# Leaving Boys Behind: Gender Disparities in High Academic Achievement<sup>1</sup>

By Nicole M. Fortin, Department of Economics, UBC Philip Oreopoulos, Department of Economics, University of Toronto Shelley Phipps, Department of Economics, Dalhousie University and CIFAR

August 2011

<sup>&</sup>lt;sup>1</sup> We would like to acknowledge Lori Timmins for her outstanding research assistance on this project. We would also like to thank Mario Small, Thomas Lemieux, and participants at the CIFAR SIIWB Workshop and at the CEA 2011 for helpful comments on this and earlier versions of the manuscript. We thank ICPSR and MTF for allowing us to use the data, and the usual disclaimer applies. The authors are grateful for CIFAR's financial support. Fortin also acknowledges funding from SSHRC Grants #410-2011-0567.

# Leaving Boys Behind: Gender Disparities in High Academic Achievement

## Abstract

Using three decades of data from the "Monitoring the Future" cross-sectional surveys, this paper shows that, from the 1980s to the 2000s, the mode of girls' high school grade distribution has shifted from B to A, essentially "leaving boys behind" as the mode of boys' grade distribution stayed at B. We appeal to a "signaling" type model of academic achievement to argue that educational expectations should play a prominent role, given a level of aptitude, in an individual's choice of optimal GPA. In an extended Oaxaca-Blinder decomposition of academic achievement, we do indeed find the most important factor accounting for the growing gender gap in high academic achievement are gender differences in post-secondary expectations, in particular the growing proportion of girls who aim for a post-graduate degree. We also find that high achieving girls are "swimming upstream" relatively unfavorable family environments.

Keywords: Academic achievement, educational expectations, gender differences, high school grades, boys' problem.

JEL codes: I210 (Analysis of Education), J16 (Economics of Gender), J24 (Human Capital)

# 1. Introduction

Women now far outnumber men among recent college graduates in most industrialized countries (OECD, 2008). Their growing dominance in education attainment raises new questions about gender disparities arising throughout school-ages.<sup>2</sup> Girls have long obtained better grades, on average, in high school than boys.<sup>3</sup> As shown in Figure 1, the average gender gap in GPA among high school seniors (scaled out of 4 points) hovers steadily around 0.2 between 1976 and 2009.<sup>4</sup> But because historically these achievements never translated into higher levels of educational attainment or better labor market outcomes for women relative to men, earlier research has instead concentrated on explaining the remaining gaps in women's performance, particularly in mathematics (e.g. Guiso et al., 2008; Bedard and Cho, 2010, Niederle and Vesterlund, 2010). Conversely, the relative underperformance of males, especially in reading, has attracted little attention until recently (LoGerfo, Nichols, and Chaplin, 2006).

Interest in the academic performance gap favoring women is changing for a number of reasons. As the manufacturing sector is shrinking, the "good jobs" that men once obtained without a post-secondary education are becoming harder to find. The new gender enrollment imbalance in some university disciplines, such as medicine, has generated concerns about future labor shortage in some specialties that attract few women, such as orthopedic surgeons, and has lead to reverse affirmative action. Repercussions for the marriage prospects of college-educated women (Blank, 2011), and concerns among boys' parents about a "failure to launch" (Bell, Burtless, Gornick and Smeeding, 2007) are among the other preoccupations behind the growing interest for the issue. There are now numerous books and articles in the popular press discussing the relative underperformance and under motivation of boys, and their late entry into adulthood.<sup>5</sup>

The goal of this paper is to first document changes in gender disparities in academic performance of high school students (12<sup>th</sup>, 10<sup>th</sup>, and 8<sup>th</sup> graders) over the last three decades using

 $<sup>^{2}</sup>$  According to OECD (2008), the average share of the student population in tertiary education in OECD countries accounted for by women reached 55% in 2005. Only four countries are likely not to achieve at least parity between men and women by 2015: Korea, Turkey, Japan and Switzerland.

<sup>&</sup>lt;sup>3</sup> This is observed in other countries as well. See Machin and McNally (2005) for Britain, Lai (2010) for China.

<sup>&</sup>lt;sup>4</sup> The gender gap in GPA from the MTF match (within standard errors) the numbers from the NAEP High School Transcript Study for 1990, 2000, 2005 and 2009, also reported in NCES (2004), as well as the numbers reported in Cho (2007) for 1984 from the HS&B survey.

<sup>&</sup>lt;sup>5</sup> See among others, Tyre (2008), Gurian and Stevens (2007), Sax (2008). By contrast, popular books in the 1990s were concerned about girls being disadvantaged by school system (e.g. Sadker and Sadker, 1994; AAUW,1992).

survey data from the "Monitoring the Future" (MTF) project.<sup>6</sup> We show, in particular, that the progressive erosion of "grading on the curve" has resulted in an increasing proportion of girls among the very top achievers. As the percentage of 12<sup>th</sup> grade students reporting in the MTF that they earn A's (93-100%) almost doubles, from 8.5% in the 1980s to 16.6. % in the 2000s, the difference between the fraction female to the fraction male in this category also doubles from 3.2% to 5.4%.<sup>7</sup> From 1990s to the 2000s, the female advantage in the proportion of 10<sup>th</sup> and 8<sup>th</sup> graders earning A's also increases, from 3.6% to 4.8% and from 4.9% to 5.5%, respectively.

The second aim of the paper to identify the relative importance of three sets of determining factors, which changed over time, in accounting for the growing gender disparities in high academic achievement.<sup>8</sup> These include family environment, working while in school, and future aspirations. The data does not allow us to consider the effect of the teachers' gender (Dee, 2005, 2006) or the effect of teaching styles (Algan, Cahuc, and Shleifer, 2010) which have attracted recent attention.<sup>9</sup> But some studies (Kramarz, Machin, and Ouazad, 2008) suggest that these effects are smaller in magnitude than those of student ability or family background variables, which we include in our analysis. Moreover, to the extent that teaching styles may be linked the teacher's gender, CPS data shows that the proportion female (ranging from 55% to 59%) among secondary school teachers has changed very little over the time period considered.

As first observed by Ben-Porath and Welch (1976) boys and girls are raised in different family environments. Our study also shows important differences in family composition. Families with girls are, on average, larger (in line with Angrist and Evans, 1998), have less educated parents, more absent fathers (Dahl and Moretti, 2008), and more working mothers, with the last two gender gaps in family characteristics increasing over time.<sup>10</sup> These time trends would seem to disadvantage girls, suggesting that high achieving girls are actually "swimming upstream" perhaps reacting to more difficult circumstances. Another changing factor is the decreasing labor force participation of boys during high school; among high school seniors, participation went from 85% in the 1980s to the 76% in the 2000s, when about three-quarters of

<sup>&</sup>lt;sup>6</sup> To the best of our knowledge, Jacob and Wilder (2010) is the only other paper using the MTF to study educational expectations. They focus on their impact on educational attainment (college going).

<sup>&</sup>lt;sup>7</sup> In the MTF, an A grade corresponds to a percentile grade in the 93-100% range. The exact years are 1976 to 1988 for the 1980s, and 2000 to 2009 for the 2000s for 12<sup>th</sup> graders, and 1991-1999 for the 1990s for 10<sup>th</sup> and 8<sup>th</sup> graders. <sup>8</sup> See Buchmann, DiPrete, and McDaniel (2008) for a review.

<sup>&</sup>lt;sup>9</sup> We are however able to assess the importance of the type of high school (academic, general, vocational, etc.).

<sup>&</sup>lt;sup>10</sup> See Lundberg and Rose (2002) on the gendered effects of children on fathers' labor supply and wages.

both boys and girls had some sort of paid employment during school.<sup>11</sup> To the extent that negative effects of working during school are expected for a wide range of high school students (Tyler, 2003; Rothstein, 2007), the time trends should prevent boys from falling behind in high school grades. On the other hand, working during school in entry level jobs could play a role in motivating students to pursue higher education and may thus have different effects across the GPA distribution.

Finally and most importantly, another set of factors that changed over the last three decades are the post-secondary aspirations and expectations of high school students, as well as their choice of high school (from vocational to academic) to enact these career plans. Figure 2a shows the expectations facet of the well-known reversal in the gender gap in college enrollment analyzed by Goldin, Katz and Kuziemko (2006). It shows that from 1976 to 1983, boys and girls had similar expectations about going to college, but that a gap progressively opened thereafter stabilizing around 12% in the 2000s. Figure 2b presents that gender gap expressed as the proportion of females among students who say that they will definitively go to a four year college. Among 12<sup>th</sup> graders, that proportion was around parity up to the early 1980s, overshooting actual enrollment rates by a few percentage points in the 1990s to stabilize around 57% in the 2000s. Interestingly, the gender gap in expectations to attend a four year college is shown to emerge as early as grade 8, when it hovers around 55%. Goldin and Katz (2002) have argued that the 1970s "Pill Revolution" was crucial in allowing young women to formulate plans for higher education without the fear of interruptions for family reasons. We argue that in subsequent decades, the ongoing progress of women in the professions has continued to fuel young women's career plans involving graduate and professional schools. At the same time, with the advent of computerization and other office technologies, in the 2000s clerical work seems to have lost much of its appeal for young women.<sup>12</sup>

Our set of expectation variables includes a full range of career plans for life after high school, including vocational and graduate school as well as serving in the military. As explained below, the expectation variables are thought to subsume the effects of gender difference in the returns to college (Jacob, 2002; Cho, 2007) or the anticipated dispersion regarding those returns (Charles and Luoh, 2003). As pointed out by Manski (2004), when using expectations data, it is

<sup>&</sup>lt;sup>11</sup> However as shown below, boys continue to work longer hours and earn more money.

<sup>&</sup>lt;sup>12</sup> See Table A1.

important to clarify the assumptions held about the information used to form of these expectations. In context of educational expectations, whether youths condition these expectations on returns to college, ability or financial constraints could lead to different interpretations of the results. Fortunately, the MTF surveys also include data on educational aspirations and subjective school ability which allows us to condition, at least partially, for some potentially endogenous effects, and to present bounds on the effects of educational expectations with and without these controls.

As with most studies of changes in gender differentials, we cannot appeal to an experimental or quasi-experimental setting to explore how an individual assigned a gender at random would behave and how his/her environment would respond.<sup>13</sup> We have to construct counterfactual states of the world based on the observed responses and respective endowments of males and females. We then apply decomposition methodologies (Fortin, Lemieux, and Firpo, 2011) aimed at separating endowment effects from response effects under the assumption that the distribution of unobservables conditioning on observables is independent of gender. The goal of the exercise is to assess the relative importance of the potential explanatory factors in order to identify the ones that may warrant a more careful investigation of a causal link. We focus on a distributional approach to analyze changes over time in GPA because average gender differences in GPA have not changed over the past thirty years, while the gender ratio of students admitted to college, those with high GPA, has changed dramatically. In so doing, we also contribute a distributional understanding to increases in average GPA over time.<sup>14</sup>

The paper is organized as follows. Section 2 presents the simple "signaling" type model of academic achievement that motivates the estimation procedure. Section 3 introduces the MTF surveys and presents some descriptive statistics about gender disparities in academic achievement and in the explanatory factors, as well as changes over time therein. The decomposition methodologies are explained in Section 4. Section 5 presents the decomposition results and their interpretation. Finally, section 6 concludes.

<sup>&</sup>lt;sup>13</sup> Notable exceptions are audit studies where the gender of the individual is hidden (Goldin and Rouse, 2000), and the authors study how the environment responds without knowing the gender of the individual.

<sup>&</sup>lt;sup>14</sup> By contrast, grade inflation should refer to changes in the price of grades (e.g decrease in study time). Because they compress the grade distribution, rising grades are different from rising nominal prices. The term "grade inflation" may be an imperfect analogy.

### 2. A Simple Model of High School Performance

We begin by presenting a simple model of academic performance to illustrate that, despite the fact that girls have long had better school ability than boys, it is their changing aspirations for the future that led girls to capture a larger proportion of high grades. The model also helps rationalize the relative underperformance of boys as the consequence of career choices that require lower levels of educational attainment. Models of high school performance in economics are usually set as derivatives of the Mincerian human capital investment model where individuals choose their level of schooling to maximize their life-cycle earnings. Here we want think of decisions taken earlier in life when labor market outcomes are not as concrete or narrow as returns to college, but would come out as the answer to the typical question: "What do you want to do when you grow up?"<sup>15</sup> At this stage, parents are likely still involved in the education of their children, perhaps actively assisting them with homework, helping them set goals, and manage their time. The model thus borrows from the model of intergenerational transmission of income status (Becker and Tomes, 1979), the idea that the other generation's utility, the parents in this case, enters the decision maker's objective function. The parents' utility from their offspring educational expectations depends on their own characteristics, such as their own level of education, as well as on the school ability or aptitude of each child. Over the last three decades, the exogenous changes in the opportunities for women in the labor market have led many parents to have higher educational expectations for their daughters. As shown in Chen, Fortin, Phipps, and Oreopoulos (2011), in the 2000s, parents of primary school students had higher educational expectations for their daughters than for their sons.

As in signaling models of educational choices (Spence, 1973), we want to allow for the fact that, given a level of aptitude, an individual may find it optimal to aim for the minimum GPA needed to reach a career or educational goal. Implicit in our framework is the fact that basic school ability is revealed quite early in the pupil's schooling experience and that the level of aptitude of the individual figures more prominently in educational/career choices than it does in human capital investment models.<sup>16</sup> This also contrasts with learning models (Stange, 2008)

<sup>&</sup>lt;sup>15</sup> For example, DiPrete and Buchmann (2006) suggest that gender differences in college enrollment may arise because women's marriage market possibilities and social status may be enhanced by college going above and over the simple returns to college.

<sup>&</sup>lt;sup>16</sup> There is an emerging consensus in the psychology literature that the grade level at which students form reliable perceptions of academic competency is around 5<sup>th</sup> grade (Herbert and Stipek, 2005).

where academic ability for college is revealed slowly over time and where individuals revised their schooling decisions. We think that both models are adequate representation of the behavior of some subsets of individuals. The updating of educational expectations is perhaps more salient among college/university students who face more fateful choices about which major to pursue or whether or not to pursue their studies (Zafar, 2011) than among high school students for whom the salient choice is once they reach minimum school leaving age. Despite updating by some individuals, results in Stange (2008) and Jacob and Wilder (2010) support the idea that a large proportion (40%-66%) of individuals is actually successful at enacting their early educational plans. More importantly, because we use cross-sectional data, the model is set in a static framework.<sup>17</sup>

A pupil comes to secondary school with a basic aptitude for school  $(A_i)$  that was largely revealed during elementary school and with plans for future education likely influenced by parental desires, and aims for a GPA  $(G_i)$  in the range  $R_j(G_i) = I[l_j \le G_i \le u_j]$ , where *I* is the indicator function, that will allow him/her to pursue further schooling/career plans  $(S_j, j =$ 1, ..., *J*). As illustrated in Figure 3, this simplified functional form for  $R_j(G_i)$  assumes that with probability one that a student with a GPA in the indicated range will be able to pursue her education plans.<sup>18</sup> The student chooses a level of effort and target GPA to maximize her utility from schooling plans minus costs,  $C(G_i)$ , plus an intergenerational utility component,

$$Max_{\{G,E\}} U_i\{S_j - C(G_i)\} + \alpha U_i^p(S_j, A_i),$$
  
subject to  $S_j(G_i) = I[l_j \le G_i \le u_j], \quad j = 1, ..., J$   
 $C(G_i) = f(E_i, A_i).$ 

The component  $U_i^p(S_j, A_i)$  represents the utility to the parents of having an offspring of ability  $A_i$ in reach of educational level  $S_j$ , and  $\alpha$  is the weight placed by the student on parental utility. This last parameter is potentially important is assessing gender differences, as psychologists argue that girls place more importance on pleasing adults than boys. Assuming separability of schooling plans and costs, however, parental utility merely acts to scale the rewards of a

<sup>&</sup>lt;sup>17</sup> We do not exclude the possibility that some students revise their plans, but because we do not have access to the MTF longitudinal data, we cannot explore this avenue.

<sup>&</sup>lt;sup>18</sup> In reality, the discontinuities do not need to be as sharp as illustrated in Figure 3.

schooling plan. The effect of parents' other characteristics might have similar rescaling effects on the cost function.<sup>19</sup>

Importantly, the cost of getting a particular grade,  $C(G_i) = f(E_i, A_i)$ , is decreasing nonlinearly with ability,  $\frac{\partial f(E_i,A_i)}{\partial A_i} < 0$ , and  $\frac{\partial^2 f(E_i,A_i)}{\partial A_i} > 0$ . The cost of academic achievement is increasing with effort, but there may be some complex non-linear interactions between effort and ability, possibly different by gender, that we do not attempt to model directly here, but leave for future research.<sup>20</sup> The mapping  $R_j(G_i)$  of GPA into educational plans may include a more complex step function than the one above, where there are different probabilities of attaining educational choice  $S_j$  by GPA level.<sup>21</sup> What is important in leading some students to optimally choose lower levels of GPA are the discontinuities in access to educational choice by GPA, as shown in Figure 3.

In Figure 3, the utility of three educational choices for student *i*,  $U_i(S_j) = w_i * I[l_j \leq G_i]$ , are displayed as simple step functions, for *j*=two-year college ( $w_i = 4, l_{2-yr} = 2$ ), four-year college ( $w_i = 7.5, l_{4-yr} = 3$ ), and graduate school ( $w_i = 10, l_{grad} = 4$ ). The cost functions illustrated in Figure 3, subsumed in their functional forms the level of effort needed for high, medium and low ability students to achieve that higher GPAs, showing that it is are more costly lower ability students to obtain high GPAs. Thus the choices of GPA,  $G_i^*(A_i, S_j)$ , which maximize the utility net of achievement cost for each ability level, are the lower bound of each educational choice. That is, the low ability student will target a GPA of 2 to access two-year colleges, the medium ability a GPA of 3 to access four-year colleges, and the high ability student a GPA of 4 aiming to attend graduate school. Letting  $G_i^{max}(A_i) = G_i^*(A_i, S_j)$  be the highest grade that a student, with a given level of ability  $A_i$ , can attained when the student has the highest educational aspirations j = J (e.g. graduate school), then a student's optimal choice of GPA may reflect potential educational under-achievement  $G_i^*(A_i, S_k) < G_l^{max}(A_i)$ . This potential educational under-achievement students' educational aspirations are

<sup>&</sup>lt;sup>19</sup> The role of teachers in this model would be similar to that of parents in lowering the cost of academic achievement and enhancing its benefits by motivating students to succeed.

<sup>&</sup>lt;sup>20</sup> Bishop (2006) argues that there are different studying and homework cultures by gender, something like "smart boys get high marks without showing effort" or 'it is not cool for boys to work hard to get top grades".

limited by lack of information, borrowing constraints, time impatience, or other intertemporal optimization errors.

In this study of gender gaps in academic achievement, we seek to identify how the distribution of student characteristics maps into the distribution of GPAs differently by gender, taking into account the different educational expectations of the students, controlling for student ability, students' aspirations, labor market work and for the family environment. We are primarily interested in how changes over time in these determinants help account for changes over time in gender differentials in academic achievement. To the extent that women's educational expectations and aspirations have changed over time for exogenous reasons, we will be able to argue that these changes explain the changes in gender achievement gaps. We will estimate the following academic achievement equation,

$$Prob[G_i = c] = h_g^c(S_i, A_i, L_i; X_i, X_i^p), \qquad c = 1, ..., 9,$$
(1)

where  $S_i$  denote the educational expectations and  $A_i$  denote the student' school ability, ideally measured in elementary school. Here, we combine choice of high school (academic, general, vocational), school expectations and aspirations to measure  $S_i$ , and  $A_i$  is proxied using either a subjective measure of school ability or measures of cigarette smoking and alcohol binging, which may relate to time impatience.<sup>22</sup> We include a derived measure of effort, following the tradition in labor economics of deriving non-market time, here study time, as the difference between total time (T) and labor market time ( $L_i$ ):  $E_i = T - L_i$ . Some exogenous characteristics of students  $X_i$ , race and living in a SMSA, and as well as extended set of family characteristics,  $X_i^p$ , thought to be pre-determined variables, are included in the specification.<sup>23</sup>

We estimate a different linear probability model by gender for each level of GPA, which carries some advantages and disadvantages. The advantages of using a linear probability model are that we do not have to rely on the assumptions of normality of residuals and do not have to estimate the variance of residuals, a nuisance parameter, which limits the identification of the effects to relative effects. By comparison with an ordered probit model, this model allows the educational responses to be different by level of GPA. Given that the detailed decomposition of the gender differentials requires linear educational responses, this estimation procedure gives us

<sup>&</sup>lt;sup>22</sup> Educational aspirations and subjective school ability measures are available only for the 12<sup>th</sup> graders.

<sup>&</sup>lt;sup>23</sup> These family environment characteristics include absent father, absent mother, living in the same household as siblings, the number of siblings, whether the mother had a paid job while growing up (not at all, some of the time, most of the time, all the time), the level of education of the father and of the mother.

coefficients that can readily be used. By comparison with a multinomial logit, there is no need to compute the marginal effects at the mean of characteristics, which may not correspond to a representative student for some GPA levels. Among the disadvantages is the fact that the predicted probabilities are not bounded between 0 and 1. In practice, we will find some under-predictions (<0), but the predicted probabilities over GPA levels sum to 1.

### **3.** Data and Descriptive Statistics

The data used are from the "Monitoring the Future" project, which for more than thirtyfive years has been measuring behaviors, attitudes, and values of American secondary school students. These data have been collected by the Institute for Social Research, University of Michigan mainly to monitor substance abuse every year from 1976 onwards for 12<sup>th</sup> graders, and from 1991 onwards for students in Grades 8 and 10. Given higher male drop-out rates (Heckman and Lafontaine, 2008), the sample of boys in 12<sup>th</sup> grade is likely a positively selected sample of boys, which will lead to understate any gender gap favorable to girls by comparison to a wider sample of boys. It is thus useful to compare high school seniors with high school sophomores and 8<sup>th</sup> graders, who remain subject to minimum age school leaving laws. Because of the focus on drug use, those who use illicit drugs as seniors are oversampled, we are thus careful to use the sample weights provided to remove any bias resulting from that oversampling. We use the crosssectional surveys, which comprise 10,000 to 16,000 observations per grade per year for the core questions, resulting in close to half-a-million observations over the entire period. There exists a seldom accessible longitudinal component, which surveys a small subset of the students (Bachman et al., 2002). Many more attitudes and behavioral questions are asked of students answering one of 6 modules.<sup>24</sup> We focus here on the core sample because of the larger sample sizes available, which allow us to perform the breakdown by gender and GPA.

Most variables from the MTF are coded categorically. For variables with non-ordinal categories (e.g. type of high school), we simply use categorical dummies. For ordinal variables that do not have a metric but are available in n categories (e.g. likeness of attending a 4-year

<sup>&</sup>lt;sup>24</sup> The surveys contain a host of non-cognitive variables but they are asked only of a subset of students. Acknowledging that some psychologists (e.g. Duckworth and Seligman, 2006; Hicks, Johnson, Iacono and McGue, 2008) have argued that self-control and self-discipline give girls the "edge", and these issues are at the center of the ADHD debate (Elder and Lubotsky, 2009), we attempt to capture a similar notion with the "alcohol binging" variable, which is present in the core sample.

college, subjective school ability), we generally use the following formula to rescale the index from 0 to 1: Category k = 1 - (n - k + 1)/(n + 1), when k = n is highest category to be recoded into 1. This recoding presumes equal distance between the categories. For the decomposition analysis, these variables are further normalized to have a zero mean over the entire sample of boys and girls.

Our dependent variable is the self-reported school grade which is elicited from the following question: "Core 20: Which of the following best describes your average grade so far in high school? D (69 or below), C- (70-72), C (73-76), C+ (77-79), B- (80-82), B (83-86), B+ (87-89), A- (90-92), A (93-100)."<sup>25</sup> Obviously, grades from administrative data are preferable to self-reported grades because students with different characteristics may misreport their grades differently.<sup>26</sup> Self-reported grades from the MTF, however, are likely quite reliable for several reasons.

First, there are other questions in the survey of seniors asked before this one directed at getting subjective assessments of school ability (Core 16) and intelligence (Core 17), which would allow students, so inclined, to boast about their abilities. We do, in fact, use the subjective school ability response to the question: "Core 16: Compared with others your age throughout the country how do you rate yourself on school ability? Far below average, below average, slightly below average, average, slightly above average, above average, far above average" as control variable in some specifications. Not surprisingly, a smaller proportion of students state that their ability is "far below average" than is found in the question about average GPA. We note that the raw correlation between subjective school ability and self-reported grades is only 58% among seniors.<sup>27</sup> Following the results of Stinebricker and Stinebricker (2008) who find that college students are generally overconfident about their school ability when they enter college, it seems reasonable to assume that students use the subjective measure when forming educational expectations.

Second, the wording of the question on self-reported grades in terms of an upward scale is similar to commonly used questions about self-reported income where individuals are asked to

<sup>&</sup>lt;sup>25</sup> Following standard institutional practice, the self-reported grades in the 9 categories are translated in the numbers: A (93-100) 4.0, A- (90-92) 3.7, B+ (87-89) 3.3, B (83-86) 3.0, B- (80-82) 2.7, C+ (77-79) 2.3, C (73-76) 2, C- (70-72), 1.7, D (69 or below) 1, where 2.3 and 2.7 and so on, are the rounded versions of 2.333 and 2.666.

<sup>&</sup>lt;sup>26</sup> See Balsaa, Giuliano, and French (2011) on grade misreporting by alcohol-binging students.

<sup>&</sup>lt;sup>27</sup> Note that the questions about subjective school ability and educational aspirations were not asked of the  $8^{th}$  and  $10^{th}$  graders. For these students, we only have educational expectations.

declare in which income bracket their income falls and may be less prone to error than simple declarative questions. Finally, as mentioned in the introduction, when we compare the average grades of 12<sup>th</sup> graders from the MTF to those of the NAEP High School Transcript Surveys (HSTS), we find that the gender differences, as well as the grade inflation, do match within standard errors, even though the scales used are somewhat different.<sup>28</sup>

Table 1 reports simple difference-in-difference estimates of the changes over time and by gender in self-reported grades and in expectations about attending graduate or professional school of 12<sup>th</sup> graders. We study three time periods roughly corresponding to the 1980s (1976-1988), the 1990s (1989-1999) and the 2000s (2000-2009). Figure 1, Panel A of Table 1 shows little change over time in the significant female advantage of about 0.2 (on a 4 point scale) in average grades, if anything boys have made small gains (0.010-0.011) in relative grades. Panel B shows that the stability in average grades masks a significant increase in the female advantage in the proportion of students with the highest grades (A (93-100) students), which represent the pool of students who can be confident of being admitted to graduate school if they continue the good work in their undergraduate studies. Panel C shows an even greater and significant increase of the female advantage in expectations of attending graduate school. Indeed from the 1980s to 1990s, the proportion of women expecting to attend graduate school more than doubled from 10% to 21%, while the proportion of men increased by 50%, from 10% to 15%. The fact that the increase in the gender differential in expectations to attend graduate school was more sizeable (5.3%) from 1980s to 1990s, when women' progress in the labor market was sharpest, than from 1990s to the 2000s (2.6%) are in line with our conjecture that gender differences in plans for the future fuel gender differences in high academic achievement. Panel D provides additional descriptive evidence showing that the girls' higher educational expectations are driven by career choices that require graduate studies. In the smaller sample of seniors who answered module 4, the proportion of girls thinking that, at age 30, they will work as a professional with a doctoral degree (or equivalent) has grown by 11.5% from the 1980s to the 1990s, while the equivalent proportion of boys has grown only by 3.3%.<sup>29</sup> Moreover, when asked how likely it is that they

<sup>&</sup>lt;sup>28</sup> The HSTS scale has 5 categories, which include a zero: A(90–100) 4.0, B(80–89) 3.0, C(70–79) 2.0, D(60–69) 1.0, F (less than 60) 0.0.

<sup>&</sup>lt;sup>29</sup> It is interesting to note that they were no significant increases from the 1990s to the 2000s in those proportions. The questionnaire explicitly lists a few of occupations, lawyer, physician, dentist, scientist, college professor, among those requiring a doctoral degree. Note that the following occupations, engineer, architect, and accountant are listed in the professional occupations without doctoral degree, possibly accounting for gender differences. Appendix Table

will get to do that type of work, 77.3% of girls vs. 71.2% of boys state that it is as least "very likely".<sup>30</sup>

A more complete picture of changes in academic achievement is presented in Figure 4 which displays histograms, corresponding to the actual data, overlaid with a kernel density of the self-reported grades of girls and boys in 12<sup>th</sup> grade. Figures 5a and 5b report the same data for 10<sup>th</sup> and 8<sup>th</sup> graders for the 1990s (1991-1999) and 2000s (2000-2009). The figures clearly show a progressive disaffection over the past thirty years with "grading on a curve" with the alternative "competency grading" gaining in importance. In the 1980s, the mode and median of the grades distribution roughly coincided in the B range. By the 2000s, the mode of girls' grade distribution had moved from B to A, while the mode of the boys' grade distribution stayed at B.<sup>31</sup> This is what we call "leaving boys behind" although the proportion of boys in the A range has increased over time; the gender gap in the proportion of students at the very top of the GPA distribution has increased.

Figure 6 displays the female/male difference in the percentage of seniors in each GPA level for each of the three decades of interest. The lines in the figure show the raw differences for the observations for which we have complete information, each corresponding bar previews our decomposition results, which we discuss in section 5 below. In Figure 6, the largest gender gap favorable to girls is in the percentage of students with A's has increased from 3.7% to 6.0 % from the 1980s to the 2000s. The largest gender gap favorable to boys is in the percent of students with C+'s, which has decreased from 4.4% to 3.2%. Similar gender gaps for 10<sup>th</sup> graders and 8<sup>th</sup> graders are displayed in Figures 7a and 7b.

The means of our main core variables for seniors are reported in Table 2 for each of the three decades of interest.<sup>32</sup> The first 9 rows of the table report the exact numbers behind Figure 4a. The next two rows display the average school grade index and the students own evaluation of their school ability. It shows that despite having lower grades, boys rate their own school ability

A-1 reports the complete answers to the question by gender and time period. It shows a sharp decline in clerical office work as an intended occupation for girls, not match by as great of a decline in craftsman and protective services as intended occupations for boys, over the three decades.

<sup>&</sup>lt;sup>30</sup> The various choices are "not very likely, somewhat likely, fairly likely, very likely, certain, already doing it".

<sup>&</sup>lt;sup>31</sup> Similar gender differences can be found in the administrative grades available in the Add Health data.

<sup>&</sup>lt;sup>32</sup> The statistics are computed on observations with no missing variables. This reduces the sample sizes by comparison with Table 1. Descriptive statistics for 10<sup>th</sup> and 8<sup>th</sup> graders are available as supplemental material.

higher than girls, although differences are not always statistically significant.<sup>33</sup> This would lend some support to the motto of effective gender-specific teaching: "build the girls up, break the boys down" (Sax, 2007). Similar male overconfidence has been reported among college students by Stinebricker and Stinebricker (2009). They argued that even when boys are admitted to college, because of their overall lower performance, they are less likely to succeed in spite of their overconfidence.<sup>34</sup>

Demographic characteristics are presented next. They show that the sample is composed of 8% black boys vs.10-11 % black girls; this largely reflects the differential drop-out rates by gender among Blacks. Among 8<sup>th</sup> graders, the sample is composed of 11 % black boys vs. 12 % black girls.<sup>35</sup> The subsequent rows tabulate cigarette smoking and alcohol binging (how frequently did one had more than 5 drinks over the last two weeks) recoded into 4 categories. Although smoking has fluctuated somewhat differently by gender over time, boys are still more likely than girls to report these risky behaviors. In the 2000s, about 25% of boys vs. 23% of girls smoked cigarettes, about 21% boys vs. 13% of girls report two or more instances of alcohol binging over the last two weeks.

As noted above, girls tend to live in families that on the surface might be less likely to foster high academic achievement. For example, although family size has decreased over time, by comparison to boys, girls are raised in larger families. <sup>36</sup> In the 2000s, 37% of girls vs. 33% of boys report living in families with 3 or more children. Similarly, 4% more girls than boys report having an absent father, 3% more girls than boys report that their mother works all the time and about 3% more boys than girls report than their father or mother has completed college. We note that the gender gaps in family characteristics are similar in the sample without Blacks. The proportions of mothers and fathers in the various educational attainment classes provide an additional way to assess the representativeness of the sample and they do in fact correspond to proportions reported elsewhere.

<sup>&</sup>lt;sup>33</sup> Girls in 1976-1988 and boys in 2000-2009 having similar average GPA of 3, but the boys' school ability index of 0.664 is significantly greater than the girls 0.651.

<sup>&</sup>lt;sup>34</sup> Although grades by topic are not reported in the MTF, numerous studies (especially those using the National Education Longitudinal Study) show that boys continue to maintain an advantage in math test scores (but not in math grades), especially at the high end of the distribution. The boys' overconfidence may build on these scores. <sup>35</sup> Descriptive statistics for 10<sup>th</sup> and 8<sup>th</sup> graders are available as supplemental material.

<sup>&</sup>lt;sup>36</sup> See Angrist and Evans (1998).

The means of paid work, hours of work and wages, reported next in Table 2, show that the gender gap in participating in paid work has closed over time, but boys continue to work more hours and get higher pay. About 3% more boys than girls work more that 30 hours a week during school and 7% more boys than girls earn at least \$126 per week for that employment.<sup>37</sup>

The types of high school attended by seniors reported in Table 2 inform us about gender differences in the students' educational plans for the future. The number show that the gap in favor of girls in the proportion of students attending an academic high school has grown; while about 3 % more girls than boys attended an academic high school in the 1980s, that proportion increased to 7% in the 2000s, with 59% of girls vs. 52% of boys attending an academic high school. Conversely the gap in favor of girls in the proportion students attending a general high school has reversed. While 31% of girls vs. 30% of boys attended in general high school in the 1980s, in the 2000s, 30% of girls vs. 33% of boys attended a general high school. Among 8<sup>th</sup> graders, already 4% more girls than boys report that they attend a college preparatory school, although a large proportion of students (43% of both boys and girls) do not know the type of school attended.

Among 12<sup>th</sup> graders, two types of questions regarding post-secondary plans are asked. A first question asks about expectations: "Core 21: How likely (definitively won't, probably won't, probably will, definitively will) is it that you will do each of the following things after high school? a) Attend a technical or vocational school, b) Serve in the armed forces, c) graduate from a two-year college, d) graduate from college (four-year program), e) attend graduate or professional school after college?" A second question asks about aspirations: "Core 22: Suppose you could do just what you'd like and nothing stood in your way. How many of the following things would you WANT to do?" with the five options above being supplemented by none of the above. Among 8<sup>th</sup> and 10<sup>th</sup> graders, only the expectations questions are asked. Among 12<sup>th</sup> graders in particular, the expectations question raises issues of endogeneity with respect to GPA. Some students with low GPA may simply be aiming for a two-year college because of their anticipated career choices; others may have low expectations of graduating from a four-year college because of their GPA. The aspirations question attempts to circumvent that problem with the preamble if "nothing stood in your way". Controlling for subjective school ability (Core 16

<sup>&</sup>lt;sup>37</sup> The categorical data on hours and pay does not allow us to compute a gender pay gap per se.

above) and aspirations (Core 22) may help alleviate these concerns. Among 8<sup>th</sup> graders, the issue of endogeneity of educational expectations is presumably less severe as there is more time to adjust one's level of effort.<sup>38</sup> The same could be said of 12<sup>th</sup> graders about their expectations to attend graduate or professional school.

Interestingly, Table 2 shows that in the 1980s, although seniors of both genders had similar expectations about graduating from college and attending graduate school, girls already have aspirations close to 2% higher about these choices. By the 2000s, the gender differences were sizeable; the expectations index for both college and graduate school was 8% higher for girls than boys<sup>39</sup>. Gender differences in aspirations for college and graduate school are respectively 8% and 11% in favor of girls. Also in line with higher drop out rate among boys, is the fact that 6% of boys vs. 3% of girls have declared no post-secondary aspirations.

### 4. Decomposition Methodology

We follow the literature on gender wage differentials (e.g. Fortin, 2008) in applying an Oaxaca-Blinder type of decomposition to the analysis of gender differences in academic achievement, but we extend the decomposition to the overall distribution of grades and follow the approach of Fortin, Lemieux, and Firpo (2011) to analyze more carefully the impact of gender differences in the educational response functions. We now give a short summary of the formulas behind the modified decomposition reviewed in Fortin, Lemieux, and Firpo (2011). With the aggregate decomposition, we seek to determine what portion of the gender gap in grades is attributable to differences in the educational response to these characteristics. Traditionally, this latter portion has been called the "unexplained" part. Here, we argue that it corresponds to gender differences in the structural function  $h_g^c(S_i, A_i, L_i; X_i, X_i^p)$  of equation (1). With the detailed decomposition, we apportion parts of the aggregate decomposition to particular explanatory factors and responses to determine which of these explanatory factors are relatively more important.

<sup>&</sup>lt;sup>38</sup> Note however that for 8<sup>th</sup> and 10<sup>th</sup> graders, the questions of subjective school ability and post-secondary aspirations were not asked.

<sup>&</sup>lt;sup>39</sup> Comparing seniors in 1972 from the NLS72, in 1980 from the H&B, in 1992 from the NELS88, and in 2004 from the ELS2002, Ingels and Dalton (2008) also find that in 2004, more girls than boys expected to pursue graduate studies, whereas it was the opposite in 1972.

The classic Oaxaca-Blinder methodology is based on the construction of a counterfactual state of world. Assuming that grades (*G*) can be modeled as a linear (in the parameters) function of characteristics (*X*) that is different for girls (F = 1) and boys (F = 0)

 $\mathbb{E}(G|X, F = 1) = \mathbb{E}(X|F = 1)\beta_1$  and  $\mathbb{E}(G|X, F = 0) = \mathbb{E}(X|F = 0)\beta_0$ , under the zero conditional mean assumption,  $\mathbb{E}(\varepsilon|X, F) = 0$ . We can ask "What would boys' grades be if they had the same characteristics as girls?" or "What would girls' grades be if they had the same educational response as boys?",  $\mathbb{E}(G^c) = \mathbb{E}(X|F = 1)\beta_0$ .

Using the counterfactual average grades, we could write the differences between the average grades of girls and boys as

$$\begin{aligned} \Delta_{O}^{\mu} &= \mathbb{E}(G|F=1) - \mathbb{E}(G|F=1) + \mathbb{E}(G^{c}) - \mathbb{E}(G^{c}) \\ &= \mathbb{E}(X|F=1)\beta_{1} - \mathbb{E}(X|F=0)\beta_{0} + \mathbb{E}(X|F=1)\beta_{0} - \mathbb{E}(X|F=1) \\ &= [\mathbb{E}(X|F=1) - \mathbb{E}(X|F=0)]\beta_{0} + \mathbb{E}(X|F=1)(\beta_{1} - \beta_{0}) \\ &= \Delta_{X}^{\mu} + \Delta_{E}^{\mu} \end{aligned}$$

where  $\Delta_X^{\mu}$  would represent the part of the gender differences in average grades due to gender differences in endowments or characteristics, also called composition effect, and where  $\Delta_E^{\mu}$ represent the part of the gender differences due to the fact that boys and girls respond differently to characteristics, this term is called the educational response effect.

If the true conditional expectation is not linear, the OB decomposition is biased (Barsky et al., 2002). As in Fortin, Lemieux, and Firpo (2011), this issue is addressed by using a modified decomposition. We reweight the sample of boys so that the distribution of their characteristics (X) is similar to that of girls, using the following reweighting function

$$\Psi(X) = [(Prob(X | F = 1))/(Prob(X | F = 0))]$$
  
= [(Prob(F = 1 | X))/(Prob(F = 0 | X))] · [Prob(F = 0)/Prob(F = 1)].

The counterfactual coefficients  $\beta_o^1$  will be estimated on the sample of boys reweighted to look like girls { $X_0, \Psi(X_0)$ }, then the difference ( $\beta_1 - \beta_o^1$ ) reflects the true gender gap in educational responses, and the counterfactual means are computed as  $\bar{X}_0^1 = \sum \{i: F = 0\} \Psi(X_i) \cdot X_i$ . The reweighted decomposition uses the terms from the reweighted sample as counterfactuals,

$$\Delta^{\mu}_{O,R} = \mathbb{E}(X|F=1)\beta_1 - \mathbb{E}(X_0|F=1)\beta_0^1 + \mathbb{E}(X_0|F=1)\beta_0^1 - \mathbb{E}(X|F=0)\beta_0$$
$$= \Delta^{\mu}_{E,R} + \Delta^{\mu}_{X,R}$$

to obtain an aggregate decomposition as the sum of an educational response effect,  $\Delta_{E,R}^{\mu}$ , and a composition effect,  $\Delta_{X,R}^{\mu}$ . Inasmuch as grade dummies can be averaged out, this decomposition relies on the additional assumptions of common support and ignorability ( $F \perp \varepsilon | X$ ), that is, conditioning of observables unobservables are assumed to be the same across gender.

Each term of the reweighted decomposition can be further broken down into the "pure" effect and a residual term. The composition effect,  $\Delta^{\mu}_{X,R}$ , is written as the sum of a pure composition effect,  $\Delta^{\mu}_{X,p}$ , and a specification error,  $\Delta^{\mu}_{X,e}$ ,

$$\begin{aligned} \Delta_{X,R}^{\mu} &= \mathbb{E}(X_0|F=1)\beta_o^1 - \mathbb{E}(X|F=0)\beta_0 + \mathbb{E}(X_0|F=1)\beta_0 - \mathbb{E}(X_0|F=1)\beta_0 \\ &= [\mathbb{E}(X_0|F=1) - \mathbb{E}(X|F=0)]\beta_0 + \mathbb{E}(X_0|F=1)(\beta_o^1 - \beta_0) \\ &= \Delta_{X,p}^{\mu} + \Delta_{X,e}^{\mu}. \end{aligned}$$

Similarly, the educational response term,  $\Delta^{\mu}_{E,R}$ , can be written as the sum of a pure response effect  $\Delta^{\mu}_{E,p}$  plus a reweighting error  $\Delta^{\mu}_{E,p}$ ,

$$\begin{split} \Delta^{\mu}_{E,R} &= \mathbb{E}(X|F=1)\beta_{1} - \mathbb{E}(X_{0}|F=1)]\beta^{1}_{o} - \mathbb{E}(X|F=1)\beta^{1}_{o} + \mathbb{E}(X|F=1)\beta^{1}_{o} \\ &= \mathbb{E}(X|F=1)(\beta_{1} - \beta^{1}_{o}) + [\mathbb{E}(X|F=1) - \mathbb{E}(X_{0}|F=1)]\beta^{1}_{o} \\ &= \Delta^{\mu}_{E,p} + \Delta^{\mu}_{E,e} . \end{split}$$

The specification error  $\Delta_{X,e}^{\mu} = \mathbb{E}(X_0|F=1)(\beta_o^1 - \beta_0)$  corresponds to the difference in the composition effects estimated by reweighting and by using simple regressions, where  $\mathbb{E}(X_0|F=1)$  is the mean of the reweighted sample. The reweighting error  $\Delta_{E,e}^{\mu} = [\mathbb{E}(X|F=1) - \mathbb{E}(X_0|F=1)]\beta_o^1$  goes to zero in a large sample.

Because of the linearity of these expressions, the detailed decomposition or the apportionment of the composition and educational response effects to each explanatory variable is straightforward. In practice, this detailed reweighted decomposition can be obtained by running two Oaxaca-Blinder decompositions: OB1) use with sample of girls (F = 1) and the reweighted sample of boys looking like girls to get the pure wage structure effect, OB2) uses with sample of boys (F = 0) and the reweighted sample of boys looking like girls to get the pure wage structure effect.

#### **5. Empirical Results**

Before going on to the decomposition results as such, it is useful to show which of our explanatory variables are more significant and how the educational responses differ by gender. As explained earlier, we estimate models corresponding to equation (1) separately by gender, for each of the nine GPA levels and for each of the three decades to compute the decomposition results. However, to conserve space we report the detailed estimated coefficients for the 2000s only for the two GPA levels where the gender achievement gap were largest, that is for the A and C+ grades.

#### 5.1 Determinants of Top and Below Average GPA

Tables 3a and 3b report the estimated coefficients of the explanatory variables listed in Table 2. In Table 3a, the dependent variable is equal to 100 if the student gets an A, and 0 otherwise, so that the coefficients indicate the added probability of getting an A associated with the explanatory variables. In Table 3b, we estimate the covariates of getting exactly a C+. We present the estimated coefficients from two specifications. Specification 1 includes educational expectations, assuming that students take their abilities and other limitations into account while formulating their expectations. With this specification 2 explicitly controls for subjective school ability and for educational aspirations formed without possible limitations resulting from ability or other constraints. We anticipate that the inclusion of these two set of variables will reduce the explanatory power of educational expectations. To the extent that teachers' assessments and study effort stand between one's subjective school ability assessment and one's actual grade, including subjective student ability as a regressor should be different from the proverbial regressing Y on Y, but this variable should still have considerable explanatory power. As shown in Table 3a and 3b, we find that both of these conjectures are realized.

Yet, even after controlling for student ability and student aspirations, educational expectations remain among the most significant explanatory variables. Getting an A is very significantly positively associated with wanting and expecting to attend graduate school, especially for boys, and negatively associated with expecting to go to a two-year college. Conversely, consistent with the model of section 2, expecting to go to a two-year college is positively associated with the probability of getting a C+, again especially for boys. Similar effects are found for the type of high school attended, much of the impact of the variable is

captured by school ability going from Specification 1 to Specification 2, although it does remain a significant variable with effects in the 2% to 4% range. The types of high school attended, thought to be part of a student's plans for the future, do however show significant differences across genders. Girls are more likely to get A, and less likely to get C+ in academic high school than boys.

Among the other most significant variables, by comparison with non-Blacks, we find that black boys are 6-8% less likely to get an A and 6% more likely to get a C+, while black girls 9%-10% less likely to get an A and 5% more likely to get a C+. Along the lines of Balsa et al. (2011), alcohol binging is associated with a significantly lower the probability of getting an A, about -4%, and a higher probability of getting a C+, about 1-4%. Similarly effects are found for smoking variables, in the -3% to -6% range for getting an A and the +2% for a C+. We view these correlations as symptomatic of time impatience or caring less about the future.

Focusing on family background variables, we find that controlling for school ability (going from Specification 1 to 2) substantially reduces the impact of parental education on students' probabilities of getting an A or a C+, although that association remains significant for girls. To the extent that parental education is capturing the family socio-economic status, these results are consistent with the past research, Cameron and Heckman (2001) and Reynolds and Pemberton (2001), showing that the biggest influence of parental resources on the children's education operates through academic performance. Other important family influences, more impervious to addition of subjective school ability, are the actual presence of parents in the household. The absent father and the mother working have significant effects (about -1 to-4%) on the probability of getting an A, and positive effects on the probability of getting a C+ (about 1%-2%). Interestingly the effect of the absent father is somewhat greater for girls, and that of the mother working is somewhat greater for boys. As Buchmann and DiPrete (2006), we do find that these effects have increased from the 1980s to the 2000s.<sup>40</sup>

In comparison to the effects of expectations, aspirations, subjective school ability and family background, the effects of the variables related to working during school are generally less significant and show some of the non-linear pattern found in the literature. However, there are significant gender differences in the coefficients of the work variables. In Table 3a, the coefficient of "work during school" is negative in the range of -2% to -3% for boys and positive,

<sup>&</sup>lt;sup>40</sup> Regression coefficients not shown, but this result will be clear in the Figures below.

but not significant for girls. Conversely, in Table 3b, it is positive in the 1% to 2% range for girls and not significant for boys. Thus changes over time in the gender differentials associated with working during school are likely to account for the change in the gender achievement gaps in the educational response portion of the decomposition.

#### **5.2 Decomposition results**

To succinctly summarize the decomposition results, we will mostly present them in the form of graphs in order to display the entire GPA distributions. Figure 6 displays the aggregate decomposition for  $12^{\text{th}}$  graders. Based on Specification 1, it shows how the female/male differences in percentage for each GPA levels,  $\Delta_{O,R}^c$ , c = 1, ... 9, can be decomposed into composition effects,  $\Delta_{X,R}^c$ , the portion "explained" by gender differences in characteristics, and educational response effects,  $\Delta_{E,R}^c$ , the portion "unexplained" but attributed to the fact that the relationship between characteristics and GPA levels differs by gender. For each of the three time periods, positive bars indicating the excess percentage of girls in a GPA level (negative bars indicating the excess percentage of boys) are divided into two; the bottom darker (blue) portion corresponds to the composition effects and the upper lighter (beige) portion corresponds to the educational response effects. We see that the portions attributable to composition effects have generally increased over the three time periods, especially at the top of the grade distribution. Averaging over all GPA levels, the "explained" part grew from a mere 10% of the total gender differential in the 1980s to 32% in the 1990s and to 37% in the 2000s.

Figures 7a and 7b show the results of the aggregate decomposition for 10<sup>th</sup> and 8<sup>th</sup> graders, respectively. The specification of the educational responses functions for these younger students is similar to the Specification 1 used for seniors, given that the questions about subjective school ability or educational aspirations were not asked of these younger students.<sup>41</sup> The figures show that a very sizeable share of the gender differences in the percentage of students attaining each GPA level is accounted by gender differences in the explanatory variables, the composition effects. Among 10<sup>th</sup> graders, averaging over all GPA levels, the

<sup>&</sup>lt;sup>41</sup> The list of variables available for 10<sup>th</sup> graders and 8<sup>th</sup> graders is the following: dummies for race (white/non-white), SMSA, smoked cigarettes per day (4), alcohol binging last two weeks (4), sibling not same household, father not same household, mother working (3), father's education (7), mother's education (7), worked during school, average hours of work (6), average earnings (7), type of high school (4), educational expectations (army, vocational, go to college, complete 4 year college). So the main differences with Specification 1 for seniors are the absence of the number of siblings and of the expectations of going to graduate school.

"explained" part accounts for almost half of the total gender differential: more precisely 49.5% in the 1990s and 46.7% in the 2000s. Among 8<sup>th</sup> graders, the "explained" part also accounts for a large portion of the total gender differential, more precisely 33.9% in the 1990s and 51.1% in the 2000s.

For seniors, we also present a subset of results in a more classic tabular form, which includes standard errors.<sup>42</sup> As in Table 3a and 3b, Table 4a and 4b present the decomposition results for the two GPA levels where the gender differentials are the largest (A and C+) and for Specifications 1 and 2, but this time we present results for all three time periods. Table 4a presents the detailed decomposition of the composition effects and Table 4b the detailed decomposition of the educational response effects. Note that the specification errors  $\Delta_{X,e}^c$  are reported in Table 4a and the reweighting errors  $\Delta_{E,e}^c$  are reported in Table 4b. The specification and reweighting errors are generally found to be at least an order of magnitude smaller than the main effects  $\Delta_{X,p}^c$  and  $\Delta_{E,p}^c$ .

First, going through column 1 of Table 4a shows the increasing female advantage in top grades, as the female/male difference in the percentage of students getting A's increases from 3.747 in the 1980s, to 4.711 in the 1990s to 6.063 in the 2000s.<sup>43</sup> At the same time, the male percentage advantage in column 3 in the C+ grade decreases from 4.429 in the 1980s, to 3.898 in the 1990s to 3.152 in the 2000s. Second, the decomposition shows that overall, until the 2000s, gender differences in the explanatory variables did little to account for gender differences in academic achievement. There is one important exception: From the 1980s through the 2000s, gender differences in smoking and alcohol binging consistently account for 0.560 to 0.602 in Specification 1, and robustly from 0.427 to 0.489 in Specification 2 (that is controlling for subjective school ability) of the gender gap in getting A's. Smoking and alcohol binging also account for gender differences in getting C+, reliably but declining over time from -0.402 to -0.198 in Specification 1 (from 0.300% to 0.163% in Specification 2).

Third, because of the boys' overconfidence about their school ability, the effect of subjective school ability goes in the wrong direction: the female/male difference in own school ability is negative and its coefficients are positive for top grades and negative for mediocre

<sup>&</sup>lt;sup>42</sup> Because of the large number of observations statistical significance is never an issue in the decomposition results, the issue of economic significance becomes more important.

<sup>&</sup>lt;sup>43</sup> These numbers are a bit different from the ones reported in Table 1, Panel B, row 3 (3.2, 4.4, and 5.4) because for the analysis, we restrict the sample to those observations for which we have complete data.

grades. In the aggregate decomposition, this reduces the part of the gender differentials, negative for top grades and positive for mediocre grades, accounted by the explanatory variables, that is, going from Specification 1 to Specification 2 the "Total Explained" is smaller.<sup>44</sup> In the 2000s, the Total Explained corresponds to more 40% of the gender achievement gap in Specification 1, but only 17% in Specification 2. We note that race, SMSA, and family background variables are other sets of "contrarian" or "swimming upstream" variables, whose effects increase over time: these variables work to the advantage of boys (because there are more black girls, more girls with absent father, etc.,) and reduce the percentage of girls with top grades and of boys with mediocre grades. That is, if girls were as confident as boys about their school ability, if they lived in similar families, if there were as few Black girls living in SMSA as boys, the girls' grades would be even higher. For example, in the 2000s, there would be from 0.641 to 0.928 percent more girls than boys earning A's.

Fourth, from the 1980s to the 2000s, the portion of gender differences accounted for by educational expectations increased. Indeed, for the A grades, gender differences in expectations accounted for 1.132 in 1990s and 2.029 in the 2000s, or from 75% to 85% of the Total Explained in Specification 1, and more than 100% of the Total Explained in Specification 2. As the inclusion of subjective school ability removed some of the potential omitted variable bias from the educational expectations coefficients, it reduced the magnitude of the gender differentials accounted by educational expectations, but it did not reduce (actually increases in some cases) the relative portion that is accounted by educational expectations. For the C+ grades, only starting the 1990s do gender educational expectations get any bite at accounting for gender differences in academic performance; but in the 2000s, they essentially accounted for more than entire explained portion of the gender differential. Figure 6b displays the aggregate decomposition for Specification 2 and thus provides a lower bound for the effects of expectations. Nevertheless, overall the results from both Specification 1 and Specification 2 convey the same message as the one suggested by Table 1: Even after controlling for a host of other factors, changes in gender differences in educational expectations largely account for the more salient changes in gender differentials in academic achievement.

<sup>&</sup>lt;sup>44</sup> This effect is similar to the gender differences in educational attainment on the gender pay gap. In recent years, gender differences in education reduce the explained part of the gender pay gap.

It is also interesting to consider the contribution of changes in gender differences in educational responses presented in Table 4b, noting that the interpretation of these differences crucially depends on the omitted category in each case. The most persistent factor from the 1980s to 2000s is the type of high school attended, where the omitted category is "other (not specified) high school". As we saw in previous tables, not only are girls increasingly attending college preparatory high school, but they are benefiting more (in terms of grades) from it then boys. This differential educational response adds to the total effect of "plans for the future" factors in accounting for gender difference in academic achievement.

One factor that is increasingly important over the decades in accounting for gender differences in educational responses is "work during school", where the omitted categories are not working, zero hours of work and zero wages. It shows that over the three decades, working during school seems to have increasingly acted as complement rather than a distraction for high achieving girls. Consistent with a non-linear effect of work, the educational responses linked to working during school contributed to the positive gender gap in favor of boys in getting a C+, although the magnitude of this effect is smaller than the previous effect.

The effects of gender differences in educational responses associated with family background is more difficult to interpret because departures from the omitted category (families with father present, mother present, one sibling, mother not working, both parents with high school education) are a more complex affair and the results are sensitive to which number of siblings is the omitted category (especially in the 1990s).<sup>45</sup> Nevertheless, they indicate that family background generally bolsters the response of high achieving girls by comparison with boys.

On the basis of these findings, we display the results of the detailed decomposition from Specification 1 for a reduced set of four factors: individual characteristics (race, SMSA, smoking and binging), family background, working during school, plans for the future (includes type of high school and educational expectations) in Figures 8 for 12<sup>th</sup> graders and Figure 9a and 9b for 10<sup>th</sup> graders and 8<sup>th</sup> graders. As in previous figures, the lines trace the magnitude of the gender gap in academic achievement to be explained across the GPA distribution, and the bars for each GPA levels are divided into two, the darker (blue) one capturing the composition effects and the

<sup>&</sup>lt;sup>45</sup> Such sensitivity is not surprising given that even using an instrumental variable strategy that exploit exogenous variation in family size, Conley and Glauber (2006) find a strong effect of sibship size on second-born boys' grade retention, but no effect on first-born boys.

lighter (beige) the educational response effects. In some instances, either effect can be negative, as the family effects above. The distance between the height of the bars and the symbol on the line corresponds to the portion of the gender differential accounted for by the other factors presented in the different panels.

The overall message emerging from Figure 8 is the same as the one we took away from Tables 4a and 4b. The effects of "Plans for the future" are by far the most important explanatory factors contributing to both the composition and educational response effects, generally with the right signs, except for the very low GPA levels. To the extent that they account for a sizeable share of the gender achievement gap, they provide some supporting evidence for the model presented in section 3 and for the idea that given their different plans for future, youths of both genders are optimizing their academic effort to fit these plans. More girls than boys are aiming for professions that require a graduate degree, more girls are getting A's. More boys than girls are aiming for craftsman jobs and protective service occupations, more boys are getting C+'s.<sup>46</sup> Interestingly that message is even stronger among 10<sup>th</sup> graders and 8<sup>th</sup> graders. For these younger students, plans after high school are arguably further in the future and thus less likely endogenous (in the sense of resulting from cognitive dissonance issues). Youths with lower GPA are less likely to say that they will not go to college because of their lower GPA, given that many believe that there is still time for improvement. In Figures 9a and 9b, the composition effects associated with plans for the future are generally accounting for more than 50% of the gender differentials, both at the high and low end of the GPA distribution.

Figure 8 shows that the other factors of interest do contribute to a smaller extent to the gender differentials in achievement among seniors. The larger amounts of smoking and alcohol binging among boys than girls contribute to the explained part of the gender differential in getting high grades, but the girls' response is far more negative than the boys'. Family background work to the advantage of boys, but girls' response to unfavorable family backgrounds helps them achieve higher grades. The boys' response to working during school appears to prevent them from moving from the B range to the A range, while it helps girls getting straight A's. Overall, however the contribution of these other factors is very small.

The same message emerges from Figures 9a and 9b; aside from plans for the future, the other gender differences in characteristics explain only minute parts of the gender gap in

<sup>&</sup>lt;sup>46</sup> See Table A1.

academic achievement. The effects of smoking and alcohol binging are also not surprisingly smaller among the young students, as are the effects of work during school. The phenomena of high achieving girls "swimming upstream" an unfavorable family background environment found among seniors, is also present among 10<sup>th</sup> and 8<sup>th</sup> graders, but the magnitude of the effect is small. Similarly, the effect of working during school among the younger students is negligible.

In summary, the aggregate decomposition results show a marked improvement, over three decades, in the model's ability to account for gender differences in academic achievement. Surprisingly, this finding applies equally well to the three grade levels (12<sup>th</sup>, 10<sup>th</sup>, and 8<sup>th</sup> graders) considered. The detailed decomposition results point out to changes in educational expectations as the main factor accounting for the added explanatory power of the model. Given the exogenous changes in the labor market opportunities of women over the last three decades, this is not completely unanticipated. The startling aspect comes from the fact that the explanatory power of the educational expectations is greater for 8<sup>th</sup> graders than for 12<sup>th</sup> graders. To the extent that educational expectations of 8<sup>th</sup> graders are less likely endogenous with respect to GPA levels than those of seniors, this is welcome news for the validation of the model. Certainly, among 12<sup>th</sup> graders, the specification 2, which controls for subjective school ability and educational aspirations, grants less lower explanatory power to the model and to educational expectations. But it still supports the idea that educational expectations, and changes in these expectations over the last three decades, is the most likely explanation for why girls are leaving boys behind in terms of earning top grades in high school.

#### **6.** Conclusion

Using a long-lived series of detailed cross-sectional surveys of high school students, this paper set out to identify which factors among a set of plausible culprits,—family environment, labor market work during school, and plans for the future—, would be more important in accounting for the changes over the past three decades in the gender achievement gap, especially at the top of the GPA distribution. The focus on the gender gap in top grades follows from the findings of previous studies (Jacob, 2002; Frenette and Zeman, 2007; Cho, 2007; Conger and Long, 2010) showing that the lower college admission rates of men can in large part be accounted for by their lower high school performance. We note that better high school performance explains "how" more girls are admitted to college but not "why". Frenette and

26

Zeman (2007) also find that parental expectations of the highest level of educational outcome of the child account for a notable share of gender differences in university attendance. Further the role of parental expectations could explain why first generation immigrant boys do not suffer from the boys' underachievement problem.<sup>47</sup>

Over the last three decades there has been sustained effort (Manski and Wise, 1983; Manski, 1993; Dominitz, Manski, and Fischhoff, 2001; Manski, 2004) to understand the formation of students' expectations and to ascertain the importance of these expectations in their decision to enroll in college. Recent studies (Stange, 2008; Jacob and Wilder, 2010; Zafar, 2011) set in longitudinal settings have focused on learning and beliefs updating, trying to address on the first part of the puzzle. Not having the luxury of longitudinal data, we bypass the first issue of expectations formation assuming that the majority of students, by the time they leave elementary school, have fashioned some expectations about college-going. Indeed, decisions to attend an academic (college preparatory) high school, to move to a neighborhood with a better high school, and to apply to enroll in a magnet school are made early in pupils' life. Then we rely on changes over time in the educational expectations of students to evaluate the importance of these expectations for the high academic achievement that opens the door to college-going. These changes are different by gender for reasons arguably exogenous to the early education process; rather they are rooted in changes in the labor market opportunities for women.<sup>48</sup>

These considerations lead us to propose a "signaling" type model of academic achievement where educational expectations, formed in elementary school and likely influenced by parental desires, play a prominent role in determining, given a level of aptitude, in an individual's choice of optimal GPA. Thus in this model, high school students can optimality choose a GPA level lower than their maximum attainable given their aptitude level, if their postsecondary education plans can be enacted with the lower GPA level. This model contrasts with both the early childhood development (ECD) branch and education production function (EPF) branch of the literature on cognitive achievement in children (Todd and Wolpin, 2003), where the goal is to maximize achievement under some cost constraints. Here, our model is set in the

<sup>&</sup>lt;sup>47</sup> Wilson, Burgess and Briggs (2011) have also suggested aspiration-based explanations to account for ethnic differences in academic performance. See also Zafar (2011).

<sup>&</sup>lt;sup>48</sup> A caveat here is the disproportionate increase in ADHD among boys (Elder and Lubotsky, 2009; Chen, Fortin, Oreopoulos and Phipps, 2011). But the recent phenomena is likely affecting cohorts younger than the ones studied.

high school years after the revelation of basic school ability (literacy and numeracy) has taken place.

Our findings show that the predominance of girls at the top of the GPA distribution is rooted in their higher educational expectations, themselves linked to career plans that include a graduate degree (such a law or medical degree). More precisely, in the 2000s, "Plans for the Future" is most important set of explanatory factors accounting for the girls' higher share of A's at the three grade levels (12<sup>th</sup>, 10<sup>th</sup>, and 8<sup>th</sup> graders). This set of factors is important enough to account for all of the increase of 2.3%, from 1980s to 2000s, in the gender difference in the percentage of students earnings A's. A more minor, but still interesting finding is that high achieving girls are "swimming upstream", being more likely from a disadvantageous family environment.

By comparison with girls, more boys think that they are likely to enter military service or to attend a vocational school. Because the career plans of boys include more predominantly male occupations (craftsmen, protective service and military service occupations, engineers and architects) that do not require advanced degrees, their lower share of high grades is consistent with the "signaling type" model that we propose. However, these findings should not alleviate all concerns about the possibility of academic underachievement by boys. To the extent that boys are overconfident about their school ability, they may fail to enact their expected education plans.

#### **References:**

- American Association of University Women (AAUW) (1995) *How Schools Shortchange Girls: The AAUW Report-* A Study of Major Findings on Girls and Education, Marlowe & Company.
- Bachman, Jerald G., Patrick M. O'Malley, John E. Schulenberg, Lloyd D. Johnston, Alison L. Bryant, and Alicia C. Merline (2002) *The Decline of Substance Use in Young Adulthood- Changes in Social Activities, Roles and Beliefs*, Mahwah, N.J.: Lawrence Erlbaum Assoc.
- Balsaa, Ana I., Laura M. Giuliano, and Michael T. French (2011) "The Effects of Alcohol Use on Academic Achievement in High School," *Economics of Education Review*, 30: 1–15.
- Barsky, R., Bound, John, Charles, K., Lupton, J. (2002) "Accounting for the Black-White Wealth Gap: a Nonparametric Approach," *Journal of the American Statistical Association*, 97:663–673.
- Becker, Gary and N. Tomes (1979) "An Equilibrium Theory of the Distribution of Income and Intergenerational Mobility," *Journal of Political Economy*, 87:1153-89.

- Bedard, Kelly and Insook Cho (2010) "Early Gender Test Score Gaps across OECD countries," *Economics of Education Review*, 29(3): 348-363.
- Bell, Lisa, Gary Burtless, Janet Gornick and Timothy M. Smeeding (2007) "Failure to Launch: Cross-National Trends in the Transition to Economic Independence," chap. 2 in *The Price of Independence: the Economics of Early Adulthood*, edited by Sheldon Danziger and Cecilia Elena Rouse. New York: Russell Sage Foundation.
- Ben-Porath, Yoram and Finis Welch (1976) "Do Sex Really Preferences Matter? "*Quarterly Journal of Economics*, 285-307.
- Bishop, John H. (2006) "Drinking from the Fountain of Knowledge: Student Incentive to Study and Learn
  Externalities, Information Problems and Peer Pressure," in *Handbook of the Economics of Education*, Elsevier North Holland, Vol. 2: 909-944.
- Blank, Rebecca (2011) "Women in America: Indicators of Social and Economic Well-Being." Report to the White House Council on Women and Girls, Economic and Statistics Administration: U.S. Department of Commerce.
- Buchmann, Claudia and Thomas A. DiPrete (2006). "The Growing Female Advantage in College Completion: The Role of Family Background and Academic Achievement," *American Sociological Review*, 71: 515-541.
- Buchman, Claudia, Thomas A. DiPrete, and Anne McDaniel (2008). "Gender Inequalities in Education," *Annual Review of Sociology*, 34:319-37.
- Cameron, Stephen V. and James J. Heckman (2001) "The Dynamics of Educational Attainment For Black Hispanic, and White Males," *Journal of Political Economy*, 109(3): 455-499.
- Charles, K., and Luoh, M.-C. (2003), "Gender Differences in Completed Schooling", *Review of Economics and Statistics*, 85(3), 559–77.
- Chen, Kelly, Nicole Fortin, Philip Oreopoulos and Shelley Phipps (2011) "School Start Age and Hyperactivity in Canadian Children," mimeo, Dalhousie University.
- Cho, Donghun, (2007) "The Role of High School Performance in Explaining Women's Rising College Enrollment," *Economics of Education Review*, 26: 450–462
- Conger, D. and M.C. Long (2010) "Why are Men Falling Behind? Explanations for the Gender Gap in College," *Annals of the American Academy of Political and Social Science*, 627(1): 184-214.
- Conley, Dalton and Rebecca Glauber (2006) "Parental Educational Investment and Children's Academic Risk: Estimates of the Impact of Sibship Size and Birth Order from Exogenous Variation in Fertility," *Journal of Human Resources*, 41(4): 722-737.
- Dahl, Gordon B. and Enrico Moretti (2008) "The Demand for Sons." *Review of Economic Studies* 75 (4): 1085-1120.

- Dee, Thomas (2005) "A Teacher Like Me: Does Race, Ethnicity, or Gender Matter?" *American Economic Review*, 95:158-166.
- DiPrete, Thomas A. and Claudia Buchmann, (2006) "Gender-Specific Trends in the Value of Education and the Emerging Gender Gap in College Completion," *Demography*, 43: 1-24.
- Dominitz, J., C. Manski, and B. Fischhoff (2001): "Who Are Youth At-Risk?: Expectations Evidence in the NLSY-97," in *Social Awakenings:A dolescents' Behaviors as Adulthood Approaches*, ed. by R. Michael. New York: Russell Sage Foundation, 230-257.
- Duckworth, A. L., and M. E. P. Seligman (2006) "Self-Discipline Gives Girls the Edge: Gender in Self-Discipline, Grades, and Achievement Test Scores" *Journal of Educational Psychology* 98(1), 198-208.
- Elder, Todd and Darren H. Lubotsky (2009) "Kindergarten Entrance Age and Children's Achievement: Impacts of State Policies, Family Background, and Peers," *Journal of Human Resources*, 44(3): 641-683.
- Fortin, N.M. (2008) "The Gender Wage Gap among Young Adults in the United States: The Importance of Money vs. People," *Journal of Human Resources*, 43: 886-920.
- Fortin, N.M., T. Lemieux and S. Firpo (2011) "Decomposition Methods," in O. Ashenfelter and D. Card (eds.) Handbook of Labor Economics, Vol. 4A, Amsterdam: North-Holland , 2011, 1-102.
- Frenette, Marc and Klarka Zeman (2007) "Why are Most University Students Women? Evidence based on Academic Performance, Study Habits and Parental Influences," Analytical Studies Branch Research Paper Series No 303. Ottawa: Statistics Canada.
- Goldin, C. and L.F. Katz (2002) "The Power of the Pill: Oral Contraceptives and Women's Career and Marriage Decisions," *Journal of Political Economy*, 110: 730-770.
- Goldin, Claudia, Lawrence F. Katz, and Ilyana Kuziemko (2006) "The Homecoming of American College Women: The Reversal of the College Gender Gap," *Journal of Economic Perspectives*, 20: 133–156.
- Goldin, C. and C. Rouse (2000) "Orchestrating Impartiality: The Impact of 'Blind' Auditions on Female Musicians," *American Economic Review*, 90 (3): 715-41.
- Greene, J. P. & Winters, M. A. (2006). Leaving boys behind: Public high school Graduation Rates. Manhattan Institute for Policy Research, No. 48.
- Gurian, Michael and Stevens, Kelly (2007). *The Minds of Boys: Saving our sons from falling behind in school and life.* San Francisco: Jossey Bass.
- Guiso, Luigi, Ferdinando Monte, Paola Sapienza, and Luigi Zingales, (2008) "Culture, Gender, and Math," *Science*, 320(5880): 1164-1165

- Heckman, James J. and Paul A. LaFontaine (2007) "The American high school graduation rate: Trends and levels", IZA DP No. 3216.
- Herbert, Jennifer and Deborah Stipek (2005) "The Emergence of Gender Differences in Children's Perceptions of their Academic Competence," *Applied Developmental Psychology*, 26: 276–295.
- Hicks, Brian M., Wendy Jonhson, Paul Iacono, and Matt McIgue (2008) "Moderating Effects of Personality on the Genetic and Environmental Influences of School Grades Helps to Explain Sex Differences in Scholastic Achievement," *European Journal of Personality*, 22(3): 247-268.
- Ingels, S.J., and Dalton, B.W. (2008). Trends Among High School Seniors, 1972–2004 (NCES 2008-320). National Center for Education Statistics, Institute for Education Sciences, U.S. Department of Education. Washington, DC.
- Jacob, Brian A. (2002) "Where the Boys Aren't: Non-Cognitive Skills, Returns to School and the Gender Gap in Higher Education," *Economics of Education Review*, 21: 589–598
- Jacob, Brian A. and Tamara Wilder (2010) "Educational Expectations and Attainment," NBER Working Paper No. 15683, Cambridge, MA: National Bureau of Economic Research.
- Kramarz, Francis, Machin, Stephen, and Amine Ouazad (2008) "What Makes a Test Score? The Respective Contributions of Pupils, Schools, and Peers in Achievement in English Primary Education," IZA Discussion Paper No. 3866.
- Lai, Fang (2010) "Are Boys Left Behind? The Evolution of the Gender Achievement Gap in Beijing's Middle Schools," *Economics of Education Review*, 29(3): 383-399.
- Lee, Chanyoung and Peter F. Orazem (2010) "High School Employment, School Performance, and College Entry," *Economics of Education Review*, 29 (2010):29–39.
- LoGerfo, L, A Nichols, D Chaplin (2006) *Gender Gaps in Math and Reading Gains during Elementary* and High School by Race and Ethnicity, The Urban Institute.
- Lundberg, Shelly and Elaina Rose (2002) "The Effects of Sons and Daughters on Men's Labor Supply and Wages," *Review of Economics and Statistics*, 84(2): 251-268.
- Machin, Steve, and Sarah McNally (2005). Gender and Student Achievement in English Schools. *Oxford Review of Economic Policy*, 21(3), 357–372.
- Manski, Charles F., and David Wise (1983) *College Choice in America*. Cambridge: Harvard University Press.
- Manski, Charles F. (1993): "Adolescent Econometricians: How Do Youth Infer the Returns to Schooling?" in *Studies of Supply and Demand in Higher Education*, ed. by C. Clotfelter and M. Rothschild. Chicago: University of Chicago Press, 43-57.

Manski, Charles F. (2004) "Measuring Expectations," Econometrica, 72(5): 1329-1376.

- National Center for Education Statistics (NCES). (2004). The High School Transcript Study: A Decade of Change in Curricula and Achievement, 1990–2000. NCES 2004–455, by Robert Perkins, Brian Kleiner, Stephen Roey, and Janis Brown. Project Officer: Janis Brown. Washington, DC: 2004.
- Niederle, Muriel and Lise Vesterlund (2010) "Explaining the Gender Gap in Math Test Scores: The Role of Competition," *Journal of Economic Perspectives*, 24: 129-144.
- Organization for Economic Cooperation and Development (OECD) (2008). "The Reversal of Gender Inequalities in Higher Education: An On-going Trend" by Stéphan Vincent-Lancrin, Chap. 10 in *Higher Education to 2030, Vol.1: Demography.*
- Reynolds, J.R., and Pemberton, J. (2001). "Rising College Expectations Among Youth in the United States: A Comparison of the 1979 and 1997 NLSY". *Journal of Human Resources*, 36(4):703– 726.
- Rothstein, Donna S. (2007) "High School Employment and Youths' Academic Achievement," *Journal of Human Resources*, 42: 194–213.
- Sadker, M. and Sadker, D. (1994). *Failing at Fairness: How our Schools Cheat Girls*. New York: Touchstone.
- Sax, Leonard (2007) Boys Adrift: The Five Factors Driving the Growing Epidemic of Unmotivated Boys and Underachieving Young Men, Basic Books.
- Spence, A. Michael (1973) "Job Market Signaling," Quarterly Journal of Economics, 87 (3): 355-374.
- Stange, Kevin M. (2008) "An Empirical Investigation of the Option Value of College Enrollment," forthcoming in the *American Economic Journal*.
- Stinebrickner, Todd and Ralph Stinebrickner (2009) "Learning about Academic Ability and The College Drop-Out Decision", NBER Working Paper 14810, National Bureau of Economic Research: Cambridge, MA .
- Tyler, J. H. (2003) "Using State Child Labor Laws to Identify the Effect of School-to-Work on High School Achievement," *Journal of Labor Economics*, 21: 381–408.
- Tyre, Peg (2008) "The Trouble with Boys". New York: Crown Publishers.
- Todd, Petra E. and Kenneth I. Wolpin (2003) "On the Specification and Estimation of the Production Function for Cognitive Achievement", *Economic Journal*, 113(485): F3–F33.
- Wilson, Deborah, Simon Burgess and Adam (2011) "The Dynamics of School Attainment of England's Ethnic Minorities", *Journal of Population Economics*. 24(2): 681-700.
- Zafar, Basit (2011) "How Do College Students Form Expectations?" *Journal of Labor Economics*, 29(2): 301-348.

Time period	1976-1988	1988-1999	Change over time (2)-(1)	2000-2009	Change over time (4)-(2)	U U	
	(1)	(2)	(3)	(4)	(5)	(6)	
A: Average grad	les						
Girls	3.004	3.106	0.102	3.218	0.112	0.214	
	(0.002)	(0.002)	(0.003)	(0.003)	(0.003)	(0.003)	
Boys	2.804	2.907	0.103	3.030	0.123	0.225	
	(0.002)	(0.002)	(0.003)	(0.003)	(0.004)	(0.003)	
Difference	0.200	0.199	-0.001	0.189	-0.010	-0.011	
	(0.003)	(0.003)	(0.004)	(0.004)	(0.005)	(0.005)	
B: Proportion wi	ith A grade						
Girls	0.100	0.143	0.043	0.192	0.048	0.091	
	(0.001)	(0.004)	(0.001)	(0.004)	(0.002)	(0.002)	
Boys	0.069	0.099	0.030	0.137	0.038	0.068	
·	(0.001)	(0.001)	(0.001)	(0.001)	(0.002)	(0.001)	
Difference	0.032	0.044	0.012	0.054	0.011	0.023	
	(0.002)	(0.002)	(0.002)	(0.003)	(0.003)	(0.002)	
C: Proportion de	finitely will a	ttend graduat	e or profession	al school <sup>a</sup>			
Girls	0.101	0.205	0.104	0.249	0.044	0.147	
	(0.001)	(0.001)	(0.002)	(0.002)	(0.002)	(0.002)	
Boys	0.099	0.150	0.051	0.168	0.018	0.069	
·	(0.001)	(0.001)	(0.002)	(0.001)	(0.002)	(0.002)	
Difference	0.003	0.055	0.053	0.081	0.026	0.078	
	(0.002)	(0.002)	(0.002)	(0.003)	(0.003)	(0.002)	
Number of							
observations	207,152	160,403		118,173			
D: Proportion th	ink that will w	ork as a prof	essional with c	loctoral degre	ee (or equiv) v	when 30 <sup>b</sup>	
Girls	0.143	0.258	0.115	0.266	0.008	0.123	
	(0.003)	(0.004)	(0.004)	(0.004)	(0.006)	(0.005)	
Boys	0.136	0.169	0.033	0.165	-0.005	0.028	
	(0.003)	(0.004)	(0.004)	(0.004)	(0.005)	(0.005)	
Difference	0.006	0.089	0.082	0.101	0.012	0.095	
	(0.004)	(0.005)	(0.006)	(0.006)	(0.008)	(0.007)	
Number of							
observations	36,699	23,592		19,168			

Table 1. Difference-and-Differences Estimates in Academic Performance andPlans for the Future - 12th graders

Note: Self-reported grades in 9 categories (D, C-,C,C+,B-,B,B+,A-,A) are translated into the numbers 1, 1.7,2,2.3,2.7,3,3.3,3.7 and 4 following standard institutional practice.

<sup>a</sup> The numbers for other post-secondary choices are reported in Table 2.

<sup>b</sup> The numbers for other intented occupations are reported in Table A-1.

Core Variables	1976-1988			1989-1999			2000-2009		
	Boys	Girls		Boys	Girls		Boys	Girls	
Self-reported Grades:									
D (69 or below): 1	0.014	0.006	*	0.015	0.006	*	0.014	0.005	*
C- (70-72): 1.7	0.045	0.022	*	0.036	0.018	*	0.031	0.016	*
C (73-76): 2	0.104	0.065	*	0.086	0.053	*	0.065	0.039	*
C+ (77-79): 2.3	0.146	0.101	*	0.126	0.087	*	0.099	0.068	*
B- (80-82): 2.7	0.165	0.133	*	0.149	0.120	*	0.129	0.103	*
B (83-86): 3	0.203	0.218	*	0.202	0.200		0.187	0.169	
B+ (87-89): 3.3	0.154	0.201	*	0.160	0.197	*	0.175	0.190	*
A- (90-92): 3.7	0.093	0.140	*	0.119	0.165	*	0.154	0.203	*
A (93-100): 4	0.076	0.113	*	0.108	0.155	*	0.147	0.207	*
School Grade Index	2.845	3.055	*	2.944	3.144	*	3.067	3.258	*
Subjective School Ability Index									
(scaled between 0 and 1)	0.652	0.651		0.658	0.654		0.664	0.658	*
Race: Black	0.083	0.097	*	0.085	0.105	*	0.084	0.107	*
Live in MSA	0.683	0.683		0.731	0.738	*	0.755	0.759	
Smoked cigarettes per day: None	0.715	0.673	*	0.678	0.694	*	0.749	0.774	*
Less than one-half pack	0.212	0.260	*	0.258	0.260		0.217	0.201	*
One to 1½ pack	0.070	0.064	*	0.060	0.044	*	0.030	0.023	*
Two packs or more	0.003	0.002	*	0.005	0.002	*	0.004	0.002	*
Alcohol binging last 2 weeks: None	0.534	0.713	*	0.635	0.775	*	0.686	0.780	*
Once	0.129	0.111	*	0.109	0.092	*	0.100	0.094	*
Two to nine times	0.307	0.167	*	0.231	0.127	*	0.197	0.121	*
Ten or more times	0.030	0.008	*	0.025	0.006	*	0.017	0.006	*
Siblings not same household	0.243	0.235	*	0.326	0.312	*	0.329	0.311	*
Siblings: None	0.046	0.042		0.056	0.051	*	0.060	0.052	
One	0.268	0.256	*	0.326	0.311	*	0.323	0.300	*
Two	0.262	0.254		0.271	0.270		0.282	0.272	*
Three or more	0.424	0.449	*	0.344	0.365	*	0.332	0.374	*
Father not same household	0.169	0.185	*	0.201	0.228	*	0.207	0.244	*
Mother not same household	0.075	0.066	*	0.098	0.084	*	0.096	0.089	*
Mom working: No	0.312	0.299	*	0.198	0.184	*	0.146	0.140	*
Some of the time	0.312	0.302	*	0.254	0.242	*	0.206	0.196	*
Most of the time	0.175	0.164	*	0.195	0.176	*	0.185	0.170	*
All the time	0.201	0.234	*	0.353	0.398	*	0.462	0.495	*
Father education: less than primary	0.062	0.076	*	0.034	0.046	*	0.031	0.041	*
Some high school	0.145	0.154	*	0.101	0.110	*	0.098	0.108	*
Completed high school	0.320	0.320		0.285	0.298	*	0.288	0.304	*
Some college	0.156	0.153		0.195	0.191		0.182	0.180	
Completed College	0.190	0.176	*	0.230	0.214	*	0.253	0.225	*
Graduate or professional	0.127	0.121	*	0.155	0.141	*	0.147	0.142	

Table 2. Means of Select Core Variables by Gender - 12th graders

Note: Asterisk indicates statistically significant gender difference at the 5% level.

(continued next page)

Core Variables	1976-1988			1989-1999			2000-2009		
	Boys	Girls		Boys	Girls		Boys	Girls	
Mother education: less than primary	0.032	0.042	*	0.027	0.035	*	0.027	0.034	*
Some high school	0.126	0.149	*	0.082	0.101	*	0.071	0.082	*
Completed high school	0.441	0.416	*	0.339	0.333		0.277	0.280	
Some college	0.166	0.175	*	0.210	0.215		0.210	0.222	
Completed College	0.164	0.146	*	0.234	0.211	*	0.290	0.257	*
Graduate or professional	0.071	0.072		0.108	0.104		0.125	0.124	
Works over school year	0.848	0.798	*	0.801	0.792	*	0.755	0.756	
Average hours of work: None	0.177	0.222	*	0.223	0.223		0.271	0.260	*
5 or less hours	0.101	0.101		0.097	0.095		0.097	0.096	
6 to 10 hours	0.099	0.103		0.095	0.107		0.097	0.107	*
11 to 20 hours	0.262	0.300	*	0.260	0.303	*	0.252	0.281	*
21 to 30 hours	0.234	0.203	*	0.220	0.205	*	0.194	0.191	
More than 30 hours	0.128	0.071	*	0.104	0.067	*	0.090	0.065	*
Average earnings per week from job:									
None	0.227	0.281	*	0.269	0.283	*	0.311	0.314	
\$1-5	0.037	0.046		0.018	0.022		0.010	0.010	
\$6-10	0.040	0.045		0.025	0.028		0.036	0.044	*
\$11-50	0.289	0.326	*	0.176	0.215	*	0.114	0.140	*
\$51-75	0.253	0.200	*	0.128	0.151	*	0.086	0.106	*
\$76-125	0.106	0.080	*	0.222	0.209	*	0.197	0.213	*
\$126+	0.047	0.022	*	0.162	0.092	*	0.246	0.174	*
Type of high school: Academic	0.487	0.514	*	0.550	0.611	*	0.518	0.589	*
General	0.300	0.307	*	0.283	0.272	*	0.328	0.298	*
Vocational	0.155	0.120	*	0.107	0.068	*	0.081	0.049	*
Other	0.059	0.060		0.059	0.049	*	0.073	0.065	*
Education Expectations: index of liken	ess to atte	nd (scaled)	betw	veen 0 and	1)				
Army	0.281	0.102	*	0.215	0.078	*	0.202	0.079	*
Vocational	0.319	0.264	*	0.268	0.210	*	0.274	0.208	*
Two-year college	0.338	0.364	*	0.362	0.370	*	0.383	0.386	
Four-year college	0.584	0.585		0.702	0.758	*	0.737	0.816	*
Graduate or professional	0.389	0.385	*	0.471	0.530	*	0.490	0.571	*
Education Aspirations: want to attend (	binary du	mmy)							
Army	0.203	0.092	*	0.177	0.079	*	0.179	0.078	*
Vocational	0.284	0.219	*	0.207	0.141	*	0.203	0.124	*
Two-year college	0.206	0.293	*	0.214	0.256	*	0.240	0.266	*
Four-year college	0.635	0.650	*	0.744	0.810	*	0.773	0.850	*
Graduate or professional	0.416	0.432	*	0.529	0.613	*	0.519	0.625	*
Number of observations	74230	79942		60469	66875		50549	57202	

Table 2. Means of Select Core Variables by Gender - 12th graders (continued)

Note: Asterisk indicates statistically significant gender difference at the 5% level.

Dependent variable: A (93-100)	Specification 1			Specification 2				
Explanatory Variables	Boys	•	Girls		Boys	•	Girls	
Race: Black	-7.534	(-13.06)	-10.330	(-18.51)	-5.828	(-10.37)	-8.729	(-16.35)
SMSA	-5.404	(-14.96)	-8.525	(-22.02)	-5.270	(-15.02)	-8.227	(-22.23)
Subjective School Ability				· /	44.457	(53.75)	70.881	(72.68)
Smoked cigarettes per day: None (	base)					× ,		· /
Less than one-half pack	-4.928	(-12.39)	-7.789	(-17.68)	-3.635	(-9.39)	-6.107	(-14.48)
One to 1½ pack	-3.911	(-4.23)	-7.383	(-6.64)	-2.513	(-2.79)	-5.382	(-5.06)
Two packs or more	13.153	(5.27)	-1.031	(-0.25)	14.792	(6.09)	2.882	(0.74)
Alcohol binging last 2 weeks: Nor	e (base)	. ,		. ,		. ,		. ,
Once	-4.087	(-7.92)	-4.502	(-7.94)	-3.991	(-7.96)	-3.864	(-7.13)
Two to nine times	-4.984	(-12.01)	-5.247	(-9.81)	-4.664	(-11.57)	-4.007	(-7.83)
Ten or more times	-4.409	(-3.68)	-6.093	(-2.77)	-3.398	(-2.92)	-3.150	(-1.50)
Siblings not same household	0.078	(0.21)	0.243	(0.62)	0.382	(1.08)	0.295	(0.78)
Siblings: One (base)								
None	1.491	(2.10)	-1.119	(-1.38)	1.166	(1.69)	-1.847	(-2.39)
Two	-0.796	(-2.04)	-1.107	(-2.59)	-0.624	(-1.65)	-0.914	(-2.24)
Three or more	-1.615	(-4.17)	-1.797	(-4.36)	-1.493	(-3.96)	-1.476	(-3.75)
Don't know	-0.909	(-0.32)	-8.712	(-2.53)	1.816	(0.65)	-5.119	(-1.55)
Father not same household	-1.228	(-3.07)	-2.319	(-5.69)	-0.870	(-2.24)	-1.895	(-4.87)
Mother not same household	0.169	(0.31)	-1.853	(-3.05)	0.620	(1.16)	-1.460	(-2.51)
Mom working: No (base)								
Some of the time	-3.779	(-7.32)	-2.718	(-4.81)	-3.433	(-6.84)	-2.360	(-4.37)
Most of the time	-4.192	(-7.87)	-4.730	(-8.07)	-3.593	(-6.94)	-3.742	(-6.68)
All the time	-3.855	(-8.26)	-4.764	(-9.46)	-3.513	(-7.75)	-3.844	(-7.98)
Father education: less than primar	0.170	(0.16)	-2.468	(-2.52)	1.407	(1.36)	-0.826	(-0.88)
Some high school	-1.731	(-2.98)	-1.728	(-2.92)	-1.265	(-2.24)	-0.959	(-1.69)
Completed high school (base)								
Some college	0.358	(0.78)	0.581	(1.19)	0.003	(0.01)	-0.201	(-0.43)
Completed College	0.910	(2.04)	2.883	(6.00)	0.133	(0.31)	1.708	(3.72)
Graduate or professional	2.635	(4.66)	2.917	(4.88)	1.074	(1.95)	0.408	(0.71)
Mother education: less than prima	-1.717	(-1.50)	-3.926	(-3.63)	-0.144	(-0.13)	-2.827	(-2.73)
Some high school	-2.298	(-3.45)	-2.685	(-4.02)	-1.876	(-2.90)	-2.068	(-3.24)
Completed high school (base)								
Some college	-1.215	(-2.73)	0.310	(0.67)	-1.480	(-3.42)	-0.169	(-0.38)
Completed College	1.420	(3.30)	2.945	(6.27)	0.662	(1.58)	1.445	(3.22)
Graduate or professional	0.789	(1.34)	1.973	(3.16)	-0.178	(-0.31)	0.408	(0.68)

Table 3a. Coefficients of LPM on Specific Grades - 12th graders - 2000-2009

Note: Dependent variables is set to 100 if the student has a GPA of 4, and to 0 otherwise. T-statistics are in parentheses.

(continued next page)

						, 		
Dependent variable: A (93-100)		Specifi	cation 1			Specific		
Explanatory Variables	Boys		Girls		Boys		Girls	
Works over school year	-3.092	(-2.55)	0.325	(0.22)	-2.206	(-1.87)	2.141	(1.50)
Average hours of work: None								
5 or less hours	4.386	(3.96)	5.993	(4.26)	2.817	(2.62)	3.205	(2.39)
6 to 10 hours	1.529	(1.39)	2.543	(1.84)	0.687	(0.64)	0.744	(0.56)
11 to 20 hours	-0.607	(-0.58)	0.227	(0.17)	-1.310	(-1.29)	-0.955	(-0.74)
21 to 30 hours	-0.279	(-0.26)	-0.261	(-0.19)	-0.758	(-0.73)	-1.564	(-1.19)
More than 30 hours	2.062	(1.82)	-0.511	(-0.35)	1.061	(0.97)	-1.793	(-1.28)
Average earnings per week from								
job: None								
\$1-5	4.664	(2.82)	3.866	(2.20)	5.648	(3.52)	3.473	(2.07)
\$6-10	2.625	(2.60)	-2.200	(-2.20)	3.175	(3.23)	-1.550	(-1.62)
\$11-50	2.002	(2.67)	0.192	(0.25)	1.724	(2.37)	-0.249	(-0.34)
\$51-75	1.681	(2.04)	-1.556	(-1.86)	1.057	(1.32)	-2.085	(-2.60)
\$76-125	0.242	(0.32)	-2.103	(-2.69)	-0.054	(-0.07)	-2.744	(-3.67)
\$126+	-0.238	(-0.32)	-2.172	(-2.65)	-0.743	(-1.02)	-2.915	(-3.72)
Type of high school: Academic	5.905	(9.39)	9.090	(12.99)	1.287	(2.08)	2.010	(2.97)
General	-1.433	(-2.30)	-0.288	(-0.41)	-2.827	(-4.65)	-1.885	(-2.78)
Vocational	2.468	(3.12)	4.776	(4.87)	-0.251	(-0.33)	1.334	(1.42)
Other (base)								
Education Aspirations: want to att	tend (bina	ry dumm	y)					
Army					-2.522	(-5.06)	-2.214	(-3.11)
Vocational					-0.138	(-0.29)	1.152	(1.89)
Two-year college					0.235	(0.54)	-0.198	(-0.42)
Four-year college					-1.739	(-3.98)	-1.321	(-2.50)
Graduate or professional					1.644	(4.34)	2.071	(5.04)
Educational Expectations: index of	of likeness	s to attend						
Army	-2.328	(-4.47)	-0.132	(-0.16)	0.310	(0.47)	2.342	(2.40)
Vocational	-3.945	(-7.18)	-3.929	(-6.52)	-3.522	(-5.62)	-3.938	(-5.67)
Two-year college	-9.946	(-21.99)	-11.409	(-24.47)	-7.536	(-14.45)	-7.042	(-12.68)
Four-year college	3.672	(6.07)	4.384	(6.40)	0.150	(0.23)	0.572	(0.77)
Graduate or professional	13.711	(24.63)	10.543	(18.49)	8.660	(14.07)	4.102	(6.43)
Constant	25.242	(28.44)	28.758	(29.58)	-1.310	(-1.25)	-13.363	(-11.36)
R-squared	0.116	•	0.126	-	0.166		0.202	-
Number of observations	49328		56156		49328		56156	
			~~					

Table 3a. Coefficients of LPM on Specific Grades - 12th graders - 2000-2009 (continued)

Note: Dependent variables is set to 100 if the student has a GPA of 4, and to 0 otherwise. T-statistics are in parentheses.

Dependent variable: C+(77-79)	Specification 1			Specification 2				
Explanatory Variables	Boys		Girls		Boys		Girls	
Race: Black	6.870	(13.71)	5.158	(14.46)	6.020	(12.08)	4.699	(13.28)
SMSA	2.864	(9.13)	2.512	(10.15)	2.792	(8.97)	2.454	(10.00)
Subjective School Ability					-20.698	(-28.22)	-20.586	(-31.84)
Smoked cigarettes per day: None (	base)							
Less than one-half pack	1.865	(5.40)	2.941	(10.44)	1.275	(3.71)	2.431	(8.69)
One to 1½ pack	2.579	(3.21)	2.612	(3.67)	1.949	(2.44)	1.992	(2.82)
Two packs or more	-0.091	(-0.04)	-0.064	(-0.02)	-0.840	(-0.39)	-1.334	(-0.52)
Alcohol binging last 2 weeks: Non	e (base)							
Once	0.752	(1.68)	0.703	(1.94)	0.698	(1.57)	0.525	(1.46)
Two to nine times	1.533	(4.26)	1.709	(5.00)	1.373	(3.84)	1.362	(4.02)
Ten or more times	1.828	(1.75)	4.423	(3.14)	1.381	(1.34)	3.540	(2.54)
Siblings not same household	0.019	(0.06)	-0.294	(-1.17)	-0.118	(-0.38)	-0.318	(-1.27)
Siblings: One (base)								
None	-0.701	(-1.14)	0.175	(0.34)	-0.556	(-0.91)	0.372	(0.73)
Two	0.308	(0.91)	0.349	(1.28)	0.222	(0.66)	0.306	(1.13)
Three or more	0.497	(1.48)	0.440	(1.67)	0.443	(1.33)	0.350	(1.34)
Don't know	1.196	(0.48)	-0.202	(-0.09)	-0.051	(-0.02)	-1.239	(-0.57)
Father not same household	1.044	(3.00)	1.728	(6.64)	0.870	(2.52)	1.587	(6.15)
Mother not same household	0.579	(1.21)	-0.051	(-0.13)	0.392	(0.83)	-0.175	(-0.45)
Mom working: No (base)		. ,		. ,		. ,		. ,
Some of the time	0.068	(0.15)	0.841	(2.33)	-0.095	(-0.21)	0.750	(2.10)
Most of the time	1.148	(2.48)	1.641	(4.38)	0.861	(1.87)	1.373	(3.70)
All the time	1.542	(3.81)	1.509	(4.69)	1.380	(3.43)	1.269	(3.97)
Father education: less than primar	1.164	(1.26)	1.444	(2.31)	0.616	(0.67)	0.945	(1.52)
Some high school	0.826	(1.64)	2.117	(5.59)	0.624	(1.25)	1.876	(5.00)
Completed high school (base)		× ,		· · ·		× /		· /
Some college	0.153	(0.38)	-0.223	(-0.71)	0.311	(0.78)	0.015	(0.05)
Completed College	-0.302	(-0.78)	-0.222	(-0.72)	0.062	(0.16)	0.123	(0.40)
Graduate or professional	-0.444	(-0.90)	-0.382	(-1.00)	0.296	(0.61)	0.331	(0.87)
Mother education: less than prima		(-0.39)	0.172	(0.25)	-1.139	(-1.16)	-0.191	(-0.28)
Some high school	1.500	(2.59)	0.712	(1.67)	1.290	(2.25)	0.547	(1.29)
Completed high school (base)	1.200	(,)	0., <b>1</b> 2	(1.07)		()	0.017	()
Some college	-0.717	(-1.86)	-0.675	(-2.27)	-0.593	(-1.55)	-0.529	(-1.79)
Completed College	-0.977	(-2.61)	-1.504	(-5.01)	-0.631	(-1.70)	-1.073	(-3.60)
Graduate or professional	-1.299	(-2.55)	-1.062	(-2.66)	-0.852	(-1.69)	-0.620	(-1.57)
Note: Dependent spricklas is set to 100						( )		

Table 3b. Coefficients of LPM (100%) on Specific Grades - 12th graders - 2000-2009

Note: Dependent variables is set to 100 if the student has a GPA of 2.3, and to 0 otherwise. T-statistics are in parentheses.

(continued next page)

Dependent variable: C+(77-79)	Specification 1			Specifi	cation 2			
Explanatory Variables	· · · · · · · · ·			Boys	~ [	Girls		
r								
Works over school year	0.996	(0.95)	2.283	(2.40)	0.601	(0.58)	1.713	
Average hours of work: None								
5 or less hours	-2.086	(-2.17)	-3.144	(-3.50)	-1.362	(-1.43)	-2.317	
6 to 10 hours	-2.198	(-2.31)	-3.021	(-3.41)	-1.796	(-1.90)	-2.481	
11 to 20 hours	-1.714	(-1.89)	-2.463	(-2.87)	-1.383	(-1.53)	-2.084	
21 to 30 hours	-1.678	(-1.81)	-2.115	(-2.42)	-1.437	(-1.56)	-1.700	
More than 30 hours	-0.746	(-0.76)	-2.652	(-2.83)	-0.263	(-0.27)	-2.229	
Average earnings per week from								
job: None								
\$1-5	1.578	(1.10)	-0.816	(-0.73)	1.111	(0.78)	-0.637	
\$6-10	-0.270	(-0.31)	-0.211	(-0.33)	-0.507	(-0.58)	-0.404	
\$11-50	-0.367	(-0.56)	-0.483	(-1.00)	-0.244	(-0.38)	-0.343	
\$51-75	0.031	(0.04)	-0.546	(-1.02)	0.288	(0.41)	-0.392	
\$76-125	-0.350	(-0.54)	-0.605	(-1.21)	-0.225	(-0.35)	-0.407	
\$126+	-0.406	(-0.62)	-0.205	(-0.39)	-0.159	(-0.25)	-0.002	
Type of high school: Academic	-4.472	(-8.19)	-5.978	(-13.36)	-2.344	(-4.28)	-3.900	
General	-0.181	(-0.33)	-1.912	(-4.21)	0.450	(0.83)	-1.421	
Vocational	-1.720	(-2.50)	-1.507	(-2.41)	-0.456	(-0.67)	-0.635	
Other (base)								
Education Aspirations: want to at	tend (bina	ry dummy	/)					
Army					-0.290	(-0.66)	0.287	
Vocational					0.241	(0.56)	0.733	
Two-year college					-0.285	(-0.75)	0.488	
Four-year college					1.298	(3.35)	-0.761	
Graduate or professional					-0.773	(-2.30)	-0.147	
Educational Expectations: index	of likeness	s to attend						
Army	2.328	(5.15)	0.469	(0.89)	2.296	(3.96)	0.134	
Vocational	0.473	(0.99)	0.876	(2.27)	0.125	(0.23)	0.175	
Two-year college	5.951	(15.15)	3.446	(11.57)	4.884	(10.56)	1.995	
Four-year college	-5.343	(-10.17)	-3.373	(-7.70)	-4.019	(-6.92)	-1.387	
Graduate or professional	-5.101	(-10.55)	-3.123	(-8.57)	-2.762	(-5.06)	-1.606	
Constant	7.884	(10.23)	6.896	(11.10)	20.066	(21.54)	19.544	
R-squared	0.051		0.051		0.066		0.069	
Number of observations	49328		56156		49328		56156	

#### Table 3b. Coefficients of LPM on Specific Grades - 12th graders - 2000-2009 (continued)

Note: Dependent variables is set to 100 if the student has a GPA of 2.3, and to 0 otherwise. T-statistics are in parentheses.

12th graders	Specif	ication 1	Specif	fication 2
A: 1976-1988	A (93-100): 4	C+ (77-79): 2.3	A (93-100): 4	C+ (77-79): 2.3
Total Differential	3.747 (0.005)	-4.429 (0.007)	3.747 (0.005)	-4.429 (0.007)
Total Explained	0.787 (0.003)	-0.321 (0.004)	0.549 (0.003)	-0.097 (0.004)
Race, SMSA	-0.151 (0.000)	0.234 (0.001)	-0.095 (0.000)	0.190 (0.001)
Own School Ability			-0.081 (0.001)	0.068 (0.001)
Smoking, Binging	0.602 (0.001)	-0.402 (0.002)	0.479 (0.001)	-0.300 (0.002)
Family Background	-0.083 (0.001)	0.004 (0.001)	-0.032 (0.001)	-0.033 (0.001)
Work	0.143 (0.001)	0.024 (0.001)	0.149 (0.001)	0.025 (0.001)
Type of High School	0.065 (0.001)	-0.098 (0.001)	0.006 (0.000)	-0.044 (0.001)
Educ. Expectations	0.211 (0.002)	-0.083 (0.003)	0.123 (0.002)	-0.004 (0.003)
Specification Error	0.001 (0.005)	0.152 (0.007)	-0.011 (0.005)	0.172 (0.007)
B: 1989-1999	A (93-100): 4	C+ (77-79): 2.3	A (93-100): 4	C+ (77-79): 2.3
Total Differential	4.711 (0.006)	-3.898 (0.005)	4.711 (0.006)	-3.898 (0.005)
Total Explained	1.499 (0.003)	-0.713 (0.003)	0.517 (0.003)	-0.103 (0.003)
Race, SMSA	-0.258 (0.001)	0.284 (0.001)	-0.212 (0.001)	0.257 (0.001)
Own School Ability			-0.355 (0.002)	0.209 (0.001)
Smoking, Binging	0.560 (0.001)	-0.246 (0.001)	0.427 (0.001)	-0.167 (0.001)
Family Background	-0.145 (0.001)	0.062 (0.001)	-0.058 (0.001)	0.013 (0.001)
Work	-0.010 (0.001)	0.002 (0.001)	0.019 (0.001)	-0.016 (0.001)
Type of High School	0.220 (0.001)	-0.237 (0.001)	0.031 (0.001)	-0.121 (0.001)
Educ. Expectations	1.132 (0.002)	-0.579 (0.002)	0.665 (0.002)	-0.279 (0.002)
Specification Error	0.114 (0.006)	-0.085 (0.007)	0.030 (0.006)	0.021 (0.006)
C: 2000-2009	A (93-100): 4	C+ (77-79): 2.3	A (93-100): 4	C+ (77-79): 2.3
Total Differential	6.063 (0.007)	-3.152 (0.005)	6.063 (0.007)	-3.152 (0.005)
Total Explained	2.395 (0.004)	-1.224 (0.003)	1.034 (0.004)	-0.590 (0.003)
Race, SMSA	-0.293 (0.001)	0.254 (0.001)	-0.232 (0.001)	0.224 (0.001)
Own School Ability			-0.463 (0.002)	0.215 (0.001)
Smoking, Binging	0.562 (0.001)	-0.198 (0.001)	0.489 (0.001)	-0.163 (0.001)
Family Background	-0.348 (0.001)	0.194 (0.001)	-0.233 (0.001)	0.143 (0.001)
Work	0.060 (0.001)	-0.026 (0.001)	0.079 (0.001)	-0.038 (0.001)
Type of High School	0.385 (0.001)	-0.257 (0.001)	0.185 (0.001)	-0.163 (0.001)
Educ. Expectations	2.029 (0.003)	-1.192 (0.003)	1.207 (0.003)	-0.809 (0.003)
Specification Error	-0.005 (0.008)	0.102 (0.006)	-0.048 (0.007)	0.112 (0.006)

Table 4a. Detailed Decomposition Results - Composition Effects Percentage Female/Male Difference for Selected GPA Levels

Note: Standard errors are in parentheses. Reweighted decomposition follows methodology of section 4. In specification 2, educational aspirations are included among the variables in the educational expectations category.

Percentage Female/Male Difference for Selected GPA Levels									
12th graders	Specifi	ication 1	Specif	fication 2					
A: 1976-1988	A (93-100): 4	C+ (77-79): 2.3	A (93-100): 4	C+ (77-79): 2.3					
Total Differential	3.747 (0.005)	-4.429 (0.007)	3.747 (0.005)	-4.429 (0.007)					
Total Unexplained	3.099 (0.005)	-4.294 (0.005)	3.200 (0.004)	-4.435 (0.005)					
Race, SMSA	-1.243 (0.007)	0.222 (0.008)	-1.290 (0.013)	0.064 (0.008)					
Own School Ability			0.493 (0.001)	0.002 (0.000)					
Smoking, Binging	-0.701 (0.004)	0.166 (0.004)	-0.240 (0.005)	0.106 (0.004)					
Family Background	1.119 (0.018)	-1.830 (0.021)	1.196 (0.026)	-1.549 (0.021)					
Work	0.412 (0.009)	0.766 (0.011)	1.569 (0.012)	0.899 (0.010)					
Type of High School	1.700 (0.019)	-1.618 (0.021)	1.381 (0.027)	-1.758 (0.021)					
Educ. Expectations	0.006 (0.002)	-0.179 (0.002)	0.796 (0.022)	0.660 (0.012)					
Constant	1.807 (0.029)	-1.821 (0.033)	1.501 (0.046)	-2.859 (0.035)					
Reweighting Error	-0.139 (0.001)	0.034 (0.001)	-0.009 (0.002)	-0.068 (0.001)					
B: 1989-1999	A (93-100): 4	C+ (77-79): 2.3	A (93-100): 4	C+ (77-79): 2.3					
Total Differential	4.711 (0.006)	-3.898 (0.005)	4.711 (0.006)	-3.898 (0.005)					
Total Unexplained	3.035 (0.006)	-3.204 (0.005)	3.885 (0.005)	-3.627 (0.005)					
Race, SMSA	-1.772 (0.010)	-0.419 (0.009)	-2.250 (0.010)	-0.342 (0.009)					
Own School Ability			-0.046 (0.000)	0.008 (0.000)					
Smoking, Binging	-0.582 (0.004)	-0.020 (0.004)	-0.378 (0.004)	-0.123 (0.004)					
Family Background	-0.493 (0.019)	-0.996 (0.018)	-0.470 (0.018)	-0.937 (0.018)					
Work	0.998 (0.011)	0.081 (0.010)	0.583 (0.011)	0.195 (0.010)					
Type of High School	1.001 (0.026)	-2.365 (0.024)	0.862 (0.025)	-2.819 (0.024)					
Educ. Expectations	0.036 (0.002)	-0.247 (0.002)	0.887 (0.016)	-0.080 (0.015)					
Constant	3.845 (0.036)	0.762 (0.033)	4.695 (0.037)	0.471 (0.036)					
Reweighting Error	0.062 (0.002)	-0.057 (0.001)	0.279 (0.002)	-0.189 (0.001)					
C: 2000-2009	A (93-100): 4	C+ (77-79): 2.3	A (93-100): 4	C+ (77-79): 2.3					
Total Differential	6.063 (0.007)	-3.152 (0.005)	6.063 (0.007)	-3.152 (0.005)					
Total Unexplained	3.593 (0.007)	-1.933 (0.005)	4.802 (0.007)	-2.489 (0.005)					
Race, SMSA	-1.414 (0.013)	-0.493 (0.009)	-1.509 (0.013)	-0.507 (0.009)					
Own School Ability			-0.089 (0.001)	0.005 (0.000)					
Smoking, Binging	-0.467 (0.005)	0.335 (0.003)	-0.333 (0.005)	0.320 (0.003)					
Family Background	1.591 (0.027)	0.313 (0.019)	1.086 (0.025)	0.706 (0.019)					
Work	1.590 (0.013)	0.303 (0.009)	1.616 (0.012)	0.087 (0.009)					
Type of High School	2.646 (0.028)	-0.905 (0.020)	1.406 (0.027)	-1.503 (0.020)					
Educ. Expectations	-0.117 (0.002)	0.153 (0.002)	1.081 (0.021)	-1.331 (0.016)					
Constant	-0.236 (0.043)	-1.639 (0.030)	1.542 (0.045)	-0.267 (0.034)					
Reweighting Error	0.080 (0.003)	-0.097 (0.001)	0.276 (0.003)	-0.185 (0.001)					

Table 4b. Detailed Decomposition Results - Educational Response Effects Percentage Female/Male Difference for Selected GPA Levels

Note: Standard errors are in parentheses. Reweighted decomposition follows methodology of section 4. In specification 2, educational aspirations are included among the variables in the educational expectations category.

Module 4 Variable		-1988	1989-1999		2000-2009	
Kind of work respondent thinks will be doing when						
30	Boys	Girls	Boys	Girls	Boys	Girls
<b>Laborer</b> (Car Washer, Sanitary Worker, Farm Laborer)	0.63	0.09	0.50	0.06	0.64	0.08
<b>Service worker</b> (Cook, Waiter, Barber, Janitor, Gas Station Attendand, Practical Nurse, Beautician)	0.76	4.72	0.74	3.38	1.23	3.48
<b>Operative or semi-skilled worker</b> (Garage Worker, Taxicab, Bus or Truck Driver, Assembly Line Worker, Welder)	5.01	0.37	2.29	0.20	1.82	0.14
Sales clerk in a retail store (Shoe Salesperson, Department Store Clerk, Drug Store Clerk)	0.66	2.13	0.33	0.56	0.41	0.57
<b>Clerical or office worker</b> (Bank Teller, Bookkeeper, Secretary, Typist, Postal Clerk or Carrier, Ticket Agent)	1.69	19.51	1.17	8.69	0.84	2.60
<b>Protective Service</b> (Police Officer, Fireman, Detective)	4.73	1.16	7.43	2.00	6.65	2.48
Military Service	5.95	1.17	5.55	1.38	5.91	1.13
Craftsman or skilled worker (Carpenter,	18.83	0.70	13.37	0.56	11.43	0.48
Electrician, Brick Layer, Mechanic, Machinist, Tool and Die Maker, Telephone Installer)						
Farmer owner or manager	2.72	0.64	1.54	0.52	1.39	0.68
<b>Owner of small business</b> (Restaurant Owner, Shop Owner)	7.15	4.49	8.08	4.79	8.37	6.63
Sales presentative (Insurance Agent, Real Estate Broker, Bond Salesman)	2.21	1.28	2.51	1.31	2.25	1.31
Manager or administrator (Office Manager, Sales Manager, School Administrator, Government Official)	7.91	7.67	7.25	7.07	6.16	4.59
<b>Professional without doctoral degree</b> (Registered Nurse, Librarian, Engineer, Architect, Social Worker, Technician, Accountant, Actor, Artist, Musician)	28.04	34.99	32.09	41.42	36.07	47.58
<b>Professional with doctoral degree</b> (or equiv) (Lawyer, Physician, Dentist, Scientist, College Professor)	13.64	14.29	16.94	25.80	16.46	26.56
Full-time homemaker or housewife	0.06	6.80	0.22	2.25	0.36	1.68
Number of observations	17731	18968	11293	12299	8970	10198

## Table A1. Means of Non-Core Variable by Gender - 12th graders

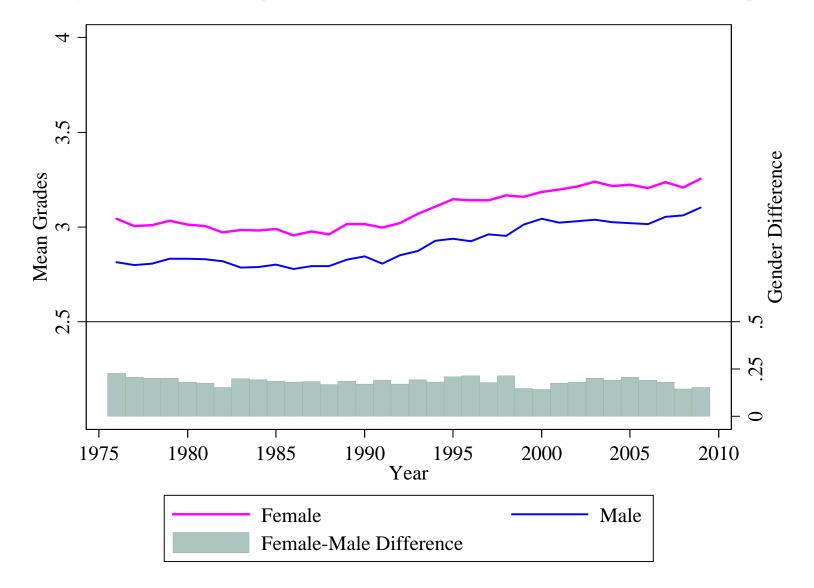


Figure 1. Average Self-Reported Grades of High School Seniors by Gender and Gender Gap

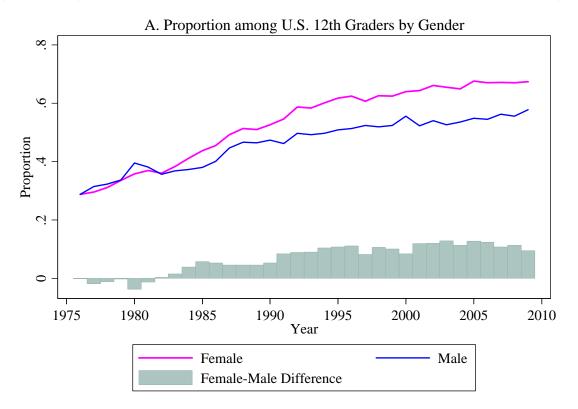
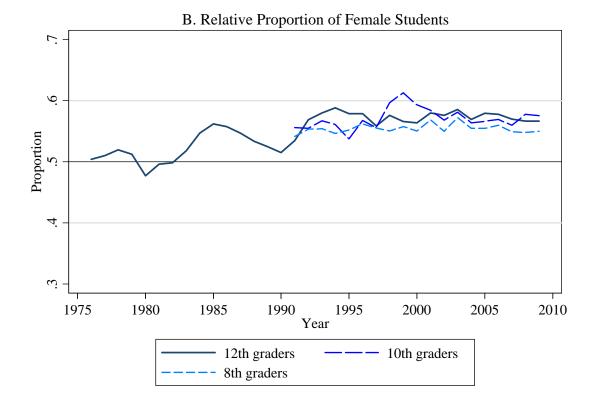


Figure 2. High School Students who say they "Will Definitively Go to College"



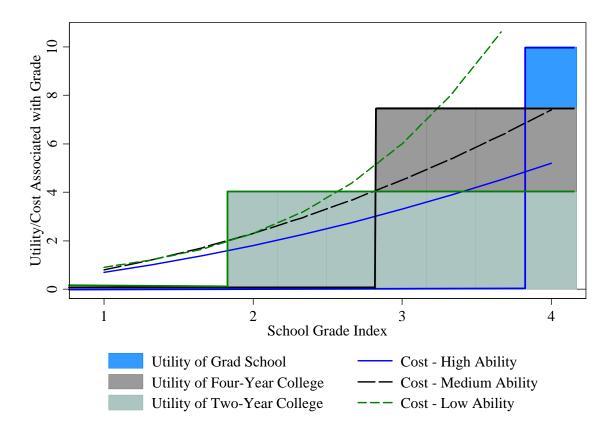
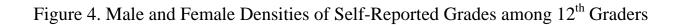
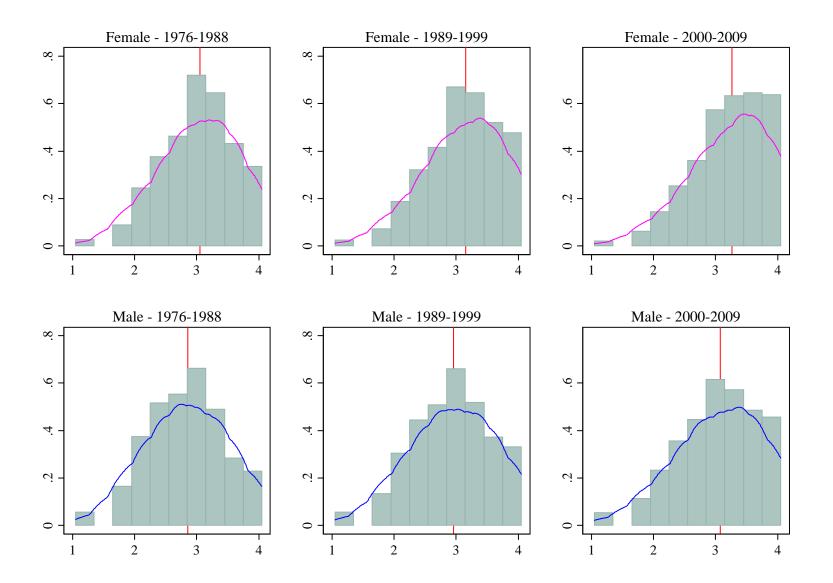


Figure 3. Utility and Cost of Academic Achievement



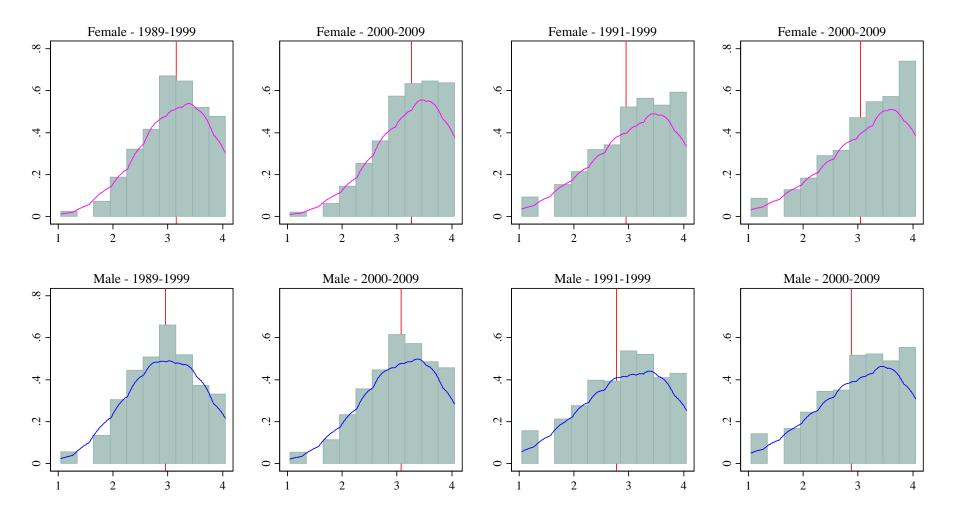


Note: Average grades is indicated by vertical line. Histogram which corresponds to actual data is overlaid with a kernel density.

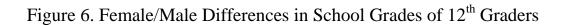
Figure 5. Male and Female Densities of Self-Reported Grades among 10<sup>th</sup> and 8<sup>th</sup> Graders

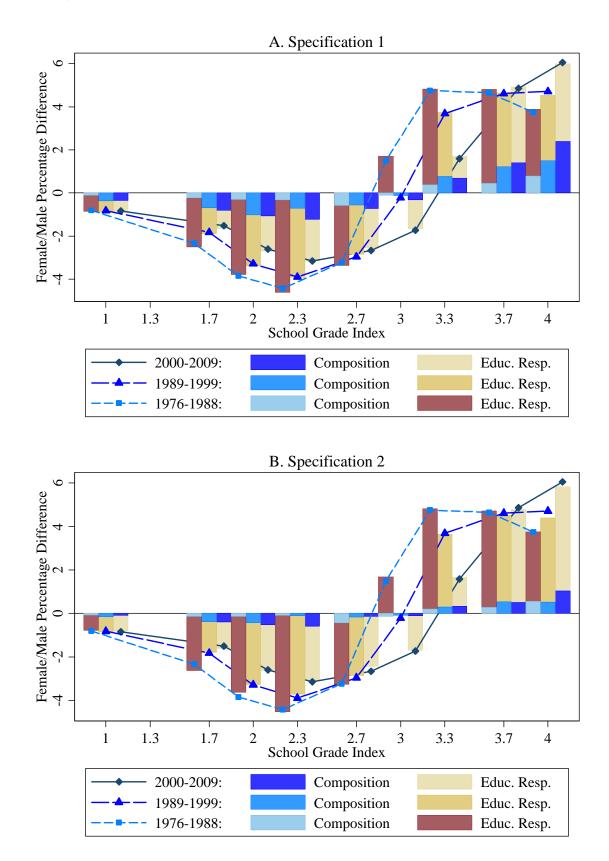


B. 8<sup>th</sup> Graders



Note: Average grades is indicated by vertical line. Histogram which corresponds to actual data is overlaid with a kernel density.





Note: Self-reported grades in 9 categories (D, C-,C,C+,B-,B,B+,A-,A) are translated into the numbers 1, 1.7,2,2.3,2.7,3,3.3,3.7 and 4 following standard institutional practice.

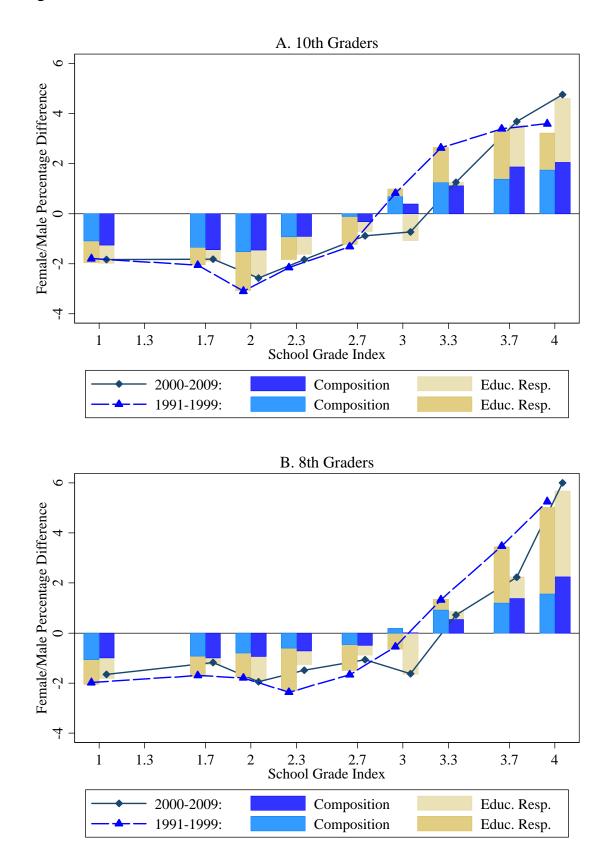
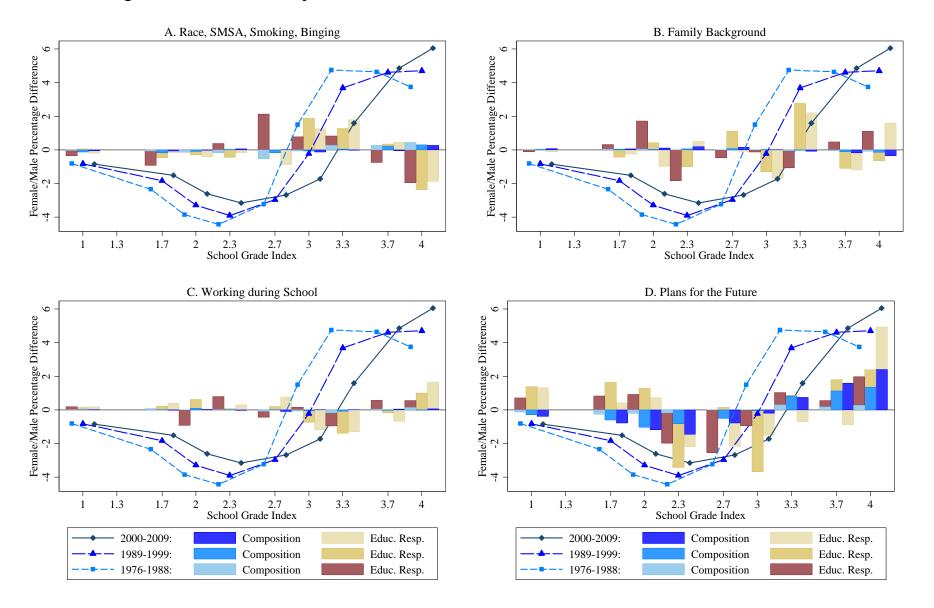
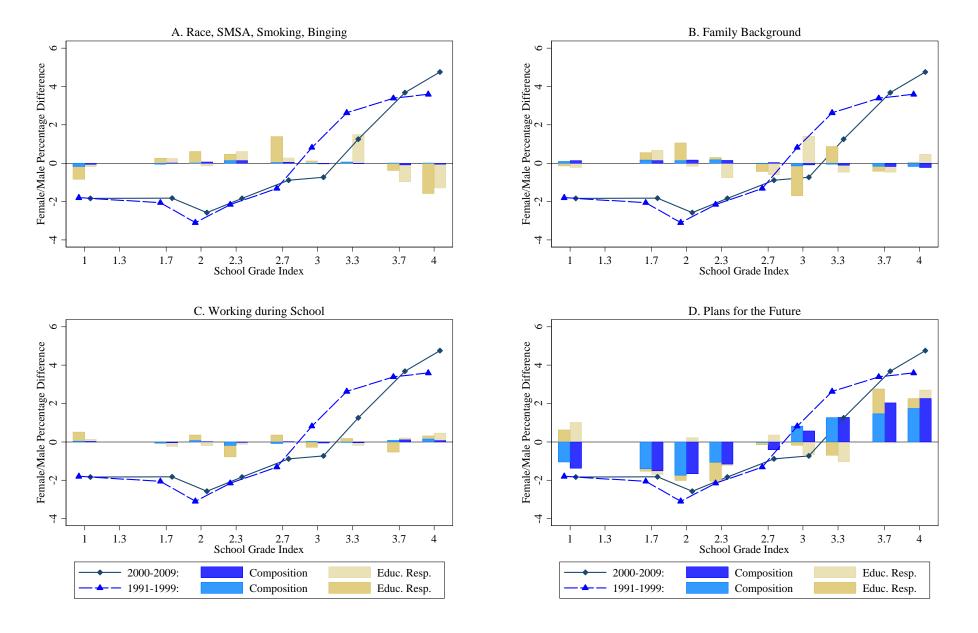


Figure 7. Female/Male Differences in School Grades of 10<sup>th</sup> and 8<sup>th</sup> Graders

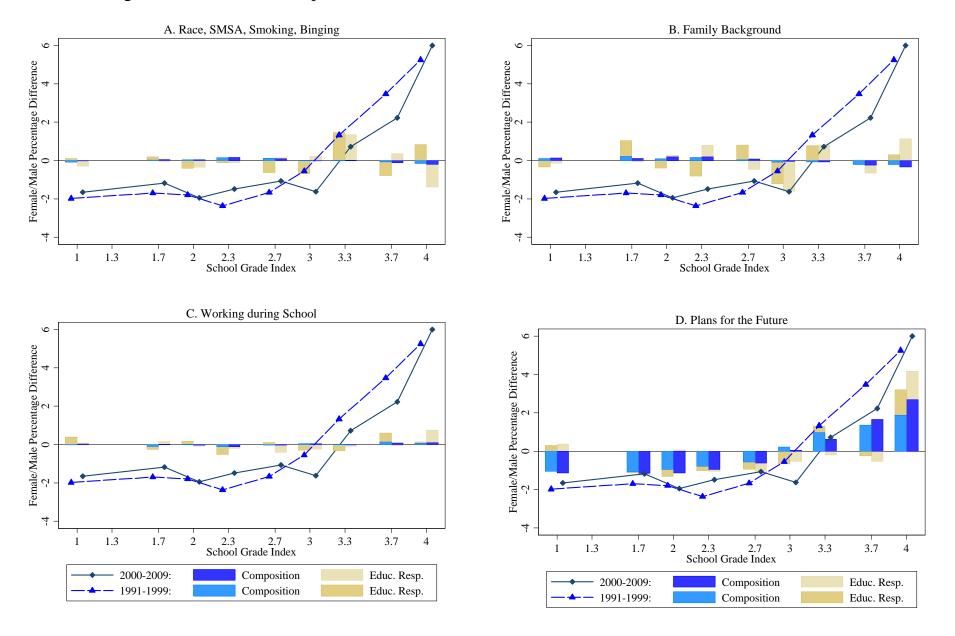
Note: Self-reported grades in 9 categories (D, C-,C,C+,B-,B,B+,A-,A) are translated into the numbers 1, 1.7,2,2.3,2.7,3,3.3,3.7 and 4 following standard institutional practice.



## Figure 8. Detailed Decomposition of Female/Male Differences in School Grades of 12<sup>th</sup> Graders



# Figure 9a. Detailed Decomposition of Female/Male Differences in School Grades of 10<sup>th</sup> Graders



#### Figure 9b. Detailed Decomposition of Female/Male Differences in School Grades of 8<sup>th</sup> Graders