

If financial development matters, then how? National banks in the United States 1870–1900

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Abstract

Despite long-standing interest in the effects of financial development, it has been difficult to determine how banks affect growth since they typically grow with the economy and engage in multiple activities. I consider a time when banks were limited to commercial loans to understand how banks mattered for growth. I construct a novel dataset tracking the size and location of every national bank in the United States from 1870-1900. A large minimum capital requirement meant that otherwise similar counties had very different amounts of banking and I use this discontinuity to estimate the effect of banking on economic development. Even though national banks could not take land as collateral, proximity to a national bank increased agricultural production per capita and tilted the composition of production away from manufacturing. More banking also increased inequality in farm sizes. Additional banking in 1870 still increased incomes 100 years later, suggesting that these results are highly persistent. Although the literature on financial development often focuses on investment as the conduit from finance to growth, this paper points to an alternative: relieving the short-term liquidity needs of commerce.

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1 Introduction

Do banks matter for growth? If so, how? These questions are difficult to answer since banks typically accompany growth and are involved in multiple activities. That makes it difficult to determine not just whether banks help create growth or simply respond to it, but which functions of banks matter for growth. There is a growing consensus that financial development contributes to growth both internationally and within nations.¹ Yet since banks do many things and may alleviate many different types of constraints, it still unclear how banks help growth. Such questions are increasingly important as some of what used to be the main activities of banks are taken over by new entities in both developed and developing countries.²

To answer these questions, I examine a period in United States history when there were strict limits on the activities of banks. From 1870-1900 the US expanded economically and geographically, settling its vast interior. National banks—banks chartered and regulated by the federal government—expanded with the rest of the country and were by far the most important financial institutions in the period. National banks could issue money directly in the form of bank notes. They could not, however, take real estate as collateral, and were limited by the banking practice of the day to make short self-liquidating loans.³ They were thus commercial banks, facilitating trade

¹See [Levine \(2005\)](#) for a summary of recent cross country literature, and the question of whether bank based or financial market based funding matters more across countries. [Burgess and Pande \(2005\)](#) and [Fulford \(2011\)](#) examine the experience of India during its large expansion of branch banking in the 1970s and 1980s. In Italy [Guiso, Sapienza, and Zingales \(2004\)](#) find that local financial institutions aid growth, [Benfratello, Schiantarelli, and Sembenelli \(2008\)](#) show they matter for process innovation, but have little impact on product innovation, and [Pascali \(2011\)](#) examines the long term importance of banks. [Dehejia and Lleras-Muney \(2007\)](#) examine state banking regulations and their effects in the US from 1900-1940. Examining a period when some of the strict limitations on banks used in this paper were relaxed, [Jayaratne and Strahan \(1996\)](#) find positive effects from allowing interstate branching. [Driscoll \(2004\)](#), however, finds that in post-war US data, changes in loans do not affect output at the state level.

²For example, mobile phone payment systems are creating new ways to move funds from place to place and payment clearing, which used to be a key function of banks using their correspondence network. For an example of one rapidly expanding network see [Jack and Suri \(2010\)](#) on the expansion of M-Pesa in Kenya.

³This banking theory was known as the “real-bills” doctrine ([James, 1978](#), p. 59-64). Banks may have occasionally skirted its rules, in particular the self-liquidating requirement, by renewing loans when they became due. Nonetheless, loan maturities were short: [James \(1978](#), p. 61) suggests that the average maturity was about 60 days. Sound banking theory, and the value of commercial banking, were clear at the time. See, for example, the eleventh edition of *Practical Banking* ([Bolles, 1903](#), p. 88): “the first and most important function of a bank is, by the use of the capital which it controls, to bridge over the periods of credit which necessarily intervene between production and consumption, in such a manner as to give back to each producer, or middleman, as quickly as possible, the capital invested by him in

through short-term loans and direct money creation, not investment banks. If these banks mattered for growth during the period, particularly in the agricultural sector, it was because of their commercial, not investment, activities.

To analyze the effects of these banks, I create a rich new data set which gives the exact geographic coordinates and size of every national bank in 1870, 1880, 1890, and 1900. Charged with regulating the national banks and the money they issued, the Comptroller of the Currency collected and published the balance sheet of every national bank each year. Since national banks were not allowed to branch, the place of business listed in the accounts allows me to locate each bank precisely—and to examine local financial development with greater nuance than studies that are limited to regional aggregates. Every decade the census collected detailed data on manufacturing and agricultural output in each county, as well as the amount of land under cultivation and the size of farms. Combining the census data with the location of the banks gives me a detailed panel of banking and output in every county of the US over three decades.

Simply comparing areas that had banks with those that did not does not identify the effect of banking, since counties where banks want to enter are likely to also be areas of high economic activity. Concerned with the stability of the money supply, Congress required national banks to have a large minimum capital. The large minimum size meant that banks were limited in where they could enter profitably: not every county could support a bank of the minimum size, and many banks opened at exactly the minimum size. That meant that some counties had significantly more banking than they would have received if banks were allowed to open at their optimum size; others had much less. How much more banking? How much less? Where and when banks choose to enter from decade to decade and banks' behavior after the minimum capital requirements were reduced in 1900 provide information about the distribution of capital without the minimum capital requirement. I use this distribution to estimate the effect of the extra capital caused by the minimum such products, in order that he may use it again in new production or new purchases.” Since the national banks funded mostly short term loans, it was accepted that they could not meet the investments needs of agriculture, even if they could meet its commercial needs ([Wright, 1922](#), pp. 46, 70).

capital requirement.

Despite the lending limitations, proximity to a bank increased output per capita, largely by increasing agricultural production. For the marginal county, gaining a bank of the minimum size increased total production per capita by 9 percent. Banks were not just following growth, but helping to create it. The mix of production in counties with banks shifted towards agriculture, despite the rapid rise in manufacturing over the period and the limitations on the kind of loans banks could make. Additional capital also seems to have increased the inequality in farm sizes, either by promoting consolidation or more rapid expansion by the largest farmers. There was no effect on yields, but some evidence that the amount of land under cultivation increased. Whether it was from money creation, working capital to farmers, or credit to merchants, it is clear that the commercial activities of banks contributed to growth, even though these banks did not fund capital investments. Moreover, the counties with more banking capital early on had higher incomes even a century later which suggests that the effects of the national banks are extremely persistent.

While the focus of much of financial development theory has been on how financial institutions fund new investments, the commercial activities of banks or other financial institutions, particularly in developing countries, may be equally or even more important.⁴ In 1870, the GDP per capita of the United States would have put it someplace between India and China today ([Maddison, 1995](#), p. 196), and as [Updike \(1985\)](#) argues, the United States shares characteristics with many developing countries today. In particular, the poor transportation infrastructure made getting goods to market costly and time consuming, particularly from rural areas, which remains a problem for many developing countries today ([Bank, 2009](#)). Recent empirical work has also suggested the importance of working capital and the availability of liquidity for small enterprises in the urban areas of developing countries ([Banerjee et al., 2009](#); [de Mel, McKenzie, and Woodruff, 2008](#)).

⁴ [Aghion and Bolton \(1997\)](#) present one version of the constrained entrepreneur. The entrepreneur might be making a human capital investment as in [Banerjee and Newman \(1993\)](#) or in [Galor and Zeira \(1993\)](#) the entrepreneur is someone deciding on an occupational choice. [Galor and Moav \(2004\)](#) provide an even more nuanced growth story, with similar underlying choices. In [Greenwood and Jovanovic \(1990\)](#) and [Townsend and Ueda \(2006\)](#) the entrepreneur faces a risky high return, or a safe low return investment, and financial markets bring diversification. [Banerjee \(2001\)](#) presents a model that nests several versions of credit constraints for an entrepreneur.

While banks may engage in many activities, this paper suggests that facilitating commerce by relieving short-term liquidity needs, whether to the producer or merchant, is a key avenue for financial development to affect growth.

2 Banking and financial markets 1864–1914

A national bank affords a safe place for the deposit of all the little hoards and savings which otherwise would be unemployed. It aggregates these into a fund which becomes useful and powerful in stimulating trade and enterprise.

—Hiland R. Hulbard, Comptroller of the Currency, 1871⁵

This section briefly discusses the financial markets and the economy of the United States from 1864 to 1914, and the literature examining them. While the national banking system was an important factor in early discussions of financial development, more recent literature has largely focused on the system's role in integrating financial markets. Despite the importance of the national banking era in the development of the American economy, this paper is the first to estimate the economic effects of national banks.

The national banking system largely replaced the state-chartered banking system that preceded it. Before the Civil War (1861–1865), there was no national system of banking. States chartered and regulated, or chose not to regulate, their own banks. These state banks issued their own banknotes—bank-issued currency, backed only by the issuing bank's willingness and ability to pay—which circulated widely, and there was no central clearing system, although regional associations of banks created various clearing arrangements (Bodenhorn, 2000). The Civil War (1861–1865) allowed the Republicans in the US Congress, who no longer faced opposition from southern legislators, to move forward in creating a new banking and currency system. The National Currency Act of 1863 and the National Banking Act of 1864 allowed the newly created Comptroller of the Currency to charter national banks, which could issue national bank notes backed by US

⁵Report of the Comptroller of the Currency to the Second Session of the Forty-Second Congress of the United States, 4 December 1871, p. XIII.

treasury bills—in effect allowing these banks to issue and back US currency. State banks were slow to convert to national banks, and so in 1865, Congress passed a new act which established a 10% tax on state banknotes. Not surprisingly, over the next year almost all state banks converted to national banks ([White, 1983](#)).

The goal of the National Banking Acts was to create a uniform bank note currency that would trade at par and to help raise funds for the Federal (Northern) war effort. To help ensure the stability of the new monetary and banking system, the acts imposed several restrictions on the new banks. The acts placed minimum capital requirements to form a national bank: a national bank needed at least \$50,000 in capital to form in a town with no more than 6,000 inhabitants, at least \$100,000 in cities between 6,000 and 50,000, and at least \$200,000 in larger cities ([U.S. Congress, 1864](#), sec. 7).⁶ Moreover, the acts prohibited direct mortgage lending by national banks, and banks could not hold any mortgages obtained indirectly for more than five years ([U.S. Congress, 1864](#), sec. 28).⁷ National banks, and most state banks at the time, were not allowed to set up branches, and so all banks were unit banks, facing the same constraints.⁸

While the larger limits do not appear to have binding, the minimum capital of \$50,000 limited banks' entry into many areas. The laws allowed banks to open before they were fully capitalized, but required them to quickly become fully capitalized. In 1870, 1880, and 1890, every bank that reported less than \$50,000 had attained at least the minimum capital by the next year or

⁶The evidence does not suggest that the limits above \$50,000 were strictly enforced. Between 1870 and 1880, counties with cities with populations between four and six thousand gain somewhat more banks per capita, but the difference is statistically insignificant, despite the doubling of required capital as population increased over six thousand. The capital size appears to have been reasonably easy to circumvent by opening in a nearby town. The city populations are from the census in 1880 which also reports 1870 populations ([Census Office, 1880](#), pp. 416-425). I matched these cities to locations using the same process as with the banks as in appendix C, giving the geographic location of each city, and the 1890 county it falls in. While it seems possible to use the change in capital requirements around 6,000 as a discontinuity to estimate the effect of banking, it does not have good explanatory power.

⁷[Keehn and Smiley \(1977\)](#) suggests that the ban of mortgage lending was not perfect, but was nonetheless extremely restrictive. Loans secured by mortgages or real estate were less than two percent of total value before restrictions were relaxed in 1914, and rose afterwards.

⁸The second comptroller of the currency interpreted the National Banking Act as prohibiting forming branches. Although the exact language of the act does not necessarily prohibit forming branches, the rules stayed in effect until the 1920s ([White, 1983](#), pp. 14-15). State banks were similarly constrained by state laws. For the positive effects of relaxing these laws a century later, see [Jayaratne and Strahan \(1996\)](#).

shut down. As suggested by [Sylla \(1969\)](#), the best evidence that the minimum capital requirement was binding even at the end of the period is what happened when it was loosened. The Gold Standard Act of 1900 reduced the minimum capital requirement to \$25,000 for towns under 3,000, and over the next decade thousands of new national banks were formed with capital below \$50,000.

Immediately after 1865, the number of national banks grew quickly as state banks converted into national banks. Growth in the total number of national banks then slowed, before accelerating in the 1880s during a boom in banking. Figure 1 shows the growth of national and state banks. Until the late 1880s, national banks had few and ineffective competitors. In the late 1880s as deposit banking became more important and states allowed banks to form without a special charter, the number of state banks increased rapidly, filling an apparent void left by the national banking system.⁹ Yet there still seems to have been a strong desire for national banks, as the surge of smaller banks after 1900 suggests. Some of these new national banks may have been former state banks, which did not find it profitable to open with the full \$50,000 in capital as national banks, but did with a smaller required capital.

The total number of banks hides a more complex process of entry and exit and spread of geographic extent. Even as the number of banks was growing, there was substantial exit from decade to decade. While most of the converted state banks were in the Northeast, new banks spread west as the Midwest and Western states and territories became increasingly populated and productive. Figures 2 through 5 show this spread over time and space. Despite the substantial growth in national banks, many counties did not have a national bank even by 1900. In particular, the South, whose banking system had been largely destroyed during the Civil War, and the sparsely populated West, lacked banks.

Deflation and panics were important aspects of the financial markets during the Gilded Age. The United States faced a prolonged period of deflation following the war as it first resumed the

⁹See [James \(1978, pp. 29–39\)](#) and [Barnett \(1911, pp. 11–12, 32-33\)](#) for a discussion of the spread of state banks. After around 1890, many states had less stringent capital requirements than the National Banking Act ([James, 1976c](#)), which allowed state banks to open more easily.

gold standard at pre-war prices and then maintained it, notwithstanding periods of bi-metallism ([Friedman and Schwartz, 1963](#)). Notes issued by different national banks, backed as they were by treasury deposits and uniform regulations, traded at par with each other and with currency issued directly by the government. Although aggregate notes issues were initially capped, the cap was removed in 1875 as part of the Resumption act. National banks generally chose to limit their note issues below the maximum, however, which contributed to the relative scarcity of money ([James, 1976b](#)). There were major banking crises in 1873 and 1893, and smaller disturbances in 1884 and 1890 ([Wicker, 2000](#)). At the end of the period, the crisis of 1907 prompted a reform of the system and the creation of the Federal Reserve. National banks held reserves and interbank deposits in regional reserve city banks, which in turn held reserves in New York (and to a lesser extent Chicago), which made the entire system sensitive to disturbances in New York ([Cagan, 1963](#), pp. 36–37). The 1873 and 1893 crises seem to have accompanied a cyclical downturn ([Wicker, 2000](#), pp. 8–11).

The national banking system played an important role in the early discussion of the importance of money, banking, and credit.¹⁰ One area of particular concern was how poorly the national banks seemed to serve agriculture. For example, [Wright \(1922\)](#) argues that the large minimum size meant that national banks could not profitably enter many rural areas; the prohibition from taking real estate as collateral limited the ability of farmers to borrow; and the requirement to lend only on a short-term basis meant national banks could not fund long-term agricultural investments.

More recent work has focused on the supposed instabilities of the national banking system, which led to the creation of the Federal Reserve in 1913, and its role in the integration of capital markets. This paper is the first to examine empirically whether and how the national banks mattered

¹⁰For example, the *Journal of Political Economy* published a four part series in 1918 , and a comment and reply in 1919, on commercial banking and capital formation ([Moulton, 1918a,b,c,d](#); [Watkins and Moulton, 1919](#)) whose primary source of information about what banks actually do comes from national banks and their regulator the Office of the Comptroller of the Currency. The new *Review of Economics and Statistics* published a four part series from 1924 to 1927 solely on national bank statistics ([Young, 1924, 1925a,b, 1927](#)). One of the most successful textbooks on banking ([Dunbar, 1892](#)) devotes as much attention in its first edition to the national banks of the US as to the banking systems of France, England, Germany, and the Bank of Amsterdam, despite the novelty of the national banking system. The 1917 editions drops the Bank of Amsterdam in favor of discussing the new Federal Reserve.

in the economic growth of the United States. In one of the few papers that considers non-financial effects, [Campen and Mayhew \(1988\)](#) describe the importance of the national banks in Knoxville, Tennessee. Much of the literature on national banks and monetary matters after the Civil War focuses on explaining regional variations in interest rates. [Davis \(1965\)](#) documents that national banks in the Mid-Atlantic region charged a lower average discount rate, as well as had lower returns, than banks in other regions. These gaps seem to have narrowed sometime before 1900, which [Davis \(1965\)](#) attributes to the development of a national commercial paper market which allowed capital to move more easily across regions. While capital flows may have increased from the more developed East, [Sylla \(1969\)](#) suggests that where national banks did exist in rural areas, they could act as monopolists since the minimum capital requirements and branching restrictions made it difficult to acquire sufficient funds to enter. Moreover, in many rural areas the available capital for deposit in a bank was not sufficient to make it profitable for one bank to enter and put up the minimum capital, much less a second one to offer competition. Suggesting that monopoly power may have been important, [James \(1976a\)](#) finds that the number of banks (including state, national, and private banks) per capita at the state level is negatively related to the interest rate between 1893 and 1901, taking into account the risk as measured by the variance of the loss rate. [Binder and Brown \(1991\)](#) provide somewhat more formal tests of the convergence of returns based on the timing of institutional changes, and suggesting that the timing of the 1900 relaxation of national bank capital requirements does not seem to have been important. [Sullivan \(2009\)](#) suggests that profits are a better measure of possible monopoly power than interest rates or returns, and finds that differences in regional profits fell after 1884.

3 Identification strategy

There are two difficult problems to identifying the effects of banking: endogeneity and dynamic effects. Suppose banks can enter freely and we observe more banking activity in wealthier areas. Should we conclude that banking causes wealth, or wealth attracts banking? Most likely the answer

is some of both, but we might still like to know the effect of encouraging or discouraging banking, particularly for marginal areas likely to be affected by the policy. Does forcing, or subsidizing, banks to enter areas they might not otherwise want to enter increase productive activity, and by how much? India, for example, for years maintained a “social banking” policy which forced banks to open branches in rural areas with the express intent of fostering additional access to credit in rural areas and so promote growth ([Panagariya, 2008](#), pp. 224-8). Such thinking is also behind the recent push for subsidized microcredit: the profits are not sufficient to bring in profit maximizing firms, but the benefits to credit are assumed to be large.

I use the observation that the minimum capital requirement forced some areas to have too much banking while restricting others to separate the endogenous effects of banking from the causal effects. I build the argument in pieces, so that the assumptions necessary for identification are clear. First, I show why a minimum capital requirement induces a discontinuity as profit maximizing banks decide where to enter. I next estimate the demand for banking given the observed behavior of banks. I then set up the identification problem for determining the effect of banking on output. The discontinuity induced by the minimum capital requirement would allow identification using standard regression discontinuity arguments if it were possible to observe what banks would have done without the minimum capital requirement. It is still possible to identify the effect of the additional capital caused by the minimum capital requirement by integrating out the unobserved county demand for banking.

3.1 Profit maximizing banks with a minimum capital requirement

To understand how a minimum capital requirement causes a discontinuity, this section takes a simple version of the profit function of a bank with market power considering opening in some county. The monopoly power of “country banks” has been studied by [Sylla \(1969\)](#). Minimum capital requirements mean that some banks open in areas with more than optimal capital, while some areas do not have any banks. Then the key insight into any model of profit maximization is that the fur-

ther away the minimum capital requirement forces a bank to be from its profit maximizing capital, the lower are its profits, until at some point profits are negative, and the bank will not open. So a minimum capital requirement turns the capital in a county from a continuous function of banking conditions in that county to a discontinuous function. While this section describes a particular and simple model which gives a zero profit point, the empirical strategy only depends on being able to express the dividing line between entry and not in terms of capital, not on the particular underlying functions.

Bank profits come from the difference between the return on the loans it makes and the cost of raising the capital to make those loans: $\pi = r_L(L)L - r_B(K)K$ where K is total capital and L is the loans, and banks loan out all of their capital so $K = L$. Banks can raise capital by equity (capital stock) and by collecting deposits. Banks have local market power in loans so that issuing more loans reduces the interest rate the bank can charge. The cost of capital increases with the amount raised.

Banks may raise capital using capital stock (equity) or deposits (debt). To pin down the capital structure, the cost for raising any given level of deposits is decreasing in the amount of capital stock; depositors are aware that banks may make bad loans, and so are willing to accept a lower return on deposits in banks with higher equity. Since the bank is also a local monopsonist for deposits, it has to offer higher interest rates to get more deposits. To fix ideas, write the capital stock as a fraction ω of total capital $C = \omega K$, so deposits are $(1 - \omega)K$. Then the cost of capital $r_B(K) = \rho_c\omega K + (\rho_D - \theta\omega)(1 - \omega)K$ where the dividend rate on the capital stock is $\rho_c\omega K$ and the interest rate the bank must pay on deposits given its capital structure is $(\rho_D - \theta\omega)(1 - \omega)K$. Both are increasing as the bank attempts to raise more funds. Depositors demand a higher rate when the bank is less well capitalized. Cost minimization of the capital structure for a given total capital then implies that $\omega^* = (\rho_D - \rho_C + \theta)/(2\theta)$, so that the capital structure is a fixed constant. More capital stock is expensive but reduces the costs of deposits (θ_D), and the bank chooses ω so that these effects balance out, which is independent of the total capital raised. The relationship

between capital stock and total capital is approximately linear in the data, so it seems like a reasonable way to describe the capital structure. This approach to capital structure gives a simple version of the model, but any assumption that makes debt and equity incomplete substitutes so that the bank decides its size by deciding its capital stock would give similar results.

Local demand for loans is downward sloping, but depends on the business conditions and population in a given county. Counties with a large population engaged in activities which need banking services demand more loans. Then in county c at time t with population P_{ct} the interest rate on loans is $r_L(L) = \alpha_0\alpha_t P_{ct}\eta_c\epsilon_{ct} - \alpha_1 L$. In some counties η_c is high, and the need for banking services is higher, in others η_c is low, and even large populations may demand few loans. The overall demand for loans may change over time, and counties may receive idiosyncratic shocks ϵ_{ct} which change their demand for loans in a given time.

Ignoring minimum capital requirements for the moment, a bank can choose its optimal size and so make money anywhere the demand for loans is positive. Maximizing profits and solving for the optimal capital gives:

$$C_{ct}^* = \frac{\alpha_0\alpha_t P_{ct}\eta_c\epsilon_{ct}}{2(\alpha_1 + \rho_\omega)/\omega^*} = \alpha\alpha_t P_{ct}\eta_c\epsilon_{ct} \quad (1)$$

where α absorbs all of the other model parameters and $\rho_\omega = \rho_c\omega^* + (\rho_D - \theta_D\omega^*)(1 - \omega^*)$. Although the model has been structured as one of demand shocks, since it seems reasonable that demand for banking services varies while the cost of capital may be similar across counties, counties may face different and time varying costs of capital as well. The end result that C_{ct}^* depends on local conditions does not change, only the underlying parameters.

With a minimum capital requirement, the bank must open with at least \$50,000 in capital, and so faces a constrained maximization problem. To simplify notation, I express capital in thousands of dollars. Consider how what a bank actually does in the face of the minimum capital requirement C_{ct} varies depending on what it would like to do if it were unconstrained C_{ct}^* . Figure 6 show four different profit functions which might hold in four different counties or for a county over time. In panel (A), profits for that county are less than 0 if a bank enters with 50 in capital, and so it does not

enter and $C_{ct} = 0$ even though C_{ct}^* is positive. A bank would enter with a small amount of capital if it could, but if forced to choose between opening with 50 and not opening at all chooses not to enter. In panel (B), the bank is just indifferent between entering and not, since profits are zero either way. In panel (C), the bank makes positive profits if it opens at 50 capital, but would prefer to enter with less, and so opens with exactly 50 in capital. In panel (D) the bank is unconstrained and opens with its optimal capital. Using the linear assumptions for supply and demand which creates a quadratic profit function, the point where profits are zero is given when the optimal capital is 25, which is found by solving $\pi(50) = 0$ and substituting C_{ct}^* , but I will let the data decide the best point.

The model illustrates one set of assumptions that give a precise dividing line in terms of what a bank would have wanted to do if it were unconstrained. To generalize, suppose each county can be ranked by an index z of county conditions for banking so that $\pi(50, z)$ is continuous and strictly increasing in z . Then by the implicit function theorem there is a unique solution z_0 to $\pi(50, z) = 0$. Then $C^T = \operatorname{argmax}_C \pi(C, z_0)$ is the optimal unconstrained capital for entering the county where entering with 50 in capital earns zero profits. The restrictions on the profit function amount to assuming that while there may be many different conditions for a county to yield zero profits when entering with 50, all counties for which profits would equal zero have the same index z and the same unconstrained optimal capital C^T .

Having to open with a minimum capital introduces a discontinuity in capital. Counties with an optimal capital just above C^T have a lot more capital than counties with an optimal capital just below C^T . So the actual and observed capital in thousands in a county is a discontinuous function of C_{ct}^* :

$$C_{ct} = 50D(C_{ct}^* \geq C^T) + (C_{ct}^* - 50)D(C_{ct}^* \geq 50),$$

where $D(\cdot)$ is an indicator which is 1 if the condition is true and 0 otherwise. Figure 7 shows this discontinuous function. For very low optimal capital, profits are negative and so no entry occurs. As optimal capital passes C^T , the actual capital jumps to 50, and stays there until the bank is no

longer constrained. Alternatively, define $EC_{ct} = C_{ct} - C_{ct}^*$ as the extra capital (in thousands) that a county has, or the loss of capital from not having bank. It is a discontinuous function of C_{ct}^* which jumps by 50 at C^T .

3.2 Estimating optimal capital

The model of profit maximizing banks suggests that the optimum capital in a county is a log linear form driven by the county population P_{ct} and local unobservable business conditions η_c and local temporal shocks ϵ_{ct} . The previous section shows how to get equation 1 as an exact representation of profit maximizing behavior, but the log linear form is a reasonable reduced form for a wider set of models. Since banks cannot enter unconstrained, the model also gives the relationship between actual capital C_{ct} and optimal capital based on the threshold C^T . The first step to understanding the effects of the national banks is to estimate that relationship.

Counties fall in one of three categories, depending on whether there are no banks in the county, any banks with the minimum capital, or only unconstrained banks. Each county may fall into all three categories over the full panel. If there are no banks in a county, then C_{ct}^* must be below C^T ; a county with a bank with 50 capital stock must have optimal capital between C^T and 50. Counties with larger banks have exactly their endogenous optimal capital.

To estimate the relationship I employ a maximum likelihood estimator for the reduced form 1 which accounts for the constraints given by the minimum capital requirement. The MLE assumes that $\ln \epsilon_{ct} \sim N(0, \sigma^2)$ and $\ln \eta_c \sim N(0, \sigma_\eta^2)$ which put positive weight on all possible C_{ct}^* given the observed behavior for a given county. Given a dividing line C^T , I estimate $\hat{\alpha}$, $\hat{\alpha}_t$, $\hat{\sigma}$, and $\hat{\sigma}_\eta$. Maximizing the likelihood is made more difficult by the county level dependence of each time observation on η_c which requires using numerical integration to get the log-likelihood for each county conditional only on the data. Appendix A constructs the likelihood function and discusses estimation. Finally, I choose the threshold value \hat{C}^T by a grid search which maximizes the likelihood.

3.3 The identification problem

To see why the minimum capital requirement is useful, first suppose banks are unconstrained and can enter with whatever capital they find most profitable. Suppose some vector Z_{ct} determines the profitability of county c at time t for banking. So Z_{ct} might include how good the farm land is, transportation, weather, and the cultural, religious, or institutional composition of the county. While some parts of Z_{ct} may be observable, it is not possible to observe all of the important elements. The amount of banking capital in a county is thus some function of this unobservable variable: $C_{ct}^* = C^*(Z_{ct}, Y_{ct})$ and output Y_{ct} . Output is also determined by Z_{ct} or its components: $Y_{ct} = Y(Z_{ct}, C_{ct}^*)$. In addition, especially during a period when population movement was important, population P_{ct} depends on Z_{ct} as people move seeking good opportunities just as capital does. [Atack et al. \(2009\)](#), for example, examine the relationship between transportation and urbanization in the several Midwestern states as the railroads came.

Then if we wanted to know how much an increase in banking capital in a county caused economic output to increase, a regression in per capita terms of the form:

$$Y_{ct}/P_{ct} = \theta_c + \theta_t + \gamma_e C_{ct}^*/P_{ct} + U_{ct} \quad (2)$$

where γ_c and γ_t are county and time fixed effects, would not be informative. Since both Y_{ct} , and C_{ct}^* are related to the unobserved Z_{ct} and to each other, the estimate of γ_e is endogenous and does not have a causal interpretation (hence the subscript). It has bias as a causal estimate of the effect of capital on output both from the simultaneous equations and the unobserved variables. To take an extreme example, banks may allow farmers to drink away next year's crop, and so not be able to work hard the next day. In that case, banks actually reduce output, but there will still be more banking capital in areas where farmers have larger crops to drink away, so the estimate of γ_e will still be positive, even though the causal effect of banking is actually negative. The problem may become even more acute if current banking responds to predictions of future economic activity.

The central empirical difficulty of how finance affects development is how to estimate the effect of banking on economic activity so that it has a causal interpretation.¹¹

The second major difficulty in estimating the effect of finance or banking is that the effects may vary over time. Credit, by definition, allows some people or firms to bring forward investment or consumption, while others to delay it. Relieving credit constraints, or introducing a new savings option, is likely to have effects that vary over time, possibly dramatically, with what holds in the long term the opposite of what holds in the short term. Usefully, the problem is reasonably easy to deal with: [Fulford \(2011\)](#) shows that including past values of the banking variable deals with the problem by allowing the effects to vary over time. Whether or not lags matter depends on what banks do, and how people use banks, and so the structure of the lags reveals a great deal about the effect of banking.

3.4 Using the discontinuity

Suppose we could assign some places more banking (or less banking) than those counties would get based on their level of the unobserved Z_{ct} . The extra banking can then identify the causal effect of banking as long as the extra banking is not related to the unexplained economic activity U_{ct} . The minimum capital requirement that banks could only start with a minimum of \$50,000 meant that some counties received much more capital than they would have gotten without the requirements since banks, which would have entered with a profit maximizing capital less than \$50,000, instead entered with the minimum capital. Some counties, on the other hand, were denied banking that they would have had since banks did not find it profitable to enter with such a large capital stock. The minimum capital requirement thus causes the capital stock to jump discontinuously from \$0 to 50,000, even though we might expect that underlying economic activity which attracts banks behaves continuously. Small changes in the underlying economic activity thus cause big changes

¹¹See [Levine \(2005\)](#) and [King and Levine \(1993\)](#) for some attempts to deal with this problem comparing countries. [Rajan and Zingales \(1998\)](#) examine industries which are more likely to benefit from finance. [Burgess and Pande \(2005\)](#) use social banking rules in India as an instrument for which districts received more banks.

in the amount of banking.

I use the minimum capital requirement to divide the observed banking into the endogenous banking activity and the extra banking that comes from meeting the minimum capital requirement or the loss of capital and banking activity from not having a bank. From the model of optimizing behavior, the observed capital in a county is $C_{ct} = 50D(C_{ct}^* \geq C^T) + (C_{ct}^* - 50)D(C_{ct}^* \geq 50)$ and it jumps discontinuously at C^T . The extra capital caused by minimum capital requirement is then $EC_{ct} = C_{ct} - C_{ct}^*$. It is positive for counties that get a bank, but negative for those that do not. Then using the conditional expectation from equation 2:

$$E[Y_{ct}^0 | C_{ct}^*, P_{ct}] = \theta_c + \theta_t + \gamma_e C_{ct}^* / P_{ct} \quad (3)$$

$$Y_{ct}^1 = Y_{ct}^0 + \gamma EC_{ct} / P_{ct}. \quad (4)$$

For the moment, suppose it is possible to observe C_{ct}^* , and the actual capital C_{ct} . This system is shown graphically in figure 8. As C^* increases, Y increase continuously as well without the minimum capital requirement. With the minimum capital requirement Y jumps discontinuously at C^T by 50γ , and above and below increases at the rate $\gamma_e - \gamma$ which may be positive or negative (it is negative in the figure, and the estimates suggest it is in the data as well). This system assumes that the extra capital affects output, but that output and other unobserved covariates affect EC_{ct} only through C_{ct}^* .

While there is still not a causal interpretation of γ_e, γ as the effect of treating counties with optimal capital close to C^T with 50 in extra capital. As long as the density of U conditional on C^* is continuous, the standard sharp regression discontinuity result applies here (see, for example, [Imbens and Lemieux \(2008\)](#) or [Hahn, Todd, and Klaauw \(2001\)](#)). Counties just below C^T receive a lot less capital than they would get without the minimum capital requirement, which reduces their output by 50γ relative to those just above.

It is important to make the distinction between banks and counties when considering whether the continuity assumption is reasonable. Banks decide whether or not to enter based on factors

which affect Y_{ct} , but counties do not control banks, and counties are the unit of observation. The key assumption using the language of Lee and Lemieux (2010) is that counties do not have precise control over entry; counties cannot manipulate themselves around the threshold to discontinuously attract banks. Banks have control, but are attracted by county conditions, and it is the continuity of the conditions in counties which matters. Then close to C^T the distribution of the unobservables is the same above and below the threshold, and so the difference between the conditional means identifies γ . It is fine, for example, if banks know that by entering they will raise the output of a county, thus increasing the optimal capital for that county. Banks take that into account when deciding whether to enter, and are willing to enter in less profitable counties, but the counties which are just profitable enough to enter are nearly the same as the counties that are barely not profitable.

Of course, I can only observe the outcome C_{ct} not the underlying C_{ct}^* , but the panel and population of the county give a great deal of information about the optimal capital and I can use this information and still get consistent estimates of γ . The minimum capital requirement gave all counties with a bank that opened at exactly 50 some extra capital; otherwise the bank would have opened with more than 50 in capital. Similarly, all counties without a bank might have supported a bank (or a branch if branching had been allowed) and so were denied some banking. The problem is how much, not whether.

The key realization is that using the functional form of the conditional expectation function, the discontinuity identifies the causal effect γ as long as the endogenous relationship γ_e is estimated consistently. Rewrite the conditional system 3 as a moment condition :

$$E[Z_{ct}U_{ct}] = 0$$

where the instruments are $Z_{ct} = (I_c, I_t, C_{ct}/P_{ct}, C_{ct}^*/P_{ct})$, the I_c and I_t are appropriately sized vectors of indicator variables, and $U_{ct} = Y_{ct} - (\theta_c I_c + \theta_t I_t + \gamma C_{ct}/P_{ct} + (\gamma_e - \gamma)C_{ct}^*/P_{ct})$. If C^* were observable, estimating the coefficients by method of moments or least squares would be equivalent.

The optimal capital estimation provides all of the necessary information to integrate out optimal capital since the underlying unobservables must be consistent with the bank entering or not with a given capital. The estimates $\hat{\alpha}$, $\hat{\alpha}_t$, $\hat{\sigma}$, and $\hat{\sigma}_\eta$, together with population and the observed capital give the distribution the optimal capital consistent with the observed conditions in that county at that time. Call $f_{ct}(C_{ct}^* = w | C_{1870c}, \dots, C_{1900c}, P_{1870c}, \dots, P_{1900c})$ the density of optimal capital. Then by the law of iterated expectations, we can take the expectation of the moment conditions in two steps:

$$E \left[\int_w Z_{ct} U_{ct} f_{ct}(C_{ct}^* = w | C_{1870c}, \dots, C_{1900c}, P_{1870c}, \dots, P_{1900c}) dw \right] = \\ E[E[Z_{ct} U_{ct} | C_{1870c}, \dots, C_{1900c}, P_{1870c}, \dots, P_{1900c}]] = E[Z_{ct} U_{ct}] = 0.$$

The coefficients which set the original moments to zero also set the expected value of the conditional moments to zero.

To understand the mechanism, it is useful to think of the problem as one of measurement error so that results of [Fuller \(1987\)](#) apply. Call the mean of the conditional distribution \hat{C}_{ct}^* . It is specific to each county at each time. Then the realization of optimal capital is related to the mean of its conditional distribution by $\hat{C}_{ct}^* - e_{ct} = C_{ct}^*$ where e_{ct} is the error in replacing the draw C_{ct}^* with the mean of its distribution. For all except the last moment, the integration simply replaces C_{ct}^* with its mean, since by definition the error integrates out. The last moment equation introduces the variance of e_{ct} :

$$E \left[\int_w C_{ct}^*/P_{ct} (Y_{ct} - (\theta_c I_c + \theta_t I_t + \gamma C_{ct}/P_{ct} + (\gamma_e - \gamma) C_{ct}^*/P_{ct})) f_{ct}(w) dw \right] = \\ E \left[\hat{C}_{ct}^* (Y_{ct} - (\theta_c I_c + \theta_t I_t + \gamma C_{ct}/P_{ct} + (\gamma_e - \gamma) \hat{C}_{ct}^*/P_{ct})) - (\gamma_e - \gamma) \sigma_{ect}^2 \right] = 0$$

where σ_{ect}^2 is the variance of the error in the conditional distribution. If σ_{ect}^2 were constant across counties, then the estimator would be exactly the same as the standard measurement error correc-

tion when the measurement error variance is known [Fuller \(1987\)](#), pp. 13–29). Since the variance may vary with the observation, it is a simpler version of the model in [Fuller \(1987\)](#), pp. 185–200).

In practice, the mean and variance from the integration can be approximated to arbitrary precision by drawing from the distribution for each county and calculating the mean and variance (see [McFadden \(1989\)](#) and [Newey \(2001\)](#) for more complicated applications). The conditional distribution of optimal capital for each county at each time is dependent on other times, and I draw from it using a Gibbs sampler switching between drawing $\ln \eta_c$ conditional on each of the $\ln \epsilon_{ct}$ and all of the $\ln \epsilon_{ct}$ conditional on $\ln \eta_c$. Full details are in appendix B. Calculating moments by drawing from the distribution has the additional advantage that it makes calculating more complicated moments reasonably straightforward. For example, the model includes fixed effects, but it is not practical to estimate the model with a fixed effect for every county since the model no longer has a simple linear form. Absorbing them by subtracting the means of all variables introduces an error correction of the form $\int (e_{ct} - (1/4) \sum_\tau e_{c\tau})^2 f_c(e_{c1870}, \dots, e_{c1900}) d\mathbf{e}$ which is simple to calculate for each county and time using multiple draws from the joint distribution.

To summarize, the full estimation strategy takes the following steps:

1. Estimate equation 1 using panel MLE, assuming that banks will not enter if $C_{ct}^* \leq C^T$, are constrained if $C_{ct} = 50$ and so $C^T < C_{ct}^* \leq 50$ and are unconstrained if $C_{ct} > 50$, and that the unobserved factors that affect the optimal capital are distributed by $\ln \epsilon_{ct} \sim N(0, \sigma^2)$ and $\ln \eta_c \sim N(0, \sigma_\eta^2)$. Find the C^T that maximizes the log likelihood.
2. Using the estimates \hat{C}^T , $\hat{\alpha}$, $\hat{\alpha}_t$, $\hat{\sigma}$, and $\hat{\sigma}_\eta$, draw M values for each constrained county of C_{ct}^* by making M draws from the joint distribution of $\ln \epsilon_{ct}$ and $\ln \eta_c$ conditional on observing the actual capital and population. Each draw is made after N iterations in the Gibbs sampler, where N is a large number.
3. Using M draws, calculate the sample moments for each county. Find the coefficients which make the sample conditional moments as close to zero as possible using standard iterative GMM procedures.

3.5 Using 1902 capital to gain precision

An additional policy change allows me to gain precision in estimating the distribution of optimal capital. In 1900 capital requirements for national banks were halved to \$25,000. Many new banks decided to enter in the next several years at the lower capital requirements, and some existing banks reduced their capital.¹² Where banks decided to enter, and which banks chose to reduce their capital stock is very informative about the optimal capital in a county. For example, a county which gets a bank after 1900 of between \$25,000 and \$50,000 reveals exactly what the optimal capital for that county is. While the reduced capital stock cannot be used to estimate the effects of banking directly—it was not in place and so could not have had an effect on growth before 1900—it does help tighten estimates of the optimal capital. It may have taken several years for new banks to enter at the reduced capital requirement, and so I use the capital of banks in 1902. Under the assumption that the observed capital in 1902 is what banks would have wanted to do in 1900 if the lower capital requirements had been in place, I can use the 1902 capital to tighten the estimates. To do so, I estimate the optimal capital equation allowing for two different capital thresholds: C^T in place for 1870, 1880, and 1890 is the transition capital for entry at \$50,000, and C_{25}^T is the transition capital for entry in 1902 at \$25,000. When I draw from the conditional distribution, it is based on the observed behavior in 1870, 1880, 1890 and 1902. But when calculating excess capital, I use the actual capital in place in 1900. Using 1902 thus allows better estimates of optimal capital. Using 1902 changes which counties are more likely to be close to entry, and so selects a different sample of counties, and so I show estimates both using and not using the capital in 1902.

¹²Due to the application process to the Comptroller, and the requirement to raise capital, few \$25,000 banks had opened between the passing of the Gold Standard Act and the Comptroller's report. Whenever I use 1900 information, the few banks which opened below \$50,000 are excluded since they could not have been open more than a few months and so their effect on growth between 1890 and 1900 would be small.

4 Data

4.1 Sources and construction

This section gives a brief description of the sources and construction of the data, and some descriptive statistics. Appendix C gives additional details. I have created a new data set at the bank level of all national banks in 1870, 1880, 1890, and 1900. The bank level data comes from the national bank accounts collected by the Comptroller of the Currency (who is still the official regulator of national banks), and reported to Congress each year. The accounts of each bank report the town or city in which it was located—which since branching was not allowed was its only place of business. I match these towns and cities with the place names from the Graphical Names Information System maintained by the US Board of Geographic Names, which gives gives a latitude and longitude. Figures 2 through 5 show the results of this placement as banks spread over time and across space. With their location, I can match the banks to the counties (defined by 1890 county boundaries).

The census collects a wide range of demographic and economic information every ten years and reports the aggregates for counties, which are sub-unit of states, almost always with their own local governments. In keeping with the importance of agriculture during the period, the census collected detailed information on farm production, yields, and farm size, as well as some manufacturing production. [Haines and ICPSR \(2010\)](#) collected and entered this information, as well as the aggregate at the state level, and I use the National Historical Graphical Information System ([Minnesota Population Center, 2004](#)) to match the the census data with counties and the location of banks.

4.2 Sample selection

Throughout the analysis, I exclude all counties that had an urban population of 50,000 or more in 1880, which excludes counties with major cities. Counties with with large urban population were

not generally constrained by the minimum capital requirement, so the restriction mainly affects the estimates of optimal capital. Since the activities of banks in counties with large urban populations were different than banks in rural counties, it makes sense to exclude them from the analysis which focuses on the identifiable effects of additional banking in marginal counties.

For most of the analysis, I include all rural counties for which data are available. The Southern banking system was largely destroyed after the war, however, and there were far fewer banks in the South over the entire period. National banks could have had a very different effect in the South, and given the relative poverty of the Southern states, including them might affect the results. Similarly, although states were carved out of the Western territories after the Civil War, they were sparsely populated, and not consistently divided into counties, and so the Census does not provide useful data on some of them until later in the period. For comparison, I also examine results using a restricted sample of the Union (Northern) and border states during the Civil War, which are the Northeastern, Atlantic, and Midwestern states.¹³

4.3 Descriptive statistics

Table 1 shows descriptive statistics for banking and the economic variables over counties and for each decade. The table shows statistics for both all rural counties, and the rural counties in the Union states. Combining farm and manufacturing production from the census, I create a measure of total production per capita. While it does not include services, and so does not directly measure aggregate output, services would have been a small portion of the economy at the time, and the analysis includes fixed effects or first differences, and so the constant absorbs services to the extent to which services are proportional to the rest of economic activity. I do not adjust for deflation or inflation in the table, but instead allow for time effects throughout the analysis. Deflation occurred

¹³I also exclude California, Oregon, and Nevada, which were states before 1870, but might not be comparable since they have incomplete coverage by the census in early decades and are far from the rest of the states. The full list of states in the restricted sample is: Connecticut, Delaware, Illinois, Indiana, Iowa, Kansas, Kentucky, Maine, Maryland, Massachusetts, Michigan, Minnesota, Missouri, Nebraska, New Hampshire, New Jersey, New York, Ohio, Pennsylvania, Rhode Island, Vermont, West Virginia, and Wisconsin.

over much of the period, particularly around the Resumption Act of 1875, and so values in nominal dollars tend to understate growth. The average county population grows substantially over the time period. Following a dip during the 1870s, so does manufacturing and farming production per capita, and in real prices both are likely growing very quickly. Manufacturing production tends to be fairly concentrated, even when I exclude counties with an urban population of more than 50,000 in 1880. So although total manufacturing production is more than half of aggregate production in some years, in an unweighted average over counties it represents less than half of total production since many counties have little manufacturing. The average number of banks increases over time as well, particularly during the 1880s, when the banks per capita more than doubles. Moreover, banks tended to fill in gaps in geographic coverage—the average distance to a bank declines substantially from 1870 to 1890.

5 Results

5.1 Optimal capital

The first step in estimating the effects of banking is to estimate the optimal capital equation 1. The full estimation details are in appendix A. Column 1 in table 2 shows the estimates for the all counties, while column 2 shows the estimates using the capital stock for 1902 after the minimum capital requirement was halved. Columns 3 and 4 restrict the sample to rural counties in border or Union states. There is a great deal of fixed heterogeneity across counties (σ_η), and a much smaller individual county decade heterogeneity (σ_ϵ). The importance of the fixed component suggests that how good a place is for banking is largely fixed, and so entry depends particularly on whether a county has a sufficient population to sustain a bank of at least the minimum size, rather than on whether a county has an idiosyncratic shock. These results reflect the experience of the period of population growth accompanied by entry, as shown in figures 2 - 5, with relatively few counties losing banking (although individual banks might exit). Banks were not willing to enter with \$50,000 unless the population and business climate were enough to make their optimal capital

\$19,454, which suggests that entry with the minimum capital was profitable even when the optimal capital was low. Since banks could place excess capital in reserve city banks, which paid interest on it, the cost of carrying the extra capital may not have been the full cost of capital, but instead the presumably lower difference between the cost of capital and the interest paid on reserves.

With the estimates using the capital in 1902, the entry increases to \$24,359, and banks were willing to enter at the new minimum of \$25,000 if their optimal capital was at least \$14,154. The optimal capital per person also increases. Both of these changes likely come from the many banks still at \$50,000 in 1902. By assumption, banks at \$50,000 in 1902 wanted to be at \$50,000 in 1902, even if they had been forced to be there before, which tends to shift up the average optimal capital. By comparison, a bank with \$50,000 in capital in 1900 could have been anywhere between C^T and \$50,000. Shifting up optimal capital means the optimal capital when forced to enter at \$50,000 must shift up as well. While decades seem long enough that inertia or costs of changing capital should not dominate, the same may not be true of the two and a half years between the reduction of the minimum capital and the Comptroller's report in 1902.

5.2 The causal effects of banking

Table 3 shows the effects of increased banking on the natural log of total production per capita, while table 4 breaks total production into its constituent parts, and examines whether there were changes in the mix of production. All of the estimates include time effects, or are in first differences, which should sweep out overall price changes over time. Since four time periods is a short panel, first differencing allows me to relax the relatively strong assumption that the errors are orthogonal to the amount of banking in a county at all time periods required to make the within group transformation, for the weaker assumption that banking differences are orthogonal to error differences. First differencing comes at an efficiency cost, however, and since the estimates are generally close, I show the first difference results only for the total production.

For both the fixed effects and first differences, I show both the possibly biased estimate from

estimating equation 2 using least squares, the estimates using OLS only on the conditional mean of optimal capital, and the full GMM estimates which account for the full conditional distribution of optimal capital. The coefficient on capital stock is a mix between the coefficients on optimal and excess capital since it represents a mix of the endogenous effect from optimal capital, and the effect of the excess capital from the minimum capital requirement. The coefficients on the excess capital, on the other hand, are a causal effect of banking.

To make the estimates with different lags comparable, and put them in units which are meaningful, I calculate the effect of adding a bank with the minimum size capital of \$50,000 to a county with the average county population in 1880. Using the average county population may tend to underestimate the marginal effect of getting a bank on the marginal county—the best county for banking that does not already have a bank—since the marginal county likely has a lower population than the average county, but it is useful as a standard comparison. For the estimates with lags the marginal value is the total effect from both lags of either the capital stock per capita, or the excess capital per capita. I also report the p-value of the hypothesis that the total effect is zero, taking into account the unobserved C^* .

An additional \$50,000 in capital has a big effect on the production in the marginal county, increasing the total production value per capita by between 9 percent in the regressions in levels with a high level of confidence. Including lags seems to increases the total effect, which suggests that banking may have a continuous growth effect, not just a level effect. Of course, with only four time periods, it is not easy to distinguish between them. While later lags are smaller, particularly in the regressions not using the extra information from 1902, there is no strong evidence for the effect varying over time. The coefficient on the excess capital per capita is substantially larger than either the coefficients on capital stock per capita or optimal capital per capita, which suggests that the simultaneity and omitted variable bias on the endogenous effect tends to underestimate the effect of banking.

To examine how banking affects the components of total productivity, table 4 shows the GMM

estimates for manufacturing value per capita, farm production value per capita, and for the share of manufacturing in total output. For a county with the average population, gaining \$50,000 in capital increases manufacturing production, but the estimates are not statistically significant. The increase in total production seems to be particularly driven by increases farm production value per capita, which increases by seven percent with one lag, and 22 percent including two lags.

The fraction of manufacturing in the total production, shown in the last two columns of table 4, suggests that additional capital slightly decreased the share of manufacturing, particularly in the second decade. A zero or negative result is particularly striking because national banks could not lend on mortgages, and so the only direct way to promote agriculture was through financing trade. These results, of course, hold only for the marginal rural county—national banks may have promoted industrialization in the cities, while facilitating trade in rural areas.

Banks seem to have promoted agriculture largely on the extensive margin, rather than on the intensive margin. Table 5 examines how banking affected the fraction of improved farmland in a county, the yield (in dollars of production per acre), and the Gini coefficient of farm size. Improved farmland is land that has been cleared and is being tilled or is lying fallow, and includes orchards and permanent pastures, and so represents land that is actively used for agriculture. For all rural counties, there is no effect of excess banking on either the yield or the fraction of improved farmland. Limiting the sample to rural Union counties, there does seem to be a modest increase in the farmland, but none in yield (see table 7). Banks seem to have increased inequality of farm size, but only after a decade. The lag suggests that it was not that banks promoted larger farms when a county was first settled, but that they encourage consolidation after settlement.

5.3 Distance to banking

Using geography to identify the effects of banking assumes that distance somehow matters. If someone in a county far from the centers of finance can get a loan just as easily as someone close by, then the presence of a local bank should make little difference. The spread of banks with

population in figures 2 through 5 shows that local banking is important—otherwise all banking could be done at lower cost in one location. But in identifying a single effect of extra capital, the estimate ignores that some counties are much closer to other sources of banking than other counties. For example, a county may have banks in towns all along its border, and so have very good access to banking, even though the county itself may not have a bank. The effect of additional banking capital in such a county may be much smaller than in a county far from any bank. To test how distance matters, I construct a measure of distance to banking for each county in each year of the analysis. Since I know the location of each bank, I take the mean distance by area within a county to the nearest bank, which may be within the county itself.¹⁴ Since the effect of the distance to banking is likely to be highly non-linear, I interact the inverse of the distance to banking with the excess capital, and so examine whether having additional capital makes more of a difference in counties which are far away from other banks.

Table 6 shows the results of including an interaction with the inverse of distance to the nearest bank. Since the sparsely populated and poorly banked West may be driving much of the relationship, I estimate the effect of restricting the sample to the Northern and Midwestern states as well as including all rural counties. The interaction effect is negative and about the same size for both samples, suggesting that the better a county is covered by banks, the smaller difference extra capital makes. In the Union states the effect of distance is much more significant, although the point estimates are similar.

The estimates for both samples suggest that distance mattered a great deal. the benefits of opening a bank in an area where banks are on average only 11-20 kilometers away is approximately zero. Opening a bank in a county where banks are the average distance away in 1880 has an effect

¹⁴More formally, the mean distance by area to the nearest bank in a county is $(1/A) \int \int_{(x,y) \in C} \min\{D((x,y), (x_b, y_b))\} dy dx$ where A is county area, $D(\cdot, cdot)$ is the appropriate distance function for the projection, and the min is taken over all banks with locations (x_b, y_b) . In practice, I calculate an approximation to the integral using a two dimensional Riemann sum. First I create a raster with kilometer square regions (smaller sizes over the entire US sometimes failed due to the size of the resulting data set), each of which contains the shortest distance to a bank for each year from the center of the square. Using the 1890 county shapefile, I average over distances from the raster regions which fall within each 1890 county, which gives a county level mean distance to the nearest bank.

of between 13 and 16 percent.

5.4 Robustness

To check the robustness of the results, I restrict the sample in several ways and examine the effect of including the information from 1902. Table 7 shows the marginal effect and its p-value for all of the dependent variables for a sample of all rural counties and restricting the sample to union rural counties. Table 8 performs the MLE and calculates the conditional moments not including the change in rules to allow banks to open with \$25,000 in capital in 1902.

While most of the analysis includes all rural counties, it is possible that banking in the sparsely populated West and poorly banked South had a different effects than in the North and Midwest. The last three columns in table 7 restrict the sample to rural counties in Northern and Midwestern states, since these counties are the most comparable. Note that is is not necessarily a problem that the rest of the country had few banks: although counties in the West did not get many banks, they also had very low populations, and so should have had few banks. Similarly, the regressions interpret the few banks in the South as meaning that the South was not a good place for banking. Excluding the South and West does not change the results much, although including lags produces more varied but statistically insignificant results. The fraction of the county that is improved farm land increases in the rural Union counties, while for all the coefficient is small and insignificant.

The analysis has focused on using the largest sample with the most information about how much capital a county would have had with the minimum capital requirement, so table 8 examines the sensitivity of the results to these choices. Using the capital stock in 1902 as the capital counties would have had if the minimum capital had been lower in 1900 requires assuming that the banking situation in 1902 is identical to 1900 except for the variation in capital requirements, and that banks have had a chance to fully adjust by 1902. If those assumptions are met, using the changes to 1902 means the estimates of optimal capital are using more information. The estimates of the effect on total production, farm production and manufacturing production are modestly lower when

including 1902 than when not, but the results are generally similar.

Although most of the exposition has focused on the transition between no banks and a bank with \$50,000 in capital, the analysis deals with the intermediate situation where there are other banks in a county, possibly with more than the minimum capital. The presence of a bank with the minimum capital, while other banks are not at the minimum, may suggest that at least one bank is constrained. Alternatively, the total banking in the county may be entirely endogenous, and the existence of a bank with the minimum capital comes from competition or agreement among banks, rather than capital constraints. The last two columns in table 8 therefore remove the relatively few counties that have banks with \$50,000 and banks with more capital. The results are not driven by these counties.

5.5 The very long-term effects of national banking

This section briefly examines the long-term effects of national banks during the Gilded Age. While over the short term—where “short term” here is several decades—the differences in banking cause differences in economic development, over the long term one might expect any initial advantage to disappear, or even turn into a disadvantage as impatient consumers adjust to the availability of credit ([Fulford, 2011](#)). For example, over the long term mobile capital should seek the best marginal return, and so initial differences in financial development may have little long-term effect. Yet a growing literature points to the longevity of institutions and initial advantages. For example, [Pascali \(2011\)](#) finds that the presence of Jewish communities in Italy, who could lend in the middle ages, is a significant predictor of financial development in modern Italy. [Guiso, Sapienza, and Zingales \(2004\)](#) show that the structure of the banking market in 1936 still affected the supply of credit 50 years later, suggesting highly persistent effects even in modern markets.

Does banking increase income not just within decades but a century later? Table 9 shows results of estimating the effect of banking in 1900 and 1870 on income per capita in 1970. I divide up the actual banking capital in a county into optimal capital and excess capital as in the previous

analysis, and estimate only for counties where the conditional mean of optimal capital is within \$15,000 of the entry point. Since the data available in 1970 are very different from 1870-1900, I no longer use fixed effects, but instead include the total production value per capita calculated from the 1900 and 1870 census directly. In addition, I include state indicators, to make sure any estimated effect is not being driven by regional differences. To maintain comparability with the previous results, I calculate the effect of additional capital using the same value: an increase in capital per person equivalent to adding \$50,000 in capital to a county with the average population in 1880.

The results suggest that adding a bank of the minimum size in 1870 is associated with an increase in income per capita of 2% in 1970 for rural counties in the Northeast and Midwest, but not in 1900 or for all rural counties. These results are stable excluding either production per capita or state effects or both. Although these results are somewhat speculative, one way to interpret them is that additional banking during a key period of population movement and investment had long term effects since small advantages became multiplied. Since the barriers to banking were reduced substantially after 1900 as state banks spread and national banks could enter with less capital, having a capital advantage in 1900 might offer only short term advantages. Few counties had banks outside of the Northeast and Midwest in 1870, so while they may have suffered from not having banking in the short-term, lack of banking may have had few impacts in the long-term.

6 Conclusion

Although much has been written about the financial institutions during the period, this paper is the first to estimate the causal effects of banking during the Gilded Age. Banks, while certainly following growth, also contributed substantially to it: for the marginal county close to the line between getting a bank and not, the presence of a bank was worth an extra 9 percent in total production per capita. Banks seem to have promoted both farming, and to a less well estimated extent, manufacturing, but in marginal counties tilted the production mix to agriculture. The distance to a

bank is very important: banks coming into areas with few banks have a larger effect.

One way to read these results is that bad banking regulation can be very costly. An initially somewhat arbitrary decision by an administrator not to allow branching became entrenched with costly long-term consequences. The estimates are possible because of this rule, but also represent its high local cost. The no-branching rule promoted smaller unit banks, while denying many places that could have supported a branch of larger, possibly better diversified, bank the benefits of banking. In the long term the inability to branch seems to have reduced competition and harmed growth, as [Jayaratne and Strahan \(1996\)](#) find when examining the consequences of relaxing limits on inter-state branching a century later. Such continuing restrictions on capital mobility may suggest why counties with more capital in 1870 still had higher incomes in 1970.

National banks both issued loans, typically of short duration and often to fund goods in transit, and national bank notes, which as currency facilitated the exchange of goods. It is not clear whether the effects national banks had on production came from the increase in the local money supply or the local credit supply. Indeed, the two may not be separable: Schumpeter argues “that all forms of credit, from the bank-note to book-credits, are essentially the same thing” ([Schumpeter, 1934](#), p. 99), and he might be right. National bank notes traveled widely, however, and were redeemed only infrequently ([Selgin and White, 1994](#)), and so their effects on local conditions may have been small after the initial offering. In either case, since banks were generally not making loans to expand businesses or farms directly, and could not take mortgages as collateral, banks helped grease the wheels of commerce, rather than provided the capital to create new enterprises.

What lessons does this episode have for modern development? First, it seems that in the face of expensive and time consuming transport, commercial banking can be very important. Since developing countries often have limited infrastructure, particularly in rural areas, the major constraint may be less the ability to expand, which can often be done incrementally, but a liquidity constraint from the timing of payments. Farmers may have income only once a year, but have expenses all year long, merchants may have to keep a stock of goods which can be acquired only

infrequently, or at lower prices in bulk. Providing working capital and funding goods in transit are not secondary functions of banks during development, historically they are the key functions of banks during development. Second, distance to banking matters, at least when transportation or communication is an issue. While national capital markets may be important as a way to acquire capital, local financial institutions matter for local growth. Third, banking can encourage production, but it is also likely to encourage production at the most efficient scale, and by the most efficient producers, which may tend to increase inequality. The recent resurgence of interest in how financial institutions, whether banks or microfinance institutions, can help development has made it even more important to understand the mechanisms of how financial institutions actually helped development.

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A Estimating optimal capital using maximum likelihood

Because of minimum capital requirements, some counties have more capital than they would get if banks could freely choose their capital. The observed amount of capital is censored at \$50,000 and zero, however. Conditional on observing a bank with \$50,000 in capital, the optimal capital cannot be too much less than \$50,000, otherwise a bank would not want to enter. Similarly, if there is no bank in a county, the optimal capital cannot be very close to \$50,000, since otherwise it would be profitable to enter.

The banking model means that transition capital C^T is the point where banks choose to enter at 50, for an individual bank in a county acting as a monopolist. Multiple banks divide up the county among themselves and act as monopolists within their portion. Then bank i with capital C_{ct}^i gets $p^i = C_{ct}^i/C_{ct}$ of the population. A bank with $C_{ct}^i = 50$ must have $C^T < \alpha\alpha_t p^i P_{ct}\eta_c \epsilon_{ct} \leq 50$ and an unconstrained bank has $C_{ct}^i = \alpha\alpha_t p^i P_{ct}\eta_c \epsilon_{ct}$. Adding up all banks gives $C_{ct} - C^T B_{ct}^{50} < \alpha\alpha_t P_{ct}\eta_c \epsilon_{ct} \leq C_{ct}$ where B_{ct}^{50} is the number of banks with \$50,000 in capital stock.

Putting it all together, if a county has capital stock C_{ct} in thousands, which may be zero, B_{ct} banks, and B_{ct}^{50} banks with 50 capital, and the dividing line between entering and not is C^T , then:

- (1) Optimal capital is observed capital if the bank is unconstrained: $C_{ct} = \alpha\alpha_t P_{ct}\eta_c \epsilon_{ct}$ if $B_{ct} > 0$ and $B_{ct}^{50} = 0$;
- (2) Optimal capital must be between C^T and 50 if the bank is constrained but enters: $C_{ct} - (50 - C^T)B_{ct}^{50} < \alpha\alpha_t P_{ct}\eta_c \epsilon_{ct} \leq C_{ct}$ if $B_{ct}^{50} > 0$;
- (3) Optimal capital must be less than C^T if no bank enters: $\alpha\alpha_t P_{ct}\eta_c \epsilon_{ct} \leq C^T$ if $B_{ct} = 0$ and so $C_{ct} = 0$.

For any given profitability cutoff C^T , suppose that the county businesses condition shifters are distributed lognormally so that: $\ln \epsilon_{ct} \sim N(0, \sigma^2)$ and $\ln \eta_c \sim N(0, \sigma_\eta^2)$. Define D_{ct} as 1 if there are any banks in the county and 0 otherwise, D_{ct}^{50} as 1 if there is a bank with \$50,000 capital and 0 otherwise, and B_{ct}^{50} as the number of banks with \$50,000 in capital. Then the density of observing

capital C_{ct} in county c in year t , given η_c , α , α_t , and P_{ct} is:

$$\begin{aligned}
f(C_{ct}|P_{ct}, \alpha, \alpha_t, \eta_c) = & \left(\Phi \left[\ln(C^T/P_{ct}) - \alpha - \alpha_t - \eta_c \right] / \sigma \right)^{(1-D_{ct})} + \\
& \left(\left(\phi \left[\left(\ln \frac{C_{ct}}{P_{ct}} - \alpha - \alpha_t - \eta_c \right) / \sigma \right] / \sigma \right)^{(1-D_{ct}^{50})} + \right. \\
& \left. \left(\Phi \left[\left(\ln \frac{C_{ct}}{P_{ct}} - \alpha - \alpha_t - \eta_c \right) / \sigma \right] - \Phi \left[\left(\ln \frac{C_{ct} - (50 - C^T)B_{ct}^{50}}{P_{ct}} - \alpha - \alpha_t - \eta_c \right) / \sigma \right] \right)^{D_{ct}^{50}} \right)^{D_{ct}}. \tag{5}
\end{aligned}$$

The conditional density of observing \mathbf{C}_c , the vector for capital in county c for all years, conditional on η_c is the product of all of the densities for each year:

$$f(\mathbf{C}_c|\mathbf{P}_c, \alpha, \alpha_{1880}, \alpha_{1890}, \alpha_{1890}, \eta_c) = \prod_{t=1870}^{1900} f(C_{ct}|P_{ct}, \alpha, \alpha_t, \eta_c).$$

Finally, the density unconditional on η_c comes from integrating out η_c :

$$f(\mathbf{C}_c|\mathbf{P}_c, \alpha, \alpha_{1880}, \alpha_{1890}, \alpha_{1890}, \sigma, \sigma_\eta) = \int f(\mathbf{C}_c|\mathbf{P}_c, \alpha, \alpha, \eta_c) \phi[\eta_c/\sigma_\eta]/\sigma_\eta d\eta_c. \tag{6}$$

Maximum likelihood then finds the set of parameters which maximize the sum over all counties of the log of equation 6. The difficulty is that for each county the observations in each time period are dependent on each other through η_c , and so standard maximization routines which rely on each piece of the log likelihood being separate do not apply. The solution is to actually perform the integration over η_c to get the unconditional density, or at least a numerical approximation of it, and then maximize over the county level likelihood. This approach is implemented in Stata using a Gauss-Hermite quadrature to approximate the integration in equation 6 by xtintreg ([StataCorp, 2009](#), pp. 180–188).

B Drawing estimates of excess capital

Given estimates of the parameters which determine the optimal capital, and realizations of ϵ_{ct} and η_{ct} , a county with a 50 bank has $50 - \alpha\alpha_t P_{ct}\eta_{ct}\epsilon_{ct}$ too much capital. Counties without banks have $\alpha\alpha_t P_{ct}\eta_{ct}\epsilon_{ct}$ too little capital. I cannot observe ϵ_{ct} and η_{ct} , but I do know how they are distributed, conditional on observing a particular value of C_{ct} and the number of banks in a county. Reversing the constraints for the maximum likelihood:

1. $\ln(C_{ct}/P_{ct}) - \hat{\alpha} - \hat{\alpha}_t = \eta_c + \epsilon_{ct}$ if $B_{ct} > 0$ and $B_{ct}^{50} = 0$,
2. $\ln(C_{ct} - (50 - \hat{C}^T B_{ct}^{50})/P_{ct}) - \hat{\alpha} - \hat{\alpha}_t < \eta_c + \epsilon_{ct} \leq \ln(C_{ct}/P_{ct}) - \hat{\alpha} - \hat{\alpha}_t$ if $B_{ct}^{50} > 0$, and
3. $\eta_c + \epsilon_{ct} \leq \ln(\hat{C}^T/P_{ct}) - \hat{\alpha} - \hat{\alpha}_t$ if $B_{ct} = 0$ and so $C_{ct} = 0$.

Call lb_{ct} the lower bound for $\eta_c + \epsilon_{ct}$ for county c at time t which can be negative infinity, and similarly for ub_{ct} . Then the joint density is:

$$f(\epsilon_{c1870}, \dots, \epsilon_{c1870}, \eta_c | lb_{ct} < \eta_c + \epsilon_{ct} \leq ub_{ct})$$

is a multivariate truncated normal. To draw from such distribution I use a Gibbs Sampler ([Lancaster, 2004](#), pp. 207–221) which converges to the multivariate truncated normal. Since the distribution of each ϵ_{ct} is univariate truncated normal conditional on η_c , and similarly η_c is univariate truncated normal conditional on all ϵ_{ct} , the Gibbs sampler has the steps:

1. (a) Choose η_c^0 for all counties. (b) Draw ϵ_{ct}^0 from a truncated normal with bounds $lb_{ct} - \eta_c^0$ and $ub_{ct} - \eta_c^0$ and variance σ^2 .¹⁵
2. (a) Given $\{\epsilon_{ct}^{i-1}\}$, draw η_c^i from a univariate truncated normal with lower bound $\max_t\{lb_{ct}\} - \epsilon_{ct}^{i-1}$ and upper bound $\min_t\{ub_{ct} - \epsilon_{ct}^{i-1}\}$, and variance σ_η^2 . (b) Given η_c^i draw each element of $\{\epsilon_{ct}^i\}$ from its univariate truncated normal.

¹⁵Drawing from a truncated normal using standard psuedo-random number generators takes a bit of work. If u_i is a draw from a uniform [0,1] distribution, then $x_i = F^{-1}(u_i) = (1/\sigma)\Phi^{-1}[\Phi[a/\sigma] + u_i(\Phi[b/\sigma] - \Phi[a/\sigma])]$ is a draw from a truncated normal with lower bound a , upper bound b , and variance σ^2 .

3. Repeat step two N times, where N is large, and take the Nth draw of $\{\epsilon_{ct}^N\}$ and η_c^N .

I choose η_c^0 by letting $(\eta_c^0)_t = lb_{ct} + (ub_{ct} - lb_{ct})/2$ if county c in time t is a type two county and $(\eta_c^0)_t = (\ln(C^T/P_{ct}) - \hat{\alpha} - \hat{\alpha}_t)$ if the county is a type three county, and then averaging the $(\eta_c^0)_t$'s. For multiple draws in the imputations, I continue the sampler and use every Nth draw as an approximately independent draw from the multivariate truncated normal, thus letting the sampler run MN times.

C Data

C.1 County level data for 1860-1920

Counties or their equivalents are geographic subdivisions of states. For most states the county has a governmental role as a middle level of government between the state and the local governments of cities, towns, or boroughs. The US Census collects data at the county level, and for consistency I assign independent cities (such as St. Louis, Missouri) to their own counties, or to the county which contains them. The creation of several new states from territories, and their division into counties between 1870 and 1900 makes consistent geographical designations very important. In established states, counties shift boundaries and split occasionally. To create a data set which consistently refers to the same geographic area over time, I use graphically information from the National Historical Graphical Information System ([Minnesota Population Center, 2004](#)), which provides county boundaries for each decade. From the county shape file for each decade, I calculate the union with the 1890 counties, which gives the 1890 county that all counties or parts of counties belong to in a given year. I also calculate the area of each county fragment. I discard all fragments with less than 1 mile square since these fragments represent small shifts in the county polygons rather than changes in county definition. The NHGIS also provides county level census information linked to the county shape files based on the census data from [Haines and ICPSR \(2010\)](#). Since some 1890 counties are composed of pieces from multiple counties, I allocate census economic

and demographic information by county for each non-1890 census year using the area of the county fragment. So a county whose boundary shifts to include some of another county gets additional population from the other county in proportion to the area absorbed. I create 1890 county level means or distribution variables by taking the area weighted average of the county fragments that compose the 1890 county. This procedure is exact if population or other aggregates are uniformly distributed over the surface area of changing counties. While such changes are important for the sparsely settled Western states, for the established states on which I conduct much of the analysis, county boundaries are very stable.

C.2 National Bank Accounts

Each year the Comptroller of the Currency reported the accounts of the all National Banks.¹⁶ For each bank in 1870, 1880, 1890, and 1900, I have its number (matched to 1871 for 1870), its place of business, and its loans and discounts, capital stock, and total liabilities.¹⁷ The Comptroller's report includes other interesting information for banks which we have not entered, including notes outstanding, deposits, and reserves. I match the place of business to the Graphical Names Information System maintained by the US Board of Geographic Names [U.S. Board of Geographic Names \(2010\)](#). While most places with a National Bank still exist, some have merged with other towns or cities, in which case I match with the modern city. The match gives the latitude and longitude of the bank using the North American Datum of 1983. I match the bank location with the 1890 counties above, which gives the number of banks and national bank aggregates in the geographic area of an 1890 county for each decade from 1870 to 1900.

¹⁶Available in pdf from the St Louis Federal Reserve, <http://fraser.stlouisfed.org/publications/comp/>, accessed 7 July 2010.

¹⁷My excellent RA Shahed Kahn entered these series by hand, with the help of some optical character recognition to speed the process of entry. I checked the 1870 and 1880 accounts, and another RA, Mashfiqur Khan, checked the later years.

Table 1: County descriptive statistics

Year	All rural counties				Rural Union counties			
	1870	1880	1890	1900	1870	1880	1890	1900
County population	11660 (13398)	14703 (14889)	17955 (17920)	21330 (28428)	16514 (16036)	19848 (17074)	23381 (20295)	27105 (36170)
Total production	104.0 (82.5)	80.1 (64.3)	96.9 (77.2)	155.0 (312.7)	122.6 (76.2)	103.1 (67.1)	120.0 (83.8)	162.3 (97.5)
Manufacturing	38.5 (66.2)	32.3 (51.8)	44.7 (71.9)	60.4 (97.2)	52.1 (67.4)	49.2 (62.9)	65.7 (89.0)	82.6 (106.7)
Farm production	65.5 (48.4)	47.8 (37.8)	51.0 (40.5)	94.5 (304.8)	70.4 (39.8)	53.9 (32.8)	53.7 (28.4)	79.7 (47.2)
Fraction manuf. in total value	0.32 (0.28)	0.31 (0.25)	0.35 (0.28)	0.36 (0.27)	0.35 (0.24)	0.37 (0.26)	0.42 (0.28)	0.42 (0.28)
Gini farm size	0.431 (0.171)	0.392 (0.169)	0.397 (0.153)	0.456 (0.105)	0.431 (0.139)	0.374 (0.122)	0.364 (0.119)	0.417 (0.063)
Fraction improved farmland	0.36 (0.24)	0.36 (0.21)	0.53 (0.23)	0.50 (0.25)	0.46 (0.25)	0.45 (0.20)	0.65 (0.20)	0.66 (0.21)
Farm yield	15.34 (9.22)	11.00 (8.10)	7.61 (5.56)	10.66 (7.58)	14.07 (6.53)	10.02 (6.55)	6.84 (3.80)	9.19 (4.64)
Number of banks	0.41 (1.28)	0.57 (1.46)	1.06 (1.93)	1.09 (1.99)	0.82 (1.75)	1.08 (1.95)	1.68 (2.37)	1.78 (2.50)
Distance to closest bank (km)	153.2 (176.13)	89.4 (89.54)	42.5 (37.37)	40.1 (35.51)	70.3 (82.31)	52.4 (62.20)	25.9 (20.23)	24.7 (20.60)
Banks per 1000 capita	0.015 (0.042)	0.021 (0.050)	0.051 (0.094)	0.042 (0.073)	0.029 (0.053)	0.036 (0.059)	0.063 (0.080)	0.057 (0.064)
Capital stock	1.96 (6.23)	2.30 (6.11)	4.46 (8.18)	3.46 (6.14)	3.71 (8.22)	4.04 (7.74)	5.84 (8.19)	4.96 (6.52)
Loans and discounts	2.74 (8.45)	3.93 (10.75)	10.41 (20.40)	10.52 (19.95)	5.20 (11.08)	6.63 (12.47)	13.68 (20.22)	14.84 (20.10)
Counties	2665	2745	2743	2745	1261	1301	1301	1301

Notes: Standard deviations on in parentheses. The average is taken over counties and is unweighted. Values are in dollars from the census or national bank accounts and are not corrected for inflation/deflation (there was significant deflation between 1870-1880 as the US went back to a full gold backing of its currency). Rural Union counties are counties from Union or border states with an urban population of fewer than 50,000. Yield is the farm production value divided by the area improved farmland in a county, excluding extreme values driven low areas of improved farmland).

Table 2: Log-likelihood estimates of optimal capital

	All Rural	All Rural with 1902	Rural Union	Rural Union with 1902
$\ln \alpha$	-7.513*** (0.0514)	-7.252*** (0.0481)	-7.024*** (0.0596)	-6.817*** (0.0556)
$\ln \alpha_{1880}$	0.119*** (0.0341)	0.0961*** (0.0343)	0.126*** (0.0404)	0.106*** (0.0405)
$\ln \alpha_{1890}$	0.783*** (0.0329)	0.700*** (0.0329)	0.670*** (0.0393)	0.601*** (0.0393)
$\ln \alpha_{1900}$	0.604*** (0.0329)	0.764*** (0.0328)	0.552*** (0.0394)	0.720*** (0.0390)
σ_η	1.735*** (0.0393)	1.642*** (0.0352)	1.540*** (0.0452)	1.427*** (0.0405)
σ_ϵ	0.680*** (0.0101)	0.688*** (0.0100)	0.664*** (0.0123)	0.672*** (0.0122)
C^T	19.454	24.359	17.977	22.926
C_{25}^T		14.154		13.221
Observations	10804	10804	4998	4998
Counties	2745	2745	1262	1262

Notes: Estimates of the optimal capital equation $C_{ct}^* = \alpha \alpha_t P_{ct} \eta_c \epsilon_{ct}$ based on the panel using C^T as the dividing line for entering with \$50,000 in capital, and C_{25}^T for entry with \$25,000 in the columns using 1902. The full estimation details are in appendix A. The first column uses all rural counties as described in section 4.2, the second column uses all rural counties and the capital in 1902 for 1900, the third and fourth columns restrict the sample to Union and border states.

Table 3: Banks and total production per capita

		log Total production value per capita										
		Level using 1902					First difference using 1902		Level not using 1902			
		OLS	OLS	OLS	GMM	GMM	OLS	OLS	GMM	GMM	GMM	GMM
Capital Stock p.c.		15.52*** (3.268)					8.680*** (2.375)					
Excess capital p.c.		38.34*** (11.36)	32.48*** (8.218)	27.31*** (7.136)	27.89*** (4.610)		16.14 (10.22)	22.20** (8.685)	40.38*** (10.83)	37.92*** (6.363)		
L—			34.59*** (7.095)		28.12*** (4.083)					30.36*** (5.059)		
Optimal capital p.c.		7.361** (3.603)	6.287 (4.112)	11.43*** (3.296)	10.24*** (3.622)		6.356* (3.560)	4.465 (3.374)	13.13*** (2.639)	11.12*** (3.119)		
L—			5.055 (3.026)		8.389*** (2.909)					6.964** (2.997)		
Observations		10462	4090	3074	4090	3015	3015	3015	3015	8156	5990	
R-squared		0.190	0.219	0.427			0.119	0.119				
County FE		YES	YES	YES	YES	YES	NO	NO	NO	YES	YES	
Year FE		YES	YES	YES	YES	YES	YES	YES	YES	YES	YES	
Counties		2743	1048	1048	1048	1048	1042	1042	1042	2083	2083	
Marginal effect		0.0528	0.130	0.228	0.0929	0.190	0.0295	0.0549	0.0755	0.137	0.232	
p-val marginal ≠ 0		1.95e-05	0.00151	5.14e-07	0.000130	0	0.000659	0.121	0.0106	0.000193	0	

Notes: Capital stock is measured in 1000's of dollars. L— is the first lag of the variable. The OLS estimates with capital stock use all rural counties for which information is available. The GMM and OLS estimates with excess and optimal capital restrict the sample to counties with a conditional mean of optimal capital within 15 of the dividing line which maximizes the log-likelihood in table 2. See appendix for estimation details. The marginal effect is the the sum of the all lagged effects calculated for a county with the average 1880 population gaining \$50,000 in capital. Including 1902 means that the conditional moments for optimal capital include information from banks entering with \$25,000 in 1902.

Table 4: Banks and the mix of production

	log Manufacturing value per capita		log Farm production value per capita		Frac. Manufacturing in total production	
Excess capital p.c.	19.23*	17.99	23.24***	31.91***	-2.146	-3.307*
	(10.14)	(11.74)	(7.120)	(6.240)	(1.717)	(1.847)
L—		1.887		33.62***		-6.546***
		(9.919)		(5.013)		(1.515)
Optimal capital p.c.	6.526	5.555	6.969**	5.380*	-0.505	-0.317
	(5.459)	(6.288)	(2.760)	(2.879)	(0.916)	(1.007)
L—		4.321		5.031*		-0.737
		(3.464)		(2.696)		(0.827)
Observations	3963	2872	4081	3030	4090	3015
R-squared						
County FE	YES	YES	YES	YES	YES	YES
Year FE	YES	YES	YES	YES	YES	YES
Counties	1045	1045	1048	1048	1048	1048
Marginal effect	0.0654	0.0676	0.0790	0.223	-0.00730	-0.0335
p-val marg. ≠ 0	0.0580	0.294	0.00110	0	0.211	0.000902

Notes: Capital stock is measured in 1000's of dollars. L— is the first lag of the variable. All estimates use the conditional moments including banks in 1902 and all rural counties, restricting to counties with a conditional mean of optimal capital within 15 of the dividing line which maximizes the log-likelihood in table 2. See appendix for estimation details. The marginal effect is the the sum of the all lagged effects calculated for a county with the average 1880 population gaining \$50,000 in capital.

Table 5: Banks and farm production

	Fraction improved farm land		Farm yield		Gini of farm size	
Excess capital p.c.	3.532 (2.601)	1.424 (2.290)	-46.81 (88.21)	-43.13 (110.6)	1.394** (0.559)	2.792*** (0.748)
L—		-2.877* (1.676)		-84.97 (69.95)		9.316*** (2.510)
Opt. capital p.c.	2.755*** (0.988)	2.134*** (0.769)	-15.47 (34.59)	13.92 (38.10)	0.135 (0.479)	0.193 (0.423)
L—		-1.624 (1.125)		56.41* (33.99)		2.941*** (0.758)
Observations	4150	3097	4063	3014	4150	3097
R-squared						
County FE	YES	YES	YES	YES	YES	YES
Year FE	YES	YES	YES	YES	YES	YES
Counties	1048	1048	1048	1047	1048	1048
Marginal effect	0.0120	-0.00494	-0.159	-0.436	0.00474	0.0412
p-val marg. ≠ 0	0.174	0.693	0.596	0.429	0.0127	3.68e-05

Notes: Capital stock is measured in 1000's of dollars. L— is the first lag of the variable. All estimates use the conditional moments including banks in 1902 and all rural counties, restricting to counties with a conditional mean of optimal capital within 15 of the dividing line which maximizes the log-likelihood in table 2. See appendix for estimation details. The marginal effect is the the sum of the all lagged effects calculated for a county with the average 1880 population gaining \$50,000 in capital.

Table 6: Distance to banking

	All rural counties			Union rural counties		
	log Total prod. per cap.	log Manuf prod. per cap.	log Farm prod. per cap.	log Total prod. per cap.	log Manuf prod. per cap.	log Farm prod. per cap.
Excess capital p.c.	49.38*** (17.16)	24.51 (18.49)	38.86*** (13.35)	83.12*** (7.397)	47.95*** (17.87)	86.47*** (13.25)
Optimal capital p.c.	15.41* (7.928)	9.658 (8.731)	11.07 (7.206)	12.64** (5.190)	20.54*** (7.802)	10.50 (7.681)
Excess capital p.c. × 1/(Dist. to bank)	-1,027* (542.1)	-208.7 (414.6)	-631.0 (402.2)	-939.0*** (269.1)	-315.0 (280.2)	-827.9*** (320.8)
Optimal capital p.c. × 1/(Dist. to bank)	-36.14 (83.35)	-44.98 (124.6)	-49.03 (81.61)	-52.64 (59.07)	-159.1 (114.6)	-78.32 (73.87)
Observations	4090	3963	4081	2946	2884	2947
County FE	YES	YES	YES	YES	YES	YES
Year FE	YES	YES	YES	YES	YES	YES
Counties	1048	1045	1048	751	750	751
Marginal effect	0.129	0.0754	0.108	0.164	0.106	0.178
p-val marginal ≠ 0	0.00200	0.126	0.00201	0	0.00195	0
Min. dist. for pos. effect	20.80	8.49	16.24	11.30	6.57	9.57

Notes: The distance is the average over the county area distance in kilometers to the nearest bank. The marginal effect for the distance is the sum of the main effect of \$50,000 in excess capital at the mean county population in 1880, and the interaction effect of \$50,000 in excess capital times the mean (over counties) distance to the nearest bank in 1880. The last columns includes only the rural Union counties used. The first three columns include all counties for which sufficient data exists, the last restricts the sample to include only rural Union counties. The conditional expectation of optimal capital uses the capital in 1902. The window size is 15. Errors are clustered at the state level.

Table 7: Sample variations

Dependent Variable	All rural counties using 1902			Union rural counties using 1902		
	1 lag	2 lags	Difference	1 lag	2 lags	Difference
log Total production	0.093 [0.000]	0.190 [0.000]	0.076 [0.011]	0.118 [0.000]	-0.321 [0.419]	0.117 [0.002]
per capita						
log Manufacturing prod.	0.065 [0.058]	0.068 [0.294]	0.089 [0.044]	0.081 [0.003]	0.097 [0.089]	0.118 [0.000]
per capita						
log Farm production	0.079 [0.001]	0.223 [0.000]	0.089 [0.004]	0.135 [0.000]	-0.346 [0.199]	0.163 [0.000]
per capita						
Fraction manufacturing	-0.007 [0.211]	-0.034 [0.001]	-0.002 [0.722]	-0.002 [0.598]	0.024 [0.456]	-0.012 [0.063]
in total production						
Gini of farm size	0.005 [0.013]	0.041 [0.000]	0.003 [0.503]	0.006 [0.085]	0.394 [0.880]	0.008 [0.223]
Fraction improved	0.012 [0.174]	-0.005 [0.693]	0.008 [0.119]	0.046 [0.000]	0.707 [0.886]	0.024 [0.001]
farm land						
Farm yield	-0.159 [0.596]	-0.436 [0.429]	1.165 [0.012]	0.200 [0.720]	-1.424 [0.325]	1.983 [0.007]
Observations	4090	3015	3015	2946	2193	2193
Counties	1048	1048	1042	751	751	749

Notes: p-values in brackets, testing the combined effect of both lags of excess capital stock per capita. All regressions include county fixed effects, time effects, and cluster at the state level. Observations and counties are from the estimates of log total production. Only counties with at least one conditional mean within 15 of the dividing line are included.

Table 8: Robustness checks

Dependent Variable	All rural counties not using 1902		Rural union not using 1902		Counties with no more than one bank	
	1 lag	2 lags	1 lag	2 lags	1 lag	2 lags
log Total production per capita	0.137 [0.000]	0.232 [0.000]	0.101 [0.000]	0.123 [0.000]	0.085 [0.000]	0.072 [0.090]
log Manufacturing prod. per capita	0.076 [0.077]	0.016 [0.821]	0.084 [0.000]	0.107 [0.019]	0.063 [0.063]	0.127 [0.345]
log Farm production per capita	0.113 [0.000]	0.258 [0.000]	0.098 [0.000]	0.149 [0.000]	0.062 [0.007]	0.145 [0.000]
Fraction manufacturing in total production	-0.011 [0.031]	-0.038 [0.000]	0.002 [0.509]	-0.008 [0.204]	-0.005 [0.392]	-0.015 [0.169]
Gini of farm size	0.005 [0.043]	0.042 [0.000]	0.001 [0.592]	0.033 [0.000]	0.004 [0.030]	0.035 [0.000]
Fraction improved farm land	0.022 [0.033]	0.009 [0.565]	0.039 [0.000]	0.031 [0.000]	0.008 [0.324]	-0.007 [0.487]
Farm yield	-0.287 [0.351]	-0.383 [0.430]	0.261 [0.624]	0.655 [0.195]	-0.119 [0.697]	-0.269 [0.433]
Observations	8156	5990	3902	2866	3406	2502
Counties	2083	2083	1012	1012	877	877

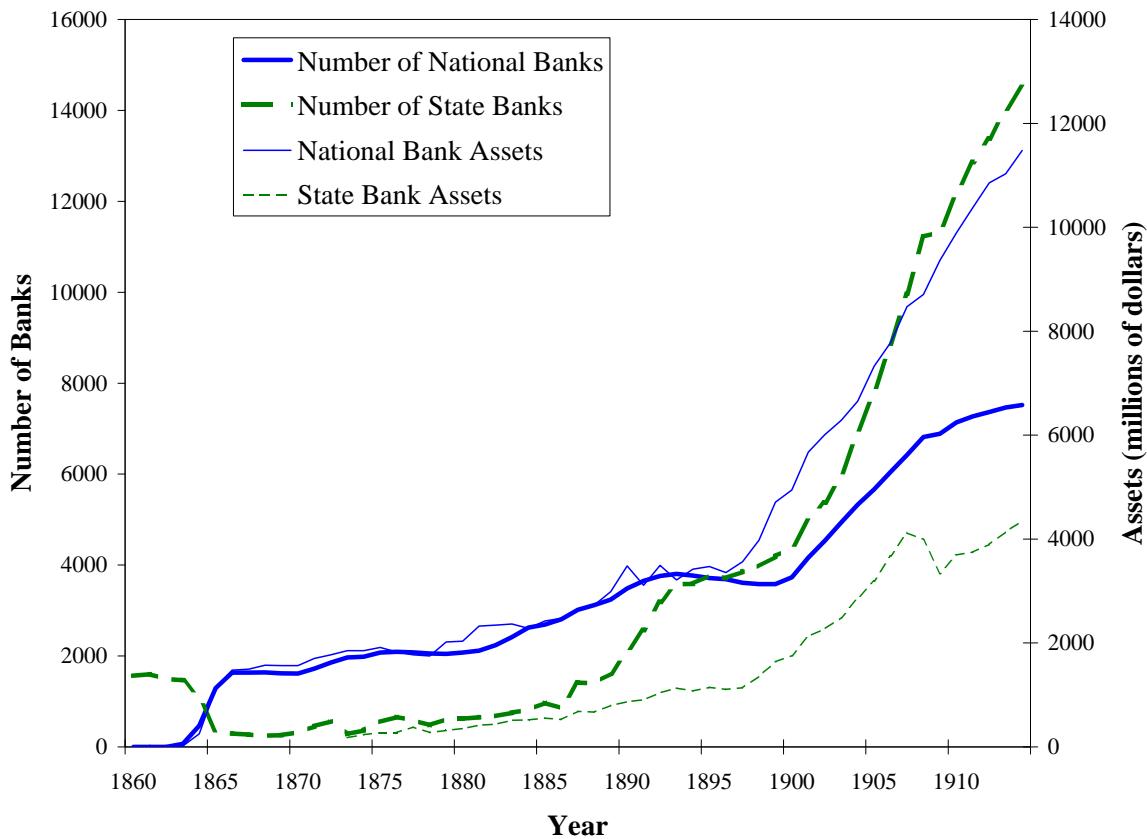
Notes: p-values in brackets, testing the combined effect of both lags of excess capital stock per capita. All regressions include county fixed effects, time effects, and cluster at the state level. Observations and counties are from the estimates of log total production. Only counties with at least one conditional mean within 15 of the dividing line are included.

Table 9: Long term differences in income

	log income per capita in 1970			
	All rural counties		Union rural counties	
	1870	1900	1870	1900
log Total prod.	0.0455*** (0.00903)	0.112*** (0.0129)	0.0654*** (0.0129)	0.146*** (0.0130)
per capita				
Optimal capital p.c.	2.448** (0.990)	4.381*** (0.920)	1.052 (1.093)	2.500** (1.116)
Excess capital p.c.	-0.187 (1.338)	0.107 (1.678)	8.313*** (2.710)	-4.583 (3.260)
State FE	YES	YES	YES	YES
Counties	970	1034	684	737
Excess marg. eff.	-0.000635	0.000364	0.0209	-0.0115
p-val excess	0.889	0.949	0.00216	0.160
Opt. marg. eff.	0.00832	0.0149	0.00265	0.00630
p-val opt.	0.0134	1.93e-06	0.336	0.0251
p-val opt. = excess	0.00777	0.0146	0.00699	0.0437

Notes: The marginal effect for both optimum and excess is the effect of increasing capital per capita by \$50,000 divided by the the average 1880 county population, which is used for comparison to earlier estimates. The marginal effect of adding capital to a county is either the optimal or the excess marginal effect, not the sum. The “p-val opt. = excess” tests whether the two effects are the same (taking into account imputations). Income per person in 1970 and county vector files comes from census data and from [Minnesota Population Center \(2004\)](#). Uses rural counties constrained at some point from 1870-1900, except those in Connecticut, Delaware, Massachusetts, Rhode Island, and Vermont. These states are well banked and small, and so have few counties which are constrained and so state effects tend to remove them. The integration of optimal capital uses the additional information from 1902. Errors are allowed to clustered at the state level. The window size is 15. Capital stock is measured in 1000’s of dollars. See appendix for estimation details on the division between optimal and excess capital.

Figure 1: Growth of National and State Banks 1860-1914



Notes: Series are from [White \(1983, pp. 12–13\)](#), but originally come from [Barnett \(1911\)](#) and the Annual Report of the Comptroller of the Currency from various years. The number of state banks is from the Comptroller's series, and likely understates the number of state banks after 1887 by around 100 compared to the Barnett count which includes additional state banks. The number of state banks in the Comptroller series is nearly 400 below the Barnett series in 1886, but increases sharply to 1887. The difference appears to be a large effort by the Comptroller to obtain better information on state banks (see Annual Report of the Comptroller of the Currency 1887, p. 38).

Figure 2: Population and national banks in 1870

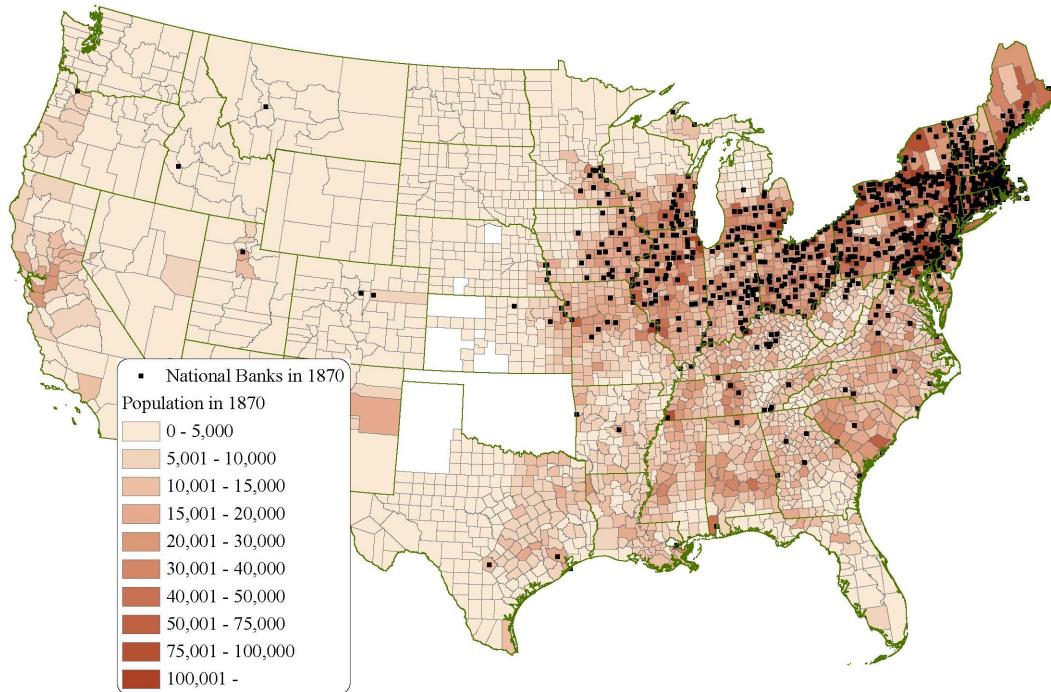


Figure 3: Population and national banks in 1880

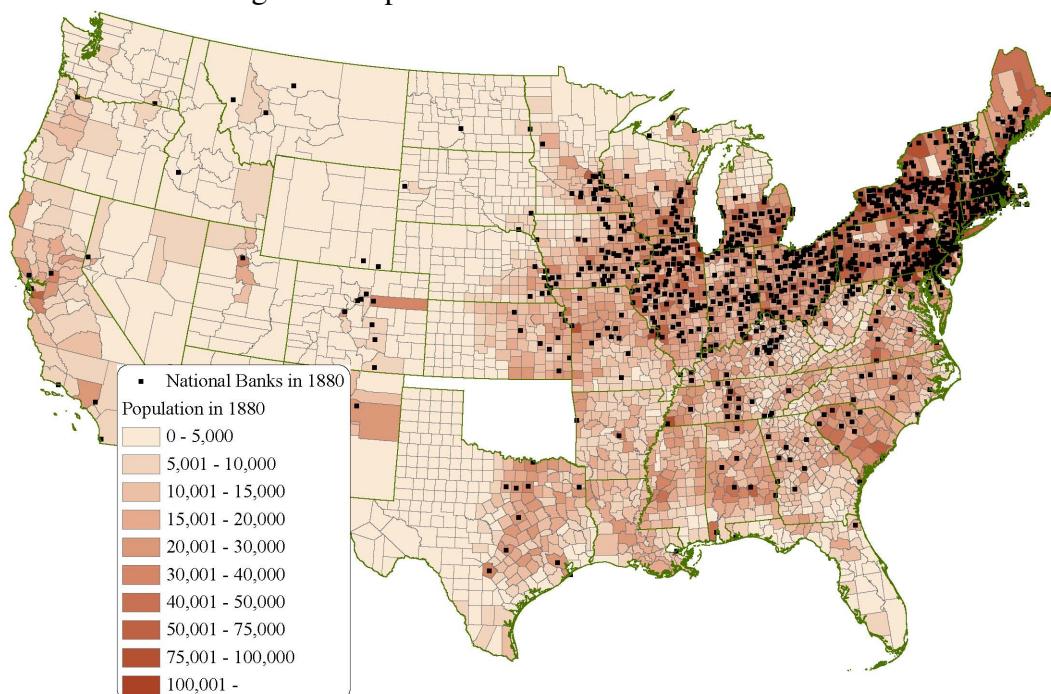


Figure 4: Population and national banks in 1890

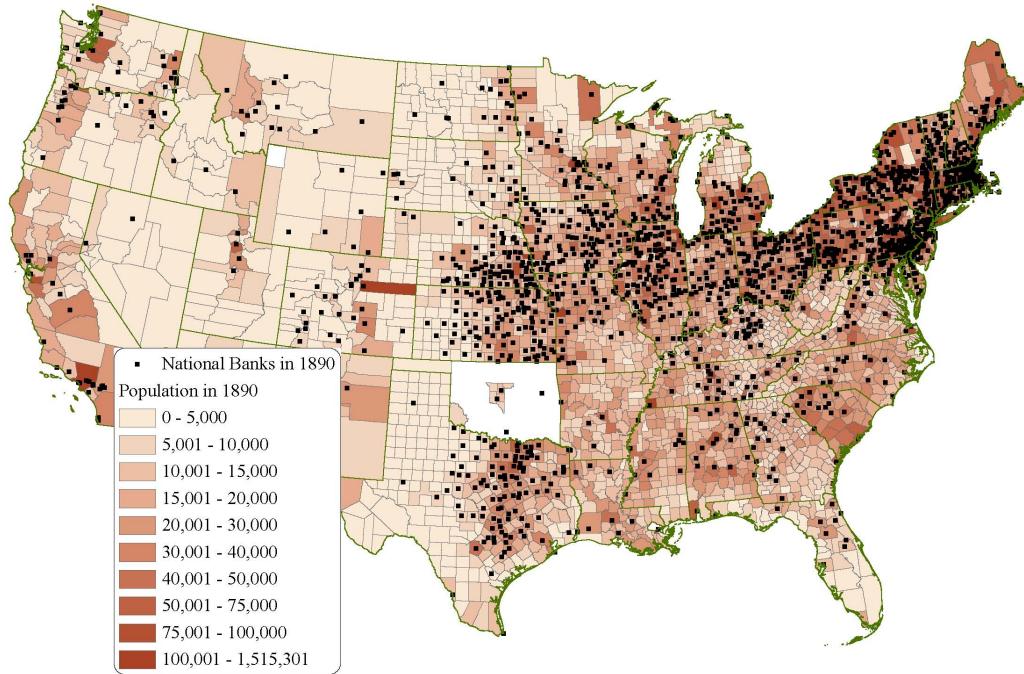
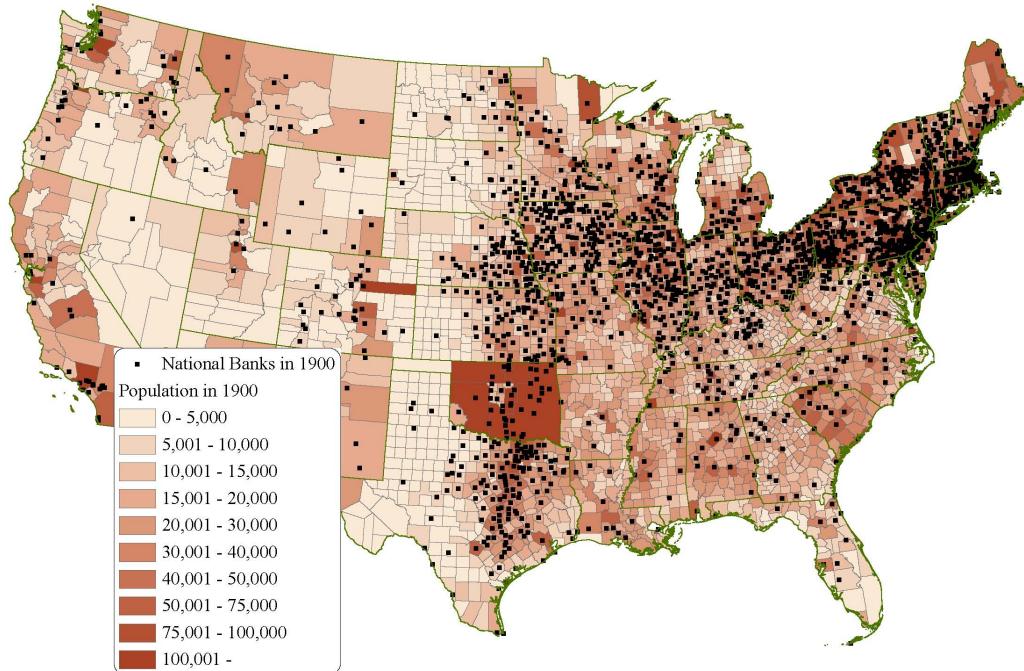


Figure 5: Population and national banks in 1900



Sources for figures 2 through 5: the national bank accounts in 1870, 1880, 1890 and 1900; [Minnesota Population Center \(2004\)](#) for the shapefiles; and [Haines and ICPSR \(2010\)](#) for the census populations. See appendix C.

Figure 6: Possible profit functions, and the observed capital

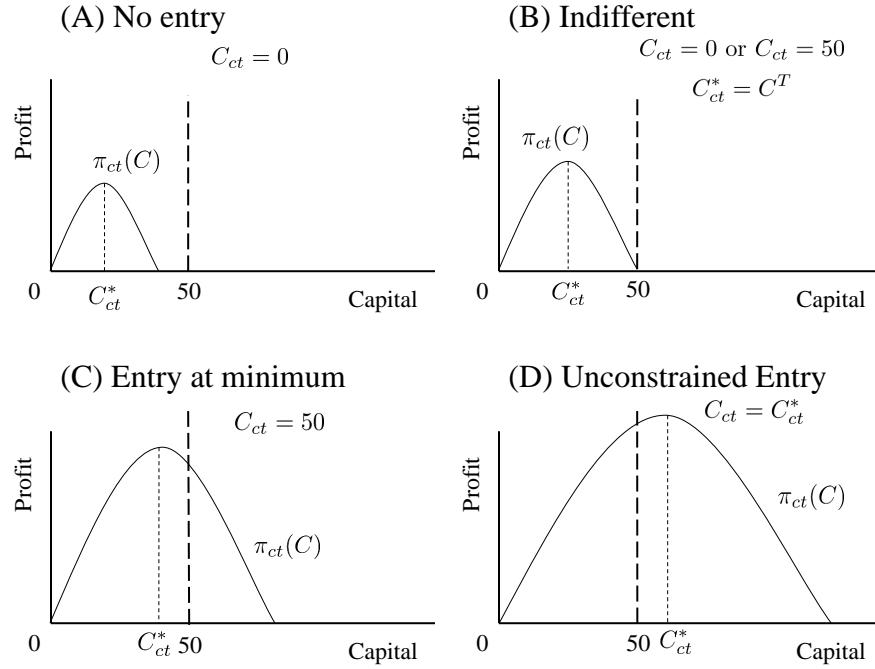


Figure 7: How observed and excess capital change with optimal capital

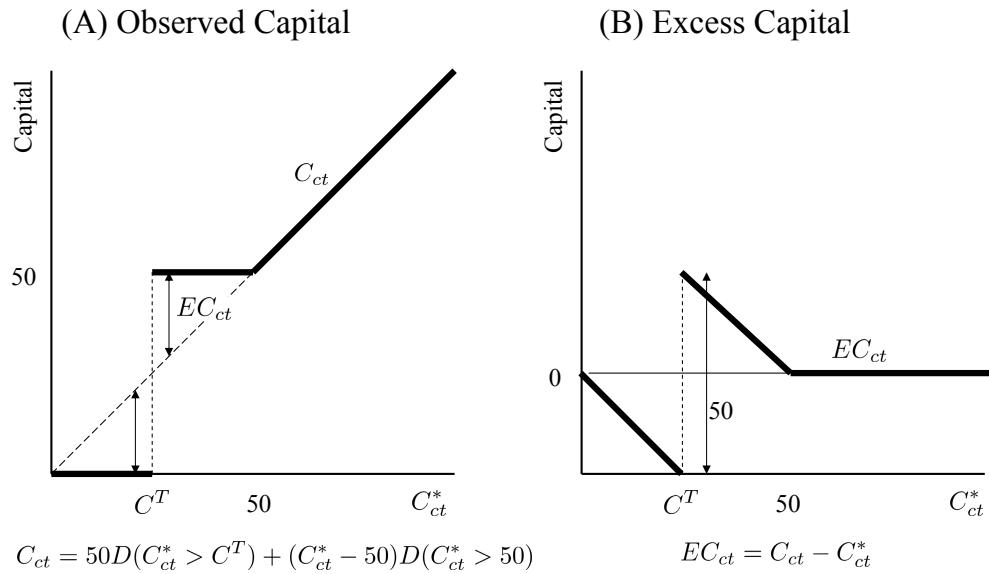


Figure 8: Conditional expectation functions

