire amount sent, while rewarding “nice” actions requires returning a non-zero amount.

The second game was a two-player asymmetric information game I call the Truth Game. The Truth Game can be thought of as modeling the interaction between a used car salesman and a potential customer. The salesman has private information about the quality of the car, while the buyer must decide to either buy the car or walk away. The salesman, of course, can claim the car is reliable, but the buyer has no way to verify the salesman’s claims. The gist of the strategic situation is that salesmen can lie to buyers for potential monetary gain.

The relevance of the Trust Game and the Truth Game to the current inquiry is that in each game normative behavior is both relatively straightforward (trust and honesty, respectively) and at odds with purely self-interested behavior. Furthermore, the games’ strategic situations are relatively simple, yielding clear equilibrium predictions when agents are purely self-interested. Further still, notwithstanding their simplicity these two games represent situations that many economists consider important.<sup>8</sup> Finally, since trust is possible in both games,<sup>9</sup> but is more clearly the normative behavior in the Trust Game, comparing behavior in these two games will provide evidence about the precise effects of social identity: i.e., whether identity affects norms and norm-compliance, or simply makes subjects more generous, trusting, etc., irrespective of the normative environment.<sup>10</sup>

There were two main results in the data. Firstly, I found that equal-status identities induced an effect that is, in retrospect, obvious, but has not previously been demonstrated: subjects held members of their own group to a higher standard. Secondly, introducing a status difference extended this own-group norm concern into a form of paternalism: high status subjects held everyone—including themselves—to higher standards. Specifically, high

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<sup>8</sup>The standard quote concerning the importance of trust, and hence the Trust Game, is due to Arrow (1972): “Virtually every economic transaction has within it an element of trust . . .” The strategic situation in the Truth Game is clearly at the heart of many situations involving asymmetric information: contracting with an expert, voting based on campaign promises, etc.

<sup>9</sup>Trust in the Truth Game would be credulity—simply believing sellers’ claims of car quality.

<sup>10</sup>Since sellers always prefer buyers to buy, one can think of increased generosity in the Truth Game as an increased propensity to buy irrespective of sellers’ messages.

status subjects were both more trusting in the Trust Game and more honest in the Truth Game; and in the Trust Game, where there is an opportunity to punish “bad” behavior, high status subjects both punished lack of trust more severely, and rewarded high levels of trust more generously, than their low status counterparts.<sup>11</sup> Across-version comparisons of behavior in the Truth Game revealed a potentially economically relevant consequence of this high status/high standards phenomenon: contrary to all purely self-interested equilibria of the Truth Game, in the unequal-status version of the experiment sellers’ messages actually benefitted buyers: sellers apparently anticipated the increased honesty from high status sellers and used this inference to their benefit. In the equal-status version of the experiment, buyers did no better than if they had ignored sellers’ messages altogether.

Finally, in the Trust Game, across-version comparisons of reciprocal behavior yielded insights into why status affects behavior: the data suggest that high status emboldened subjects to impose their values on everyone in precisely the same manner that subjects were emboldened to impose their values on their “own” people *sans* status differences. This effect, too, can be summarized succinctly: “give someone an office and they become officious.”<sup>12</sup>

To investigate the high status/high standards phenomena further, following Akerlof and Kranton (2005), I constructed a simple social identity model of subjects’ preferences having the following form, where  $a_k$  denotes player  $k$ ’s action:

$$U_j = u(a_j, a_i) - \alpha_j(a_j - a_{c_j, c_i}^{Ideal})^2 \quad (1)$$

In this two-player social identity model, player  $j$  derives utility from two sources: the purely self-interested preferences of classical economics— $u(a_j, a_i)$ —and from living up to her *ideals*— $a_{c_j, c_i}^{Ideal}$ . An ideal is the specific action prescribed by relevant social norms, and

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<sup>11</sup>The terms “punish” and “reward” are used here in the sense of leaving senders worse off or better off than if they had not trusted their co-player at all—i.e., whether the ratio of money returned to money sent is less than or greater than one. Since each dollar sent is tripled, feasible return ratios range from zero to three irrespective of the amount sent.

<sup>12</sup>The source of this expression is my father-in-law, who gleaned this phenomenon from decades of counseling individuals and dealing with large organizations as the senior pastor of several large churches throughout the U.S.

depends on the social identities of both players— $c_j$  and  $c_i$  in Equation 1.<sup>13</sup> To make the distinction clear, the relevant *norm* in the Trust Game is trust; while one possible *ideal* is “send exactly seven dollars to members of your own social identity, otherwise send nothing.” Finally, player  $j$  is willing to trade economic utility against “identity utility” at a rate measured by  $\alpha_j$ .

Given this model, I used a semi-parametric estimation technique—Censored Least Absolute Deviations (Powell 1984)—to estimate subjects’ ideals explicitly. The estimates suggest that high status subjects punished norm-deviance, ideally, more than twice as severely as low status subjects; on the other hand, when co-players fully complied with norms, high status subjects were more generous than low status subjects.

The remainder of this paper is organized as follows. First, I provide a brief sketch of related literature. Next, the experimental design is presented in detail. After this, results are presented for each game, separately. Next, I construct a simple identity model and use receivers’ actions in the Trust Game to estimate receivers’ Ideals. Subsequently, I provide some evidence that the results cannot be plausibly explained by either mood effects or wealth effects. In the concluding section, I summarize the results and suggest future avenues of research.

## 2 Related Literature

The identity inducement phase of the current experiment draws mainly on the body of literature in social psychology referred to as the “minimal group paradigm” (MGP). In MGP experiments, investigators randomly divide subjects into two groups; group affiliation constitutes a social identity. Each subject then divides a sum of money between two *other* subjects, knowing only the social identity of these two other subjects. A robust finding of hundreds of MGP experiments involving induced informationless social identities is that subjects allocate significantly more money to members of their own group—i.e. in-group

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<sup>13</sup>The “c” refers to social category; using “i” for social identity would be more natural, but also more confusing in terms of notation.



	Money Initially Sent			
	5	10	15	20
<u>Avg Money Returned</u>				
by Ashkenazic:	1.8	13	17.2	24.3
by Eastern:	2.8	14.2	16.7	23.1

Table 1: Trust Game (Fershtman and Gneezy, 2001)

bias.<sup>14</sup> For a meta-analysis of dozens of MGP experiments, see Mullen, Brown and Smith (1992). For a more recent overview of this research, see Haslam (2001). And for a more critical survey, see Brown (2000).

Of course, the MGP experiments are too cute by half. In particular, self-interest is not involved since the money-division tasks are disinterested. This leaves open the possibility that social identity effects are sufficiently weak to be drowned out whenever self-interest *is* implicated—in which case, social identity could be safely ignored by economists.

The existing experimental economics literature provides evidence concerning both the strength and nature of social identity effects when self-interest *is* implicated. Two of the clearest such clues appear in Glaeser, Laibson, Scheinkman and Soutter (2000) (GLSS) and Fershtman and Gneezy (2001) (FG). Both experiments involved mostly-standard Trust Games. Not coincidentally, GLSS provides clues about the effects of two equal-status identities; while FG provides clues concerning the effects of two unequal-status identities.<sup>15</sup>

In FG, the subjects were Ashkenazic and “Eastern” Jews. Ashkenazic Jews can be thought of as having higher status, as they “achieve higher levels of education and earnings than do Eastern immigrants. . .” [17, p. 352]. An unexplained pattern in the data was that Ashkenazic Jews (high status) both trusted and reciprocated trust more than Eastern Jews (low status). The reciprocity patterns are clear in Table 1 below (reproduced from Table 2 (pg 363) in FG): when sent a low amount, Ashkenazic Jews return less than Eastern Jews; however, when sent a high amount, they return *more* than Eastern Jews.

<sup>14</sup>When using real-world identities, the result is less robust. There are several intervening factors that have been explored by social psychologists, with status differences being among the earliest and most-researched.

<sup>15</sup>I say “clues,” rather than evidence, because neither of these experiments had the primary purpose of investigating social identity, and therefore neither have experimental designs which can rule out compelling alternative explanations.

Empirical evidence concerning such a high status/high standards phenomenon is present in Guiso, Sapienza and Zingales (2003). Analyzing survey data concerning religion and economic attitudes and controlling for income and various other demographic characteristics, the authors found a significant positive relationship between self-assessed social class and self-reported levels of trust; and a significant negative relationship between self-assessed social class and support for such non-normative behaviors as cheating on taxes and improperly claiming government benefits.

In the social psychology literature, variants of the high status/high standards phenomenon can be found as far back as Homans' analysis (1950, p. 141) of a famous and even earlier quasi-experiment, the Bank Wiring Observation Room (details appear, among other places, in Roethlisberger and Dickinson (1939).) Additionally, many of the minimal group experiments involving status are also consistent with such an hypothesis if in-group bias is taken to be the normative behavior. See, e.g., any of the overviews mentioned above, or, for the seminal work, see Turner and Brown (1978).

In GLSS, two groups of subjects involved were Asian students and white students. Being drawn from the same (Harvard) economics class, these groups likely have equal status. One pattern in the data was that subjects were less trustworthy when paired with a subject of different ethnicity. Importantly, the pattern was not "members of certain ethnic groups are less trustworthy," but rather, the implied pattern was that the *same* person could be more or less trustworthy depending on the ethnicity of their co-player.

As intriguing as these clues are, there is an obvious confound in both the experimental and empirical evidence: group affiliations can convey decision-relevant information. Specifically, the informational confound in GLSS stems from the fact that subject pairings were *not anonymous*. Additionally, patterns of association may make a person of the same ethnicity more likely to share experimental earnings outside of the experiment. In FG, group affiliation clearly conveyed information about future earnings, which could have been a factor in second-movers' decisions. Real-world identities such as social class an obviously also convey information.

The current experiment recreates the essence of the social identities in FG and GLSS, while controlling for the informational confound. Following the minimal group paradigm, I randomly assigned subjects to one of two groups. In one version of the experiment the two groups were of equal status (GLSS); in a second version of the experiment, I randomly chose one of the groups to have “high status” (FG).

A closely related paper in experimental economics that controls for the informational confound is Ball, Eckel, Grossman and Zame (2001). The authors publicly randomly assigned status levels and conducted a two-sided auction, finding that prices tended to favor high status subjects: goods traded at higher prices when sellers had high status and lower prices when *buyers* had high status. This is consistent with the high status/high standards hypothesis—it could result, for instance, if buying high and selling low was the norm. However, the authors explain their results with the assumption that agents prefer to interact with higher status agents, and the design of the experiment does not allow one to distinguish between these two explanations.

Finally, the current experiment is closely related to many experiments that have failed to find a significant effect of induced informationless social identities in the lab. The reasons are varied, but many are due to intentionally weak identity inducement procedures (e.g., Güth Levati and Ploner 2005). In a nutshell, the difficulty in identifying pure social identity effects seems to be finding an identity inducement procedure strong enough to induce a sense of shared social identity but weak enough to not raise other other serious confounds. Several experiments illustrating this balance are discussed in McLeish and Oxoby (2007).

### **3 Experimental Design**

The subjects in the experiments were recruited from undergraduates and staff at the University of California, Berkeley. During the course of the experiment, each subject played from ten to fifteen rounds of a standard Trust Game as well as ten to fifteen rounds of a costless signalling game. The experiment was programmed and conducted with the soft-

ware z-Tree (Fischbacher 2007), and all sessions were conducted in the X-lab facilities at the University of California, Berkeley. All together, eight sessions were conducted and 144 subjects participated.

The experiment consisted of two phases: an initial identity-inducement phase, followed by a game-playing phase. There were two possible versions of the identity inducement phase: the identity-only version (ID-only), and the status and identity version (S-ID). The game-playing phase was identical in each version of the experiment. Each subject participated in only one of the two possible versions.

### **3.1 Identity Inducement Phase**

In the ID-only version of the experiment, subjects were randomly assigned one of two colors—purple or orange—and then seated on one side of the room or the other, according to color group. Randomization was achieved by publicly drawing poker chips out of a canvas bag, and it was made clear that purple and orange were equally likely to be drawn. Each color can be thought of as an identity. Subjects were always separated from their immediate neighbors with opaque dividers, essentially creating an individual and private cubicle for each subject.

Each color group’s seating arrangement was the same, and the groups faced each other across approximately five feet of empty space in the middle of the room. After they were seated, I handed each subject a wristband matching their assigned color, which they were instructed to wear for the duration of the experiment. Each subject’s own color-group remained the same for the duration of the experiment. After this, subjects proceeded directly to the game-playing phase.

In the S-ID version, subjects were randomly divided into color groups in exactly the same manner as in the ID-only version. And subjects were, again, given a wristband corresponding to their color group and asked to wear it during the experiment. The difference between the S-ID and the ID-only version was that in the S-ID version, one of the colors groups was publicly randomly chosen to signify “high status,” while the non-chosen color

signified “low status.” As in the ID-only version, each subject’s own color-group—and hence their status level—remained the same for the duration of the experiment.

Status was assigned by putting one purple poker chip and one orange poker chip in a canvas bag, shaking the bag in full view of all subjects, then drawing one poker chip out of the bag. The drawn poker chip was held up by the experimenter, after which it was announced that the color of the poker chip drawn out of the bag would represent “high status” for the duration of the experiment. Status differences were reinforced by varying groups’ comfort levels and assigning status-specific tasks.

To make high status subjects more comfortable than low status subjects, high status subjects were seated three per row, while low status subjects were seated five per row. In terms of tasks, high status subjects were allowed to enjoy refreshments while low status subjects worked.<sup>16</sup> Specifically, low status subjects were assigned a boring and tedious task—taking a list of names alphabetized by last name and re-alphabetizing the list by first name, by hand. Each low status subject was provided with their own list of names, and the task was completed individually without communication with, or cooperation from, other group members. This phase lasted 10 minutes, after which materials were collected and subjects proceeded directly to the game-playing phase.

Out of a total of eight experimental sessions conducted, three of these sessions were ID-only and five sessions were the S-ID version. Overall, 60 subjects participated in the ID-only version of the experiment, while 84 subjects participated in the S-ID version.

### **3.2 Game-Playing Phase**

The game-playing phase was identical in both versions of the experiment: all subjects played multiple rounds of the Trust Game and multiple rounds of the Truth Game. Before each round each subject was randomly paired with a co-player, so that all pairings potentially changed in every round. Furthermore, pairings were not restricted by color group, so that a player could be paired with an “orange” player one round and a “purple” player the

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<sup>16</sup>The refreshments were modest, consisting of a glass of water or lemonade as well as some crackers, cheese and grapes.

next round. Save for the information about one’s partner’s color group, pairings were anonymous.<sup>17</sup>

The Trust Game is a two-player sequential game of perfect information. In this game, first-movers are called “senders” while second-movers are called “receivers.” Senders are endowed with \$7 while receivers are endowed with \$0. The game begins with the sender deciding to send some, all or none of his or her endowment to their anonymous co-player. Each dollar sent is then tripled by the experimenter, and allocated to the receiver. Finally, the receiver then decides how much of this tripled amount to send back. Any money not sent back is kept by the receiver. All purely self-interested subgame perfect equilibria of the Trust Game involve the receiver returning nothing, and hence, the sender sending nothing as well.

The Truth Game—a two-player costless-signalling game—can be thought of as modeling the situation faced by a used car salesman and a prospective buyer. The salesman has private information—whether the car is reliable or a lemon. After observing this information, the salesman can send one of two messages to the buyer: he can either tell the buyer that the car is reliable, or that it’s a lemon. The buyer observes the seller’s message and then can take one of two actions—buy or walk away. The buyer prefers buying the car only if it’s reliable, otherwise walking away is the buyer’s preferred action. The salesman, on the other hand, prefers the buyer to buy irrespective of the car’s quality. Payoffs in the Truth Game depend solely on the quality of the car and the buyer’s decision. Monetary payoffs are given in Table 2 and the game tree can be found in the appendix. One feature of the Truth Game is that it has no purely self-interested equilibria in which sellers’ messages convey any information. Subjects’ instructions for the Truth Game, as well as the Trust Game, can be found in the appendix.

In total there are 760 observations for each game in the data. There are 300 observations for each game stemming from the ID-only version of the experiment, and 460 observations

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<sup>17</sup>Even though subjects were seated by color group, there were many people in each color-group in each session, and rounds of game-play went fairly quickly (ten rounds of each game in addition to 15 minutes of setup took approximately one hour) so that it is unlikely players could have deduced who their (different) partner was in each round.

	Buyer Action	
	Buy	Walk Away
Actual Quality		
Reliable	(12, 12)	(10, 10)
Lemon	(12, 10)	(10, 12)

(a,b) = (\$ Seller, \$ Buyer)

Table 2: Truth Game Payoffs

for each game come from the S-ID version of the experiment. Since these totals reflect multiple rounds of game-play, steps were taken to minimize any possible dynamic effects. After each round of each game, subjects were randomly (and anonymously) re-paired and roles within each game were randomly reassigned. For example, a particular subject could be a buyer in the Truth Game one round and a seller the next, and be paired with a high status co-player in one round and a low status co-player in the next round. Furthermore, subjects were never informed of the outcomes of any of the rounds of either game. These steps appear to have been successful: overall, subjects' actions in early rounds of each game did not differ significantly from their actions in later rounds. Therefore, all reported results incorporate aggregating over all rounds within each game.

## 4 Results

### 4.1 The Trust Game

#### 4.1.1 Summary Statistics

As in previous research on the Trust Game, senders exhibited a substantial amount of trust. Similar to the original Trust Game experiment (Berg, Dickhaut and McCabe, 1995), senders sent about half the maximum possible amount. While the average amount sent was similar in the ID-only and S-ID versions of the current experiment, a Kolmogorov-Smirnov test suggests that the *distribution* of amounts sent in the ID-only version was significantly different from the distribution of amounts sent in the S-ID version ( $p = 0.007$ ).

The key difference between the distributions of amounts sent is the propensity to send all or nothing. Introducing status seems to have swayed fence-sitters, making both complete trust and complete lack of trust more likely at the expense of middling trust levels. A summary of the distributions is presented in Table 3.

Proportion of Maximum Sent		
	<u>ID-only</u>	<u>S-ID</u>
Average Proportion Sent	0.51	0.55
Standard Error	(0.022)	(0.019)
10th percentile	0	0
20th percentile	0.07	0
30th percentile	0.21	0.14
40th percentile	0.36	0.43
50th percentile	0.50	0.57
60th percentile	0.71	0.86
70th percentile	0.86	1
80th percentile	1	1
90th percentile	1	1
N	300	460
1. Maximum possible was \$7, so reported value is $\frac{AmountSent}{7}$ .		

Table 3: Trust Game Senders, proportions sent

For their part, receivers returned close to 80% of the amount senders’ sent—excluding observations where senders initially sent zero.<sup>18</sup> Again, this is in line with receivers’ decisions in Berg, Dickhaut and McCabe (1995), where the average proportion returned was 89.5%.

Although a Kolmogorov-Smirnov test suggests that the distributions of proportions returned varied across versions ( $p = 0.058$ ), this fact does not have much meaning by itself as it could be an artifact of variation in senders’ actions across versions. One way to get a sense of whether, and how, receivers actions varied between the ID-only and S-ID versions of the experiment is to divide the data into two categories—one in which the amount initially sent was above the median, and another in which it was below the median. Using these categories, we see that when faced with a low amount sent, the most generous receivers—i.e.,

<sup>18</sup>Since money sent was tripled, the proportion returned can take values from 0 to 3.



the top 10%—in the ID-only version were more generous than the most generous receivers in the S-ID version. On the other hand, when faced with a high amount sent, receivers in the S-ID version were more likely to display high levels of generosity than receivers in the ID-only version.

Kolmogorov-Smirnov tests fail to reject the null hypothesis of equal distributions when initial amounts sent were less than the median; while, for observations where the initial amount sent was above the median, the distributions were significantly different ( $p = 0.009$ ). A summary of the distributions of proportions returned by receivers is presented in Table 4.

	Proportion Returned					
	Overall		$\leq$ Median Sent		$>$ Median Sent	
	ID-only	S-ID	ID-only	S-ID	ID-only	S-ID
Average Return Proportion	0.77	0.78	0.55	0.54	0.92	0.92
Standard Error	(0.045)	(0.039)	(0.075)	(0.064)	(0.054)	(0.048)
10th percentile	0.00	0.00	0.00	0.00	0.00	0.00
20th percentile	0.00	0.00	0.00	0.00	0.14	0.00
30th percentile	0.14	0.00	0.00	0.00	0.62	0.29
40th percentile	0.43	0.33	0.00	0.00	1.00	0.71
50th percentile	0.75	0.67	0.25	0.25	1.00	1.00
60th percentile	1.00	1.00	0.43	0.50	1.14	1.29
70th percentile	1.14	1.30	0.67	0.83	1.29	1.43
80th percentile	1.33	1.43	1.00	1.00	1.42	1.50
90th percentile	1.50	1.50	1.90	1.50	1.50	1.50
N	249	355	102	131	147	224

1. Since money sent was tripled, the return proportion can take values from 0 to 3.
2. ID-only median sent = 3.50; S-ID median sent = 4.00.
3. Includes only observations where money initially sent was greater than zero.

Table 4: Trust Game Receivers, proportion returned

#### 4.1.2 Patterns in Senders' Actions

In both versions of the experiment, the evidence of in-group bias on the part of first-movers was weak, but consistent. Each group sent more, on average, to members of their own

Proportion of Maximum Sent						
	ID-only			S-ID		
	overall	to in-group	to out-group	overall	to in-group	to out-group
Overall	0.513 (0.022)	0.539 (0.031)	0.487 (0.031)	0.553 (0.019)	0.576 (0.027)	0.531 (0.027)
High Stat Sender				0.609*** (0.026)	0.651 (0.037)	0.569 (0.037)
Low Stat Sender				0.495 (0.028)	0.500 (0.039)	0.490 (0.040)

1. Reported value is  $\$Sent$  as a proportion of maximum (i.e.,  $\frac{\$Sent}{7}$ )
2. Standard errors in parentheses
3. Significance comparison is with respect to low-status senders in S-ID, and overall ID-only senders.

Table 5: Trust Game Senders' Actions

group than to members of the other group. And, in line with previous research in the minimal group paradigm, the introduction of status increased the magnitude of in-group bias among high status senders and decreased it among low status senders, relative to the ID-only version of the experiment. However, none of these patterns rose to the level of statistical significance.<sup>19</sup>

This is not to say social identity had no significant effects, however. Pronounced patterns in senders' actions are present in the S-ID version, where unequal status was introduced. Consistent with the hypothesis that high status senders in the Trust Game held themselves to higher standards, high status senders trusted their co-players significantly more on average. Overall, high status senders were significantly more trusting than low status senders ( $p < 0.01$ ); and significantly more trusting than their ID-only counterparts. This "extra" trust was extended to both high and low status receivers, as high status senders sent more to both types of S-ID receivers than their low status counterparts. These patterns are evident in Table 5.

<sup>19</sup>Although the in-group bias among high status senders becomes significant ( $p = 0.012$ ) when individual sender fixed-effects are incorporated; in-group bias remains non-significant among low status subjects.

### 4.1.3 Patterns in Receivers' Actions

Considering that trust is the norm in the Trust Game—hence, the name—the role of the receiver is largely to reward and punish trusting behavior.<sup>20</sup> The most pronounced patterns in the data are related to this reciprocal role. In both the ID-only and S-ID versions of the experiment there was significant identity-related variation in reciprocity.

The baseline case is the ID-only version of the experiment. For Trust Game receivers in the ID-only version of the experiment, to investigate reciprocity directly I estimated receivers' average return ratios as a function of the amount initially sent to them by their co-players. I used return ratios as the dependent variable in the estimation to avoid much of the heteroskedasticity associated with estimating return *amounts* directly.

Specifically, I estimated return ratio functions separately for subjects involved in same-color pairings and subjects involved in different-color pairings. Wald tests suggested the average return ratio function associated with same-color pairings was linear, while the function associated with different-color pairings was quadratic, in money initially sent. Estimating return ratios using these functional forms—the estimates can be found in the appendix—revealed a significant difference between how in-group members and out-group members were treated: ID-only receivers held members of their *own* social identity to a higher standard. Notice in Figure 1 that when an in-group member sent a low amount they were punished more harshly than if they were not a member of the receiver's social identity; and when an in-group member sent a very high amount, they were rewarded more generously than if they were not a member of the receiver's social identity.

That is, subjects were willing to spend money in order to provide incentives for members of their own group to live up to their ideals. In the real world, such a pattern is consistent with giving to one's church for aid programs, rather than directly to poor people, since church programs are typically both targeted to fellow church members, and often come with incentives to improve one's behavior.

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<sup>20</sup>Punish and reward in the sense of leaving the sender worse off, or better off, than they started. This is possible because sent amounts are tripled, meaning return ratios can vary between 0 and 3. A return ratio of 1 leaves the sender exactly as well off as they started.

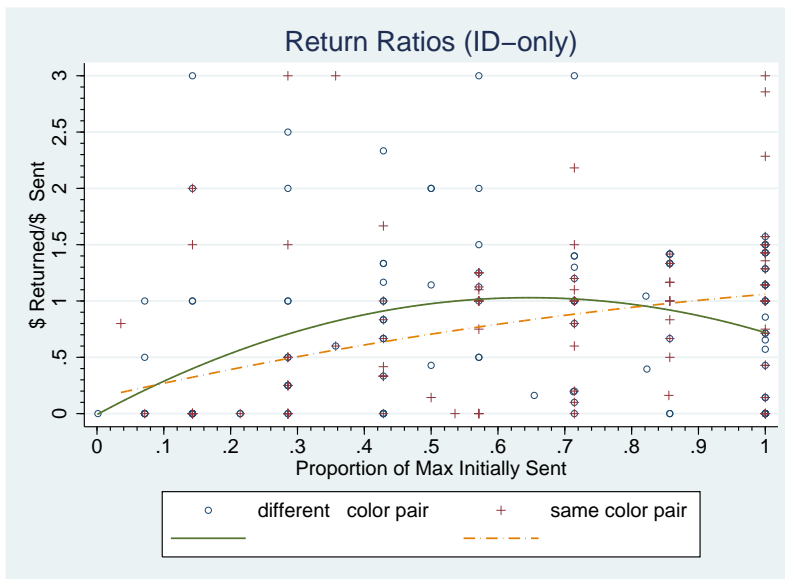


Figure 1: Return Ratios, ID-only

Next, consider the version of the experiment where a status difference was introduced between groups. In the S-ID version of the experiment, high status receivers appeared to hold everyone to a higher standard. To show this, I again estimated receivers' return ratio functions. To pin down a reasonable functional form, I separately estimated return ratio functions for high status and low status receivers as quintic functions of the amount initially sent. Wald tests failed to reject the null hypothesis that these return ratio functions were in fact linear.<sup>21</sup>

Given this, Table 6 presents linear estimates of S-ID subjects' return ratio functions. The estimates reveal that high status receivers acted in a more reciprocal manner: compared to low status receivers, high status receivers' return ratios were more than twice as responsive to the amount initially sent—in some of the estimations, high status receivers were more than three times as reciprocal by this measure. Furthermore, the patterns proved robust to including individual receiver random effects; to adjusting the variance estimates by clustering on individual receivers; and to accounting for censoring by using a Tobit estimation procedure. The latter two estimates are presented in the appendix.<sup>22</sup>

<sup>21</sup>This was true for Wald tests on all subsets of powers of amount received greater than one as well.

<sup>22</sup>Further still, accounting for session fixed effects as an added specification test did not change the patterns

Dependent Variable = Return Ratio				
	(1)	(2)	(3)	(4)
Constant	0.451*** (0.144)	0.399*** (0.134)	0.513** (0.199)	0.420** (0.177)
$\$Sent$	0.056** (0.027)	0.046** (0.020)	0.043 (0.035)	0.046* (0.028)
High Status Receiver	-0.355** (0.174)	-0.459*** (0.172)	-0.447* (0.228)	-0.426* (0.222)
High Stat Rec $\times \$Sent$	0.086*** (0.033)	0.118*** (0.025)	0.106** (0.042)	0.116*** (0.034)
High Stat Sender			-0.146 (0.288)	-0.045 (0.205)
High Stat Sender $\times \$Sent$			0.030 (0.054)	0.000 (0.040)
H.S. Sender $\times$ H.S. Rec'r			0.229 (0.362)	-0.132 (0.268)
H.S. Sender $\times$ H.S. Rec'r $\times \$Sent$			-0.047 (0.068)	0.016 (0.050)
Receiver Random Effects	no	yes	no	yes
N	355	355	355	355
$R^2$	0.107	0.106	0.108	0.103

1. Dependent variable, return ratio, can take values from 0 to 3.
2. Robust standard errors in parentheses.
3. Estimates include only observation where  $\$Sent > 0$ .

Table 6: Trust Game receivers, S-ID

As is evident in Table 6, the senders' status level was irrelevant to both high and low status receivers. This makes it more difficult to explain the results by simple distributional equity concerns. That is, one obvious explanation for the results is that the status manipulation simply introduced an unequal wealth distribution, since the low status group "earned less" by not receiving refreshments. However, receivers apparently ignored whether they were returning money to a "poorer" subject.

That high status subjects reward trust, and punish lack of trust, to a greater degree than low status subjects is clear in Figure 2. Also evident is one economically relevant consequence of the difference in reciprocity patterns. From a purely pecuniary standpoint, trusting a low status subject was a losing proposition: low status subjects' average return

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qualitatively.

ratios were below the break-even value of 1 no matter how much was sent. On the other hand, it was possible to earn a positive return on one’s “investment” by trusting high status receivers. Such a pattern has potentially important dynamic effects, as, over time, subjects could learn to only trust high status individuals. Since I tried to eliminate dynamic effects in this experiment, however, I cannot presently test this intuition.

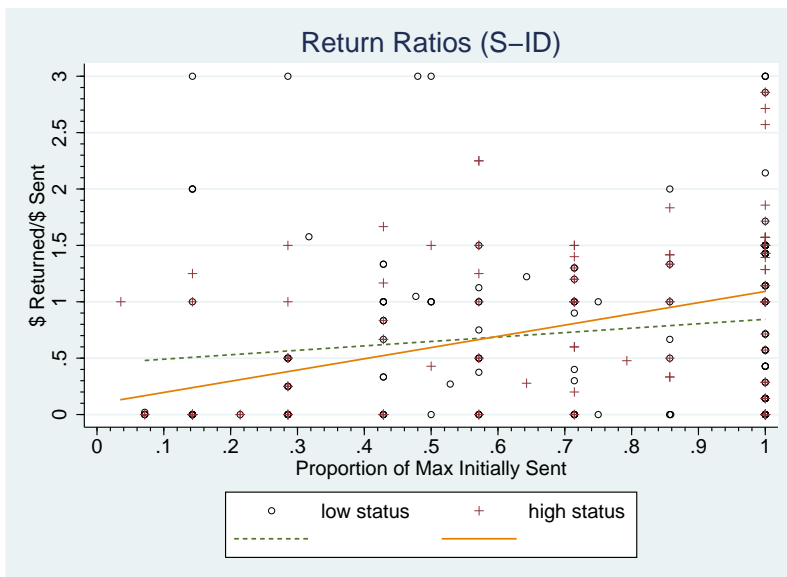


Figure 2: Trust Game Receivers, S-ID

Comparing reciprocity patterns across the experimental conditions yields insights into the working of status. High status S-ID subjects held *all* subjects to a higher standard in the same way that ID-only subjects held members of their *own* identity to higher standards. At the same time, low status S-ID receivers were more reluctant to punish and reward trust in the same way that ID-only receivers were reluctant in punishing and rewarding members of the *other* identity.

In fact, these comparisons can be made more precise. Statistically, high status receivers in the S-ID version are indistinguishable from receivers involved in same-color pairings in the ID-only version: comparing the estimated return ratio functions for these two subsets of the data yielded no significant differences. That is to say, a Wald test failed to reject the null hypothesis of identical estimated return ratio functions ( $p = 0.6460$ ). In the same

respect, low status S-ID receivers' behavior was statistically indistinguishable from ID-only subjects engaged in *different-color* pairings ( $p = 0.7488$ ). Furthermore, it is *not* the case that the data are just too noisy to distinguish among return ratio functions in any of these subsets: high status receivers' estimated return ratio functions *were* significantly different from the the return ratio functions of ID-only subjects engaged in different-color pairings. A Wald test rejected the null hypothesis that return ratio functions were identical in these two cases ( $p = 0.0259$ ).<sup>23</sup>

Putting these comparisons together, it seems as though status worked as follows: high status emboldened subjects to impose their values on *everyone* in the same way that subjects in the ID-only version were emboldened in their interactions with in-group members. Low status, on the other hand, makes subjects more reserved in their judgments: low status S-ID receivers withheld both punishment and reward in the same way ID-only receivers were less judgmental when dealing with members of the “other” group. In short, high status makes individuals more confident that their way is the right way, while low status is humbling.

## 4.2 The Truth Game

Next, I'll examine subjects' behavior in the Truth Game data in isolation. Afterwards, I'll draw connections between the patterns present in both games, and explain them in a social identity framework.

### 4.2.1 Summary Statistics

On the salesman side, the distribution of messages was surprisingly similar across versions—as Table 7 shows. Also, there was about a five percentage point increase in honesty in the S-ID version for each quality level: S-ID sellers are more likely to send the message “reliable” when the car is actually reliable; and more likely to send the message “lemon” when the car is actually a lemon.

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<sup>23</sup>Reported Wald test results are when all return ratios were estimated as linear. The Wald tests were also conducted after all return ratios were estimated as quadratic in *\$Sent*. The significance patterns did not change.

Proportion of Messages Sent				
	ID-only		S-ID	
	“reliable”	“lemon”	“reliable”	“lemon”
<u>Car Quality</u>				
Overall	0.790	0.210	0.796	0.204
Reliable	0.896	0.104	0.942	0.058
Lemon	0.705	0.295	0.657	0.343

Table 7: Truth Game Sellers’ Messages, overview

Buyers’ Propensity to Buy/Walk				
	ID-only		S-ID	
	Buy	Walk	Buy	Walk
<u>Message</u>				
Reliable	0.751	0.249	0.770	0.230
Lemon	0.333	0.666	0.255	0.745
Overall	0.663	0.337	0.665	0.335

Table 8: Truth Game Buyers, overview

For their part, buyers apparently conditioned their decisions on what sellers had to say, even though sellers’ messages must reveal nothing about car quality in any purely self-interested equilibrium. In both versions of the experiment, using a Kolmogorov-Smirnov test, we can reject the idea that the distributions of buyers’ actions were the same for both possible messages ( $p < .01$ ). Buyers’ actions are summarized in Table 8.

#### 4.2.2 Truth Game Sellers

There is evidence supporting the high status/high standards hypothesis among Truth Game Sellers. For sellers, the norm is clear: a “decent” person *should* tell the truth. Therefore, if high status sellers hold themselves to a higher standard, we would expect high status sellers to be more honest than low status sellers. In fact, this was true.

In the S-ID version of the experiment, high status sellers were significantly more honest than low status sellers. Furthermore, as in the Trust Game, this “extra” honesty extended across versions: high status S-ID sellers were also more honest than ID-only sellers (Table 9).

Further still, high status sellers in the S-ID version of the experiment were more honest



Sellers' Propensity to Tell the Truth				
Buyer Status	ID-only	S-ID		
		Overall	High	Low
Overall	0.563 (0.029)	0.635** (0.022)	0.683*** (0.031)	0.590 (0.032)
High		0.644 (0.031)	0.699 (0.045)	0.603 (0.042)
Low		0.624 (0.033)	0.669 (0.043)	0.573 (0.049)

1. Standard Errors in parentheses
2. Statistical significance is with respect to ID-only
3. High status vs low status sellers overall:  $p = 0.038$

Table 9: Truth Game Sellers' Honesty

irrespective of their private information. They were more honest when they knew (privately) the car was reliable, and more honest when they knew the car was a lemon. Considered separately, the former pattern even rose to the level of statistical significance (Table 10).

One aspect of Table 10 is puzzling at first glance: why would anyone lie when the car is actually reliable?<sup>24</sup> After all, sellers' and buyers' preferences are aligned in this instance. One would think that "holding oneself to a higher standard" should imply more honesty in the face of bad news. And while this pattern *is* present, it would be reassuring if this bit of extra honesty were statistically significant.

The answer lies in the equilibria of the Truth Game. There is, in fact, an equilibrium where lying about good news is rational behavior: the incredulous equilibrium. The gist of this equilibrium is that if sellers think buyers will think sellers are always lying, then self-interested sellers will lie about good news and tell the truth about bad news. Specifically, in this equilibrium sellers with good news lie; sellers with bad news tell the truth; the only message that buyers ever hear, therefore, is bad news so buyers learn nothing from sellers' messages.

There is some evidence in the data suggesting incredulous equilibria play a role in the

<sup>24</sup>While one obvious explanation for this would be inexperience with the game, the fact that this pattern persists in all rounds suggests something else was at play.

Honesty Propensity, by Private Information		
Seller Status	Car Quality	
	Reliable	Lemon
High	0.991*** (0.009)	0.361 (0.046)
Low	0.892 (0.030)	0.328 (0.042)
ID-only	0.896 (0.027)	0.295 (0.036)

1. Standard errors in parentheses  
2. Statistical significance comparison is within-column

Table 10: Truth Game Sellers’ Honesty

observed propensity to lie about good news. To see this, label each sender who lied at least once when they knew (privately) they had a good car for sale a “cynical seller.” Among the observations in which sellers had private information that the car was a lemon, truth-telling was significantly *more* frequent if the observation involved a cynical seller ( $\chi^2(1) = 11.3314, p = 0.001$ ). And among the observations in which the seller had private information that the car was reliable, truth-telling was much *less* likely if the observation involved a cynical seller: 100% by definition among non-cynical sellers vs. 60.9% among cynical sellers. This is exactly the pattern we’d expect if some of these subjects were engaged in the incredulous equilibrium: honesty after bad news and dishonesty after good news.

Furthermore, there is an apparent status bias in seller cynicism. In the S-ID data, of the 10 senders and 65 observations involved, only 1 sender—involved in 5 observations—was high status. Apparently, low status sellers expected to be disbelieved and developed a crude strategem to get their way.

Now, excluding all observations involving cynical sellers is clearly too blunt an instrument—as, for instance, this eliminates all the variation in honesty after good news. However, it is interesting to note that even this blunt tool reveals evidence in favor of the “high status/high standards” hypothesis: excluding cynical sellers, high status subjects are significantly more honest about bad news than both low status sellers and ID-only sellers (Table 11). Such a

Honesty about “Bad News”		
	Seller Status	
ID-only	High	Low
0.266	0.364*	0.257
(0.037)	(0.047)	(0.044)

1. Standard errors in parentheses

Table 11: Honesty, Excluding “Cynical Sellers”

a pattern is more along the lines of what most people would consider “holding yourself to a higher standard.”

### 4.2.3 Truth Game Buyers

It’s not clear what a norm for Truth Game buyers would be. And, since buyers in the Truth Game have no real punishment opportunity, it’s not clear what would constitute evidence of holding others to a higher standard. Therefore, one wouldn’t expect buyers’ actions to change much across versions—and they don’t (Table 8). But the Truth Game buyers still provide valuable insights. Specifically, I’ll use the Truth Game buyers’ data to highlight one consequence of introducing status (S-ID) relative to a “different-but-equal” (ID-only) environment: buyers unambiguously benefitted from the introduction of unequal status.

First, consider the purely self-interested equilibria of the Truth Game. In all self-interested equilibria of the Truth Game, buyers learn nothing from sellers’ messages. Hence, we would expect buyers to choose the correct action—buy reliable cars and walk away from lemons—no more frequently than random chance allows (here, 50 percent of the time). This is exactly what happened among ID-only buyers—there was no evidence that buyers learned anything from sellers’ messages. However, S-ID buyers *were* significantly more likely to choose the correct action than random chance would suggest (Table 12). Thus, the introduction of unequal status improved the functioning of this simple used-car market.<sup>25</sup>

How did this improvement happen? Notice in Table 12 that, overall, S-ID buyers did

<sup>25</sup>As a specification check, I tested whether it was senders’ messages that receivers conditioned their actions on, or if somehow senders’ private information was revealed irrespective of messages—indicating a flaw in the experimental design. A multinomial logit estimation revealed no significant effects of the actual private information once messages were controlled for. The estimation appears in the appendix.

Buyers' Propensity to Choose "Correct" Action				
	<u>ID-only</u>	<u>S-ID</u>		
		Buyer Status		
<u>Seller Status</u>		Overall	High	Low
Overall	0.490 (0.029)	0.600*** (0.023)	0.594 (0.032)	0.606 (0.033)
High Stat Seller		0.624 (0.033)	0.641 (0.048)	0.610 (0.045)
Low Stat Seller		0.577 (0.032)	0.559 (0.043)	0.602 (0.048)

1. Standard errors in parentheses
2. Indicated significance is with respect to overall ID-only
3. All S-ID values except (low-seller+low-buyer) are significantly different from .5 at  $\geq 95\%$  confidence level

Table 12: Buying Reliable Cars and Walking Away from Lemons

not fare significantly worse against low status sellers. Coupling the overall credulity evident among buyers in Table 8 with low status sellers' pronounced dishonesty, this seems odd. One obvious answer is that sellers somehow anticipated the honesty differential between high status and low status sellers. This seems to be exactly what happened: in the S-ID version of the experiment, overall, buyers were significantly more likely to believe messages from high status sellers than from low status sellers. This pattern in belief held for both high status and low status buyers, considered separately (Table 13).

Furthermore, a chi-square test failed to reject the null hypothesis that the observed distribution of messages was the same for both high and low status senders ( $\chi^2(1) = 1.43, p = 0.232$ ), suggesting that this pattern of belief cannot be explained by high status sellers simply sending "more believable" messages. Further still, the belief patterns are not about simply being nicer to low status sellers by buying more frequently from them: ignoring messages, separate chi-square tests for high status and low status buyers fail to reject the null hypothesis that buyers are equally likely to buy from low and high status sellers ( $p > 0.80$  in both cases).

At the same time, ID-only buyers correctly gleaned the *lack* of variation in honesty. ID-only buyers were (correctly) no more likely to believe in-group members than out-group

Buyers' Propensity to Believe Messages			
S-ID			
Buyer Status			
<u>Seller Status</u>	Overall	High	Low
High	0.814** (0.026)	0.825 (0.037)	0.805 (0.037)
Low	0.720 (0.029)	0.721 (0.039)	0.718 (0.045)

1. Standard errors in parentheses  
2. Statistical significance is within-column

Table 13: Truth Game Buyers' Credulity (S-ID)

members (Table 14).

Sellers' Honesty and Buyers' Credulity		
ID-only		
<u>Type of Pairing</u>	Honesty Propensity (Sellers)	Belief Propensity (Buyers)
Same Colors	0.553 (0.041)	0.713 (0.037)
Different Colors	0.573 (0.041)	0.753 (0.035)

1. Standard errors in parentheses

Table 14: Honesty and Credulity (ID-only)

Considered as a whole, patterns in Truth Game buyers' actions provide evidence about the value of unequal status to the functioning of markets with asymmetric information and no verification opportunities. Buyers correctly anticipated the patterns in seller honesty implied by the introduction of status, and were able to use this inference to their advantage, thereby increasing the likelihood of mutually beneficial trades regardless of who the seller was.

## 5 A Social Identity Explanation

Recall that in a social identity framework agents put themselves and others into social categories. Social categories are identities. Identities prescribe ideals—specifically, how

agents *should* act. Since agents care about living up to their ideals and ideals are enmeshed with categorizations, in this type of model the *same* person can behave as if they are maximizing a different utility function when social identities change, even though nothing else decision-relevant has changed.

Such a framework is consistent with the observed patterns in the Truth Game: heuristically, adding a concern for honesty to individuals' utility functions transforms the Truth Game from a *costless* signalling game into a costly signalling game. Identity concerns make signals involving a lie more “costly” than signals involving truth-telling. Whenever identity makes lying sufficiently costly—which will be a function both of subjects' ideals, and how much weight they put on these ideals vis-a-vis pecuniary incentives—truth-telling can be supported in equilibrium.<sup>26</sup> Letting the weight subjects place on normative concerns be completely idiosyncratic, more truth-telling by high status subjects is consistent with high status subjects having ideals which prescribe a higher propensity to tell the truth.

To get beyond such a heuristic argument, however, and actually estimate subjects' ideals in a simple identity model requires a less complicated strategic environment than the Truth Game provides. Therefore, to demonstrate variation in ideals more concretely, I take advantage of the simplicity of the strategic situation faced by receivers in the Trust Game: Trust Game receivers are essentially dictators, deciding how much of a fixed sum to share with their partner, so their decision of how much to return is purely a choice between ideals and self-interest. I further restrict my attention to Trust Game receivers in the S-ID version of the experiment, since this is where I have the most, and cleanest, data.

The social identity model I constructed for Trust Game receivers' utility is as follows. Denoting senders' actions by  $s$ , receivers' actions by  $r$  and money payoffs by  $x$ , agent  $j$ 's overall utility is given by Equation 2, where  $\alpha_j$  is a parameter capturing how much agent  $j$  cares about her identity relative to pecuniary incentives:

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<sup>26</sup>In Sanchez-Pages and Vorsatz (2007), the authors make this argument more formally, and, using a very similar cheap talk game, provide evidence that the “tension between normative social behavior and incentives for lying” explain the “excessive” truth-telling found in their experiment. I had no knowledge of this paper when constructing the current experiment.

$$U_j = x_j(r, s) - \alpha_j(r - r_{c_j}^{Ideal}(s))^2 \quad (2)$$

In Equation 2, agent  $j$  cares about both her money earnings— $x_j(r, s)$ —and the distance between her actual action,  $r$ , and her *ideal* action,  $r_{c_j}^{Ideal}(s)$ . Agent  $j$ 's ideal, in turn, depends on  $s$ —to incorporate reciprocity—as well as her social category,  $c_j$ . Of course, agent  $j$ 's ideal could also depend on her co-player's social category, but since there was no evidence of this in the S-ID version of the experiment, this possibility is not modeled.

While this model of utility might look unfamiliar, in fact, it can be thought of as a generalization of a widely-used form of social preferences: inequity aversion (Fehr and Schmidt, 1999). To see this, consider a non-linear variant of the standard Fehr and Schmidt formulation of inequity aversion, which is also a close cousin to the prototypical example used in Bolton and Ockenfels (2000):

$$U_j = x_j(r, s) - \alpha_j(x_j(r, s) - x_i(r, s))^2 \quad (3)$$

Thus, agent  $j$ 's cares about both her money earnings, and how unequal the distribution of earnings is. In the specific Trust Game used in the experiment, receivers' money earnings are given by  $3s - r$  and senders' earnings are given by  $7 - s + r$ . Plugging these facts into (3) and simplifying yields:

$$U_j = 3s - r - \alpha_j(4s - 2r - 7)^2 \quad (4)$$

From Equation 4, simple algebraic manipulation allows one to re-cast the model of inequity aversion given by Equation 3 in an identity-utility form:

$$U_j = x_j(s, r) - \tilde{\alpha}_j(r - r^{Ideal}(s))^2 \quad (5)$$

In Equation 5,  $r^{Ideal}(s) = 2s - \frac{7}{2}$  and  $\tilde{\alpha} = 4\alpha$ .<sup>27</sup> Thus, inequity aversion can be thought

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<sup>27</sup>It looks rather out of place, but the 4 in  $\tilde{\alpha}$  is an artifact of factoring  $-2$  out of  $(4s - 2r - 7)$  to get the expression in the parentheses,  $(4s - sr - 7)$ , into an  $r - r^{Ideal}(s)$  format. This is also where the  $\frac{7}{2}$  term in

of as an identity model with the additional assumption that ideals are constant across social identities, coupled with a specific prediction about what receivers' ideals are. Because of this connection, I will use as my null hypothesis the idea that the data in the current experiment are well-explained by inequity aversion.

Proceeding with the estimation of subjects' ideals, recall that in the current context, receivers' money payoffs in the Trust Game were  $3s - r$ . After plugging this fact into the identity model of Equation 2, first-order conditions imply that receivers' optimal (interior) money-return rule was given by:

$$r_j^*(s) = r_{c_j}^{Ideal}(s) - \frac{1}{2\alpha_j} \quad (6)$$

Thus, estimating a subjects' average optimal return rule,  $r^*(s)$ , is equivalent—up to an unknown constant—to estimating subjects' ideals. To simplify matters, I assumed that ideals were linear in  $s$  and that there was precisely one ideal for each status level— $r_L^{Ideal}(s)$  for low status receivers, and  $r_H^{Ideal}(s)$  for high status receivers. All individual-level heterogeneity, then, comes from the  $\alpha_j$ 's. To further simplify matters, I assumed that all  $\alpha_j$ 's were drawn from the same underlying distribution—as would be expected if this parameter captures a stable, individual trait. Call this random variable  $\alpha$ .

The most straightforward way to estimate  $r^*$  is to use Equation 6 directly as a regression function, with  $\frac{1}{2\alpha_j}$  serving as an error term. Since  $r^*$  was possibly censored—whenever  $r^*$  falls below zero, I observed zero<sup>28</sup>—and it's not clear what distribution the error term should have, I used a semi-parametric estimator that accounts for censoring and is robust to a wide range of error distributions: Censored Least Absolute Deviations (CLAD) (Powell, 1984).<sup>29</sup> It is sufficient, for instance, to assume that  $\frac{1}{2\alpha}$  is a well-defined random variable with a

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$r^{Ideal}$  comes from.

<sup>28</sup>Censoring from above is also possible, but not quite as worrying as there are very few observations in the data where the maximum possible amount was returned.

<sup>29</sup>The main error-term assumption required for CLAD to be consistent is that errors have median zero, which is quite a bit less restrictive than the standard assumption of normality and homoskedasticity. The tradeoff is that an assumption must be made about the data—roughly speaking, there must be “enough” uncensored observations. As will be clear from the estimation, this is likely to be satisfied in the present case.



Dependent Variable = Dollars Returned				
	CLAD (1)	CLAD (2)	Tobit (1)	Tobit (2)
Constant	-0.60 (0.688)	-0.67 (0.789)	-2.18* (1.145)	-3.53** (1.378)
$\$Sent$	0.80*** (0.297)	0.67** (0.322)	0.96*** (0.225)	1.219*** (0.253)
High Status Receiver	-6.90*** (2.369)	-3.33 (2.805)	-2.76 (1.687)	-2.62 (2.063)
H. S. Rec'r $\times \$Sent$	1.70*** (0.541)	1.33** (0.609)	0.70** (0.316)	0.72* (0.376)
High Stat Sender		-0.33 (1.726)		-0.13 (2.100)
H.S. Sender $\times \$Sent$		0.33 (0.612)		0.07 (0.379)
H.S. Rec'r $\times$ H. S. Sender		1.00 (3.998)		0.52 (3.237)
H.S. Rec'r $\times$ H. S. Sender $\times \$Sent$		-0.67 (0.789)		-0.194 (0.568)
$N_{initial}$	355	355	355	355
$N_{final}$	302	315	355	355
Pseudo $R^2$	0.187	0.216	0.090	0.068

1. Standard errors in Parentheses  
2. Estimates include only observation where  $\$Sent > 0$

Table 15: Optimal Return functions, S-ID

unique finite median.

The CLAD estimates of the optimal return functions are presented in Table 15. The estimated return functions are significantly different across status levels—allowing us to rule out inequity aversion as an explanation. In addition to CLAD, Table 15 includes Tobit estimations for comparison. However, the regularity conditions necessary for the consistency of maximum likelihood estimation are not obviously satisfied, so the Tobit estimates may be biased.

Since none of the controls related to senders' status are significant—neither individually nor jointly—I will focus on the simplest CLAD estimates from here on. Call the estimated optimal return rule among high status subjects  $\widehat{r}_H^*$ , and define  $\widehat{r}_L^*$  analogously. Estimating receivers' ideals is a matter of simply shuffling terms from one side of Equation 6 to the other. Specifically:

$$\widehat{r}_H^{Ideal} = \widehat{r}_H^* + \frac{1}{2\bar{\alpha}} \quad (7)$$

$$\widehat{r}_L^{Ideal} = \widehat{r}_L^* + \frac{1}{2\bar{\alpha}} \quad (8)$$

Plugging the CLAD estimates,  $\widehat{r}_H^*$  and  $\widehat{r}_L^*$ , into Equations 7 and 8 yields the estimated ideals:

$$\widehat{r}_H^{Ideal} = 2.5s - 7.5 + \frac{1}{2\bar{\alpha}}$$

$$\widehat{r}_L^{Ideal} = 0.8s - 0.6 + \frac{1}{2\bar{\alpha}}$$

Notice that high status subjects' ideals are not terribly different from what would be expected from inequity-averse agents. Further, recall that high status subjects act just like subjects in the ID-only version of the experiment engaged in in-group pairings. This suggests one reason for the frustrating lack of external validity characteristic of many previous experimental investigations of trust and reciprocity: merely bringing subjects into the lab succeeded in creating a shared social identity.

## 5.1 An Interpretation of the Estimated Ideals

One way to think of the receiver's role in the Trust Game is as an enforcer of normative behavior—punishing deviations from the norm and rewarding norm conformance. With this in mind, and denoting the sender's norm with  $s^{Ideal}$ , we can re-write receivers' estimated ideals in a particularly simple format (Equation 9). Here,  $\gamma$  captures a base-line level of generosity, while  $\beta$  measures concern for others living up to their ideals:

$$r^{Ideal}(s) = \beta(s - s^{Ideal}) + \gamma \quad (9)$$

To make this more concrete, suppose that in the Trust Game all senders *should* exhibit full trust—i.e.,  $s_H^{Ideal} = s_L^{Ideal} = 7$ . Then the estimated ideals can be re-written in an

especially tidy manner:

$$r_H^{Ideal}(s) = 2.5(s - s^{Ideal}) + (10 + \frac{1}{2\bar{\alpha}}) \quad (10)$$

$$r_L^{Ideal}(s) = 0.8(s - s^{Ideal}) + (5 + \frac{1}{2\bar{\alpha}}) \quad (11)$$

Here we see receivers’ true colors. High status receivers care much more about others’ norm conformance—punishing senders by more than two dollars for every dollar senders fall short of their ideal. But they are also much more generous when it is warranted. This last point can be seen by considering the ideal return amount when senders exactly conform to their norm. In this case,  $(s - s^{Ideal}) = 0$ , implying that high status receivers (ideally) reward senders much more lavishly than low status receivers.

## 5.2 Ruling out Wealth Effects and Mood Effects

One obvious alternative explanation for the patterns in the data is some type of wealth effect: it could be that the status manipulation in the S-ID version of the experiment simply made the “high status” subjects feel wealthier. The exact effect of extra wealth would depend on the model used, but heuristically one might expect the “wealthier” subjects to be simply more generous. While the increased generosity of high status subjects when sent a large initial amount is consistent with such a wealth effect, the relative *decrease* in generosity observed when the “wealthier” subjects were sent low amounts seems contrary in spirit to the idea that high status subjects simply feel wealthier.<sup>30</sup> Furthermore, it would be difficult to explain the connection between the ID-only patterns and the S-ID patterns with wealth effects alone since there was no systematic “wealth” difference created by the identity inducement phase in the ID-only version of the experiment.

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<sup>30</sup>Pushing this intuition a bit further, if the proposed wealth effect resulted in an approximately uniform increase in generosity, we should see “wealthier subjects” return a fixed amount more to their co-player for each amount initially sent than their poorer counterparts. Since this constant extra generosity would have greater impact on return *ratios* for low return amounts, this uniform extra generosity should imply *flatter* return ratio functions for high status/wealthier subjects. This was exactly the *opposite* of the observed patterns.

Another possible explanation for the patterns in the data is a mood effect. One might think that the manipulation I used to reinforce status simply put half the subjects—the high status subjects—in a better mood. And, there is some experimental research in economics suggesting that mood might have an effect on “other-regarding” behavior. I again, however, have a few reasons for believing mood effects do not explain the results. Firstly, I will reiterate that social identity had a significant impact even in the ID-only version of the experiment where it’s less clear that moods varied. Secondly, while clean results are hard to come by with respect to mood’s impact on behavior, there is one experiment in this vein that provides compelling evidence that mood does not explain the current results. In Kirchsteiger, Rigotti and Rustichini (2001), the authors induced two different moods—good or bad—by having subjects watch either a funny or a sad movie, respectively. Subjects then played a standard gift-exchange game.<sup>31</sup> The authors found that subjects in a good mood are *less* reciprocal and *more* generous than players in a bad mood. That is to say, players in a *bad* mood mirror high status subjects in the present experiment; and players in a *good* mood mirror low status subjects.<sup>32</sup> Thus, mood differences are likely to have worked *against* the reciprocity patterns in the current paper rather than providing an alternative explanation for them.

## 6 Conclusion

On the most basic level, I have shown that subjects’ ideals vary with social identity, rather than being stable *individual* traits. Beyond this, the results provide evidence for specific effects of simple social identities that are reasonable and ring true: we hold our own people to higher standards; and high status emboldens us to hold everyone—including ourselves—to higher standards.

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<sup>31</sup>A gift exchange game is similar to the Trust Game. There are two players. One player moves first and chooses how much of a fixed sum of money to transfer to his or her co-player—player 2. Player 2 observes the amount transferred and decides how much “effort” to exert. Effort decreases player 2’s earnings but increases player 1’s earnings. Reciprocity is measured as the responsiveness of the effort decision to the transfer decision. Generosity is measured by the amount of effort exerted when the initial transfer is zero.

<sup>32</sup>They also mirror, by the way, the hypothetically “wealthier” subjects outlined above.

The social identity effects above tie together results across unrelated experiments—trustworthiness patterns in Glaeser, Laibson, Scheinkman and Soutter (2000) and Fershtman and Gneezy (2001), as well as price patterns in Ball, Eckel, Grossman and Zame (2001). At the same time, the results provide one plausible factor—the unintended creation of a shared social identity—in the frustrating lack of external validity associated with many social preferences experiments in economics.

Additionally, the current experiment raises many questions for future research. First of all, the social identities induced in this experiment were assigned rather than chosen, and status was assigned rather than earned. This opens the question of whether the effects of social identity and status generalize to earned status and/or chosen social identities.

Secondly, the current experiment was designed to shut down dynamic effects. Opening up the investigation of social identity to such dynamic effects raises potentially interesting questions. Most directly, patterns in the Trust Game suggest that subjects may learn, over time, to trust only high status individuals, as completely trusting high status individuals was the only way to earn a positive expected return-on-investment.

Another class of questions raised by the current experiment relates to multiple social identities. In this experiment I attempted to induce unique identities. How individuals prioritize multiple social identities, particularly when these identities have conflicting ideals, is an obvious extension that would provide a closer analogy to the real world.

Finally, the current experiment dictated the associational patterns: with whom subjects interacted was (randomly) assigned. However, one might suspect that, for instance, the patterns in punishing norm-deviance might have a significant impact on with whom subjects *choose* to interact when these associational decisions are voluntary. To the extent that social networks are economically important, such associational biases may have significant economic consequences, and are thus worthy of investigating.

## Appendix

### A Instructions

*The following instructions appeared on each subject's computer screen immediately before each game was started:*

#### A.1 The Trust Game

In each round of the following game, you will be randomly paired with a co-player and will be randomly assigned one of two roles: **Sender** or **Receiver**.

Role of Sender:

The **Sender** will be given \$7.00 and may choose to send any amount of this money to the **Receiver**. The amount chosen by the **Sender** will be tripled and given to the **Receiver**.

Role of Receiver:

The **Receiver**, upon learning how much money is available to him or her, will be able to send any amount of that money back to the **Sender**.

Earnings per round:

The **Sender** earns  $7.00 - (\text{dollars sent to Receiver}) + (\text{dollars sent back by Receiver})$ ;

The **Receiver** earns  $3 * (\text{dollars sent by Sender}) - (\text{dollars sent back to Sender})$ .

When you've read and understand the Instructions, click on the button to proceed to the game.

#### A.2 Truth Game

**Now we begin a new game.**

As before, in each round of the following game, you will be randomly matched with a co-player and randomly assigned one of two roles: **Sender** or **Receiver**.

Role of Sender: The **Sender** will learn the result of a coin flip performed by the computer. This coin flip has equal chances of resulting in "Heads" or "Tails." The

**Sender** then sends a message about this coin flip to the **Receiver**.

Role of Receiver: The **Receiver**, does not learn the result of the coin flip. The **Receiver's** only information about the result is the message sent by the **Sender**. Once the **Receiver** receives this message from the **Sender**, the **Receiver** chooses an action: *Left* or *Right*. The **Receiver's** action, together with the result of the coin flip, completely determines both the **Sender's** and the **Receiver's** earnings for the round.

Earnings per Period: If the result of the Coin Flip is Heads, then the action Left yields earnings of \$12 for the Receiver and \$12 for the Sender. The action Right yields earnings of \$10 for the Receiver and \$10 for the Sender.

If the result of the Coin Flip is Tails, then the earnings for the action Left are \$10, \$12 for Receiver, Sender. And the earnings for the action Right are \$12,\$10 for Receiver, Sender.

This is summarized in the following table, where:  $\$a, \$b = \$a$  earnings for Receiver,  $\$b$  earnings for Sender.

Coin/Action	Left	Right
Heads	\$12, \$12	\$10, \$10
Tails	\$10, \$12	\$12, \$10

**Click the button when you're ready to proceed to the game.**

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## Tables and Figures

Dependent Variable = Return Ratio				
Variable	(1)	(2)	(3)	(4)
Constant	0.451** (0.193)	0.399*** (0.134)	0.513* (0.259)	0.420** (0.186)
$\$Sent$	0.056** (0.026)	0.046** (0.022)	0.043 (0.036)	0.046* (0.033)
High Status Receiver	-0.355 (0.219)	-0.459*** (0.170)	-0.447 (0.286)	-0.426* (0.219)
High Stat Rec $\times \$Sent$	0.086** (0.035)	0.118*** (0.031)	0.106** (0.046)	0.116*** (0.040)
High Stat Sender			-0.146 (0.244)	-0.045 (0.222)
High Stat Sender $\times \$Sent$			0.030 (0.053)	0.000 (0.044)
H.S. Sender $\times$ H.S. Rec'r			0.229 (0.327)	-0.132 (0.275)
H.S. Sender $\times$ H.S. Rec'r $\times \$Sent$			-0.047 (0.065)	0.016 (0.052)
Receiver Random Effects	no	yes	no	yes
N	355	355	355	355
$R^2$	0.107	0.106	0.109	0.103

1. Dependent variable, return ratio, can take values from 0 to 3
2. Robust standard errors in parentheses
3. Errors adjusted to include intra-receiver variation (clustering)
4. Estimates include only observation where  $\$Sent > 0$

Table A.1: Trust Game receivers, clustered, S-ID

Dependent Variable = Return Ratio				
Variable	(1)	(2)	(3)	(4)
Constant	0.052 (0.189)	0.049 (0.175)	0.074 (0.248)	0.072 (0.213)
$\$Sent$	0.098*** (0.034)	0.074*** (0.023)	0.091* (0.046)	0.077** (0.031)
High Status Receiver	-0.609** (0.286)	-0.683*** (0.254)	-0.723* (0.370)	-0.690** (0.303)
High Stat Rec $\times \$Sent$	0.131** (0.051)	0.160*** (0.034)	0.158** (0.068)	0.163*** (0.046)
High Stat Sender			-0.050 (0.380)	-0.012 (0.259)
High Stat Sender $\times \$Sent$			0.015 (0.069)	-0.008 (0.049)
H.S. Sender $\times$ H.S. Rec'r			0.285 (0.585)	-0.051 (0.383)
H.S. Sender $\times$ H.S. Rec'r $\times \$Sent$			-0.060 (0.104)	0.003 (0.070)
Receiver Random Effects	no	yes	no	yes
N	355	355	355	355

1. Dependent variable, return ratio, can take values from 0 to 3
2. Standard errors in Parentheses
3. Estimates include only observation where  $\$Sent > 0$

Table A.2: Trust Game receivers, Tobit, S-ID

Dependent Variable = Return Ratio		
Variable	(1)	(2)
Constant	-0.009 (0.199)	-0.009 0.211
\$ Sent	0.459*** (0.121)	0.459*** (0.126)
$(\$Sent)^2$	-.051*** (0.015)	-.051*** (0.016)
Same-Color Pairing	0.149 (0.306)	0.149 (0.258)
Same-Color Pairing $\times$ \$Sent	-0.267 (0.174)	-0.267* (0.152)
Same-Color Pairing $\times$ $(\$Sent)^2$	0.042** (0.021)	0.042** (0.019)
Receiver Clustering	no	yes
N	249	249
$R^2$	0.126	0.126

1. Dependent variable, return ratio, can take values from 0 to 3  
2. Robust standard errors in Parentheses  
3. Estimates include only observation where \$Sent > 0

Table A.3: Trust Game receivers, ID-only

Dependent Variable = Buyer's Action		
	Logit (1)	Logit (2)
Message	-1.83*** (0.595)	-2.17*** (0.824)
Car Quality	-0.340 (0.250)	-0.466 (0.354)
Quality $\times$ Message	-0.37 (0.675)	-0.248 (0.890)
Constant	1.36*** (0.171)	—
Rec'r Fixed Effects	no	yes
N	355	355
Pseudo $R^2$	0.151	0.289

Table A.4: Specification Test: Truth Game Buyers

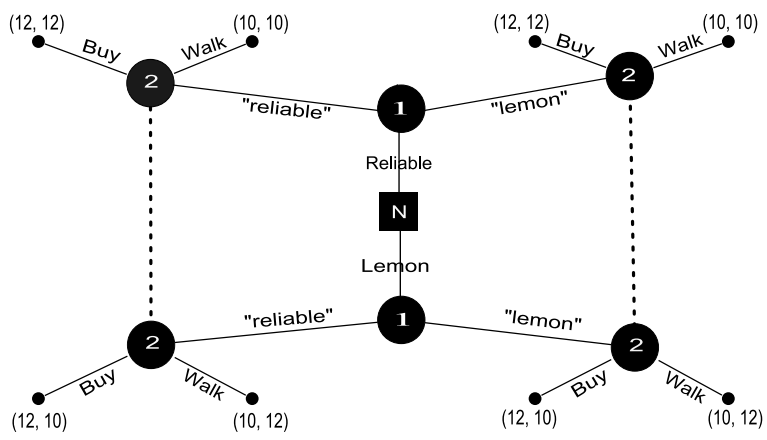


Figure A.1: Truth Game, game tree